Risk of hypoglycaemia related to the addition of DPP-4 inhibitors to sulphonylureas: systematic review and metaanalysis

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Supplement

Salvo F, et al. Risk of hypoglycaemia related the addition of DPP-4 inhibitors to sulfonylureas: systematic review and meta-analysis.

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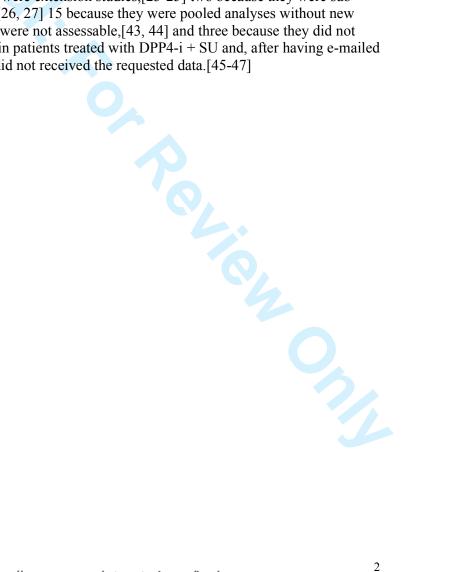
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Medline Search terms

((DPP-4[All Fields] AND ("inhibitors and inhibitors"[Subheading] OR ("inhibitors"[All Fields] AND "inhibitors" [All Fields]) OR "inhibitors and inhibitors" [All Fields] OR "inhibitors" [All Fields])) OR ("sitagliptin" [Supplementary Concept] OR "sitagliptin" [All Fields]) OR ("vildagliptin" [Supplementary Concept] OR "vildagliptin" [All Fields]) OR ("saxagliptin" [Supplementary Concept] OR "saxagliptin" [All Fields]) OR ("alogliptin" [Supplementary Concept] OR "alogliptin" [All Fields]) OR ("Linagliptin" [Supplementary Concept] OR "Linagliptin" [All Fields] OR "linagliptin" [All Fields])) AND ("randomized controlled trial" [Publication Type] OR "randomized controlled trials as topic"[MeSH Terms] OR "randomised clinical trials"[All Fields] OR "randomized clinical trials"[All Fields]

Studies excluded after full-text review: reasons for exclusion

Forty-seven studies were excluded after the full text analysis: nine because included \leq 50 patients in DPP4-i + SU group [1-9], seven because they were not RCTs, [10-16] one because there was no placebo group, [17] five because the patients were not treated with DPP4-i + SU,[18-22] three because they were extension studies,[23-25] two because they were subanalyses or post-hoc analyses, [26, 27] 15 because they were pooled analyses without new data.[28-42] two because they were not assessable.[43, 44] and three because they did not report data on hypoglycaemia in patients treated with DPP4-i + SU and, after having e-mailed authors or study contacts, we did not received the requested data.[45-47]



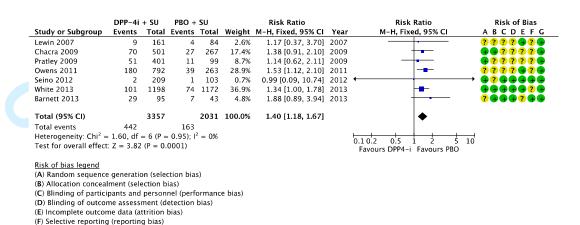
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N/A: not applicable	N/A: not applicable	Vildagliptin		
		V/A: not appl		

1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	Year	Patients with hypoglycaemia, n	Total patients, n	Treatment duration months
Feinbock et al.[49]	2003	20	111	6
Hermann <i>et al</i> .[50]	1991	12	34	6
Rosenthal &	2002	0	37	6
Mauersberger [51]				
Segal et al.[52]	1997	6	69	6
Shihara <i>et al</i> .[53]	2011	7	95	6
Spengler <i>et al</i> .[54]	1992	0	36	6
Fosi <i>et al</i> .[55]	2003	2	22	6
DeFronzo <i>et al.</i> [56]	1995	6	209	7
Charbonnel et al.[57]	2005	63	626	12
Hanefeld et al. [58]	2011	25	207	12
Kaku <i>et al</i> .[59]	2011	55	139	12
Nakamura <i>et al</i> .[60]	2006	6	18	12
Nathan <i>et al</i> .[61]	1988	0	16	9
St John Sutton et al.[62]	2002	7	99	12
Гап <i>et al</i> .[63]	2004	32	109	12
van de Laar <i>et al</i> .[64]	2004	1	50	7
ADOPT Study [65]	2006	557	1447	48
Alvarsson <i>et al.</i> [66]	2010	7	26	72
APPROACH Study [67]	2010	96	339	19
Birkeland <i>et al</i> .[68]	1994	0	30	15
Birkeland et al.[69]	2002	0	18	42
Derosa et al.[70]	2004	0	81	14
Foley & Sreenan [71]	2009	14	546	24
lain <i>et al.</i> [72]	2006	61	251	13
LEAD-3 et al.[73]	2006	60	248	45
UKPDS 33 Study[74]	1998	177	1234	120
UKPDS 34 Study [75]	1998	52	277	128

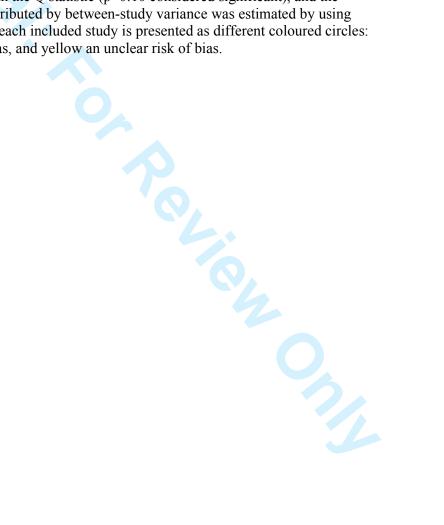
eTable 2. Trial data used to calculate the Assumed Control Risk (ACR) of hypoglycaemia; from Hemmingsen *et al.* [48]

(G) Other bias

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eFigure 1. Forest plot of the risk of hypoglycaemia in patients treated with DPP4-i + SU in comparison with those treated with placebo + SU and included in studies with low or unknown risk of bias. Risk ratios (RR) calculated for individual randomized controlled trials (RCTs) with 95% confidence intervals (CI) are presented. Arrows indicate the CI exceeding the limits of the graph. Pooled RR is also presented (black diamond). Statistical heterogeneity among studies was evaluated with the Q statistic (p<0.10 considered significant), and the proportion of total variation contributed by between-study variance was estimated by using the I² index. The risk of bias for each included study is presented as different coloured circles: green represents a low risk of bias, and yellow an unclear risk of bias.



	DPP4-i	+ SU	PBO +	SU		Risk Ratio		Risk Ratio	Risk of Bia
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	Year	M-H, Fixed, 95% CI	ABCDE
Hermansen 2007	27	222	4	219	2.1%	6.66 [2.37, 18.71]	2007		→ ?♀♀●♀●
Garber 2008	8	339	1	176	0.7%	4.15 [0.52, 32.94]	2008		→ 🗣 🗣 ? ? ? (
Chacra 2009	70	501	27	267	18.2%	1.38 [0.91, 2.10]	2009	+ 	
Pratley 2009	51	401	11	99	9.1%	1.14 [0.62, 2.11]	2009		
Kikuchi 2010	2	102	1	100	0.5%	1.96 [0.18, 21.28]	2010		→ �����
Owens 2011	180	792	39	263	30.2%	1.53 [1.12, 2.10]	2011		?????
Seino 2012	2	209	1	103	0.7%	0.99 [0.09, 10.74]	2012	+	→ - ?
White 2013	101	1198	74	1172	38.6%	1.34 [1.00, 1.78]	2013		66666
Total (95% CI)		3764		2399	100.0%	1.52 [1.27, 1.81]		•	
Total events	441		158						
Heterogeneity: Chi ² =	10.70, df	f = 7 (P	= 0.15);	$I^2 = 35$	%			0.10.2 0.5 1 2 5	10
Test for overall effect:	Z = 4.65	(P < 0.	00001)					Favours DPP4-i Favours PBO	10
Risk of bias legend (A) Random sequence (B) Allocation conceal (C) Blinding of particit (D) Blinding of outcom (E) Incomplete outcon (F) Selective reporting	ment (sele pants and ne assessi ne data (at	ection b person ment (d ttrition	ias) nel (perfe etection	ormanc	e bias)				

eFigure 2. Forest plot of the risk of hypoglycaemia in patients treated with DPP4-i + SU in comparison with those treated with placebo + SU including RCTs with a wellbalanced sex ratio among groups. Risk ratios (RR) calculated for individual randomised controlled trials (RCTs) with 95% confidence intervals (CI) are presented. Arrows indicate the CI exceeding the limits of the graph. Pooled RR is also presented (black diamond). Statistical heterogeneity among studies was evaluated with the Q statistic (p<0.10 considered significant), and the proportion of total variation contributed by between-study variance was estimated by using the I² index. The risk of bias for each included study is presented as different coloured circles: green represents a low risk of bias, red a high risk of bias, and yellow an unclear risk of bias.



	DPP4-i	+ SU	PBO +	SU		Risk Ratio		Risk Ratio	Risk of Bias
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	Year	M-H, Fixed, 95% CI	ABCDEFG
3.1.1 Hypoglycaemia	defined								
Garber 2008	8	339	1	176	4.4%	4.15 [0.52, 32.94]	2008		→ ??? _ _
Pratley 2009	51	401	11	99	59.6%	1.14 [0.62, 2.11]	2009	— — —	
Kikuchi 2010	2	102	1	100	3.4%	1.96 [0.18, 21.28]	2010	· · · · · · · · · · · · · · · · · · ·	→ @@@@@@ @
Barnett 2013	29	95	7	43	32.6%		2013		? + + + ? + 4
Subtotal (95% CI)		937		418	100.0%	1.54 [0.99, 2.42]		◆	
Total events	90		20						
Heterogeneity: Chi ² =				$^{2} = 0\%$					
Test for overall effect	Z = 1.90	(P = 0.	06)						
3.1.2 Hypoglycaemia	not defi	had							
Hermansen 2007	27	222	4	219	2.2%	6.66 [2.37, 18.71]	2007		
Lewin 2007	27	161	4	84	2.9%				2222222
Chacra 2009	70	501	27		19.7%				
Owens 2011	180	792	39	263	32.7%				2222222
Seino 2012	2	209	1	103	0.7%	0.99 [0.09, 10.74]		·	
White 2013	101	1198	-	1172	41.7%			- - -	
Subtotal (95% CI)		3083			100.0%			•	
Total events	389		149						
Heterogeneity: Chi ² =	9.15, df	= 5 (P =	= 0.10); I	² = 45%	6				
Test for overall effect	Z = 4.52	(P < 0.	00001)						
								0.10.2 0.5 1 2 5 1	
								Favours DPP4-i Favours PBO	0
Test for subgroup dif	ferences:	$Chi^2 = 0$).00, df =	= 1 (P =	: 0.95), I ²	= 0%			
<u>Risk of bias legend</u>									
(A) Random sequence				5)					
(B) Allocation conceal									
(C) Blinding of partici					e bias)				
(D) Blinding of outcor				bias)					
(E) Incomplete outcon			bias)						
 (F) Selective reporting (G) Other bias 	(reportin	g bias)							

eFigure 3. Forest plot of the risk of hypoglycaemia in patients treated with DPP4-i + SU in comparison with those treated with placebo + SU according to the presence of a definition of hypoglycaemia in the included RCTs. Risk ratios (RR) calculated for individual randomized controlled trials (RCTs) with 95% confidence intervals (CI) are presented. Arrows indicate the CI exceeding the limits of the graph. Pooled RR is also presented (black diamond). Statistical heterogeneity among studies was evaluated with the Q statistic (p<0.10 considered significant), and the proportion of total variation contributed by between-study variance was estimated by using the I² index. The risk of bias for each study included is presented as different coloured circles: green represents a low risk of bias, red a high risk of bias, and yellow an unclear risk of bias.

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ABSTRACT

Objective: Quantify the risk of hypoglycaemia associated with the concomitant use of dipeptidyl peptidase-4 inhibitors (DPP4-i) and sulphonylureas (SU) in comparison with placebo (PBO) and SU.

Design: Systematic review and meta-analysis. The Cochrane Collaboration tool for assessing risk of bias in randomized trials was used for quality assessment. The Risk Ratio (RR) of hypoglycaemia with 95% confidence intervals (95%CI) was computed for each study and then pooled. The Number Needed to Treat for one person to be Harmed, NNT(H), was estimated.

Data source: Medline, ISI Web of Science, SCOPUS, Cochrane Central Register of
Controlled Trials, and clinicaltrial.gov were searched without any language restriction.
Eligibility criteria for selecting studies: PBO-controlled randomized trials with at least 50
Type II diabetic patients treated with DPP4-i + SU.

Results: The 10 studies included represented a total of 6,546 patients (4,020 received DPP4-i + SU, 2,526 PBO + SU). The RR of hypoglycaemia was 1.52 (95% confidence interval 1.29 to 1.80) with a corresponding NNT(H) of 10 (6 to 17). The subgroup analysis by dose did not reveal any difference between full and low DPP4-i doses: the RR related to full dose DPP4-i was 1.66 (1.34 to 2.06), with a corresponding NNT(H) of 8 (5 to 15). The increased RR related to low dose DPP4-i did not reach significance (RR 1.33; 0.92 to 1.94).

Conclusions: Addition of DPP4-i to SU in patients with type II diabetes would lead to about a 50% increase in risk of hypoglycaemia and to a supplementary case of this for every 10 patients treated. This highlights the need to respect recommendations for a decrease in SU dose when initiating DPP4-i and to assess the effectiveness of this risk minimization strategy.

What this paper adds

What is already known on this subject

Hypoglycaemia is a serious event that could be related to increased morbidity and mortality in Type II diabetic patients. The risk of hypoglycaemia is known to increase when DPP4 inhibitors are used concomitantly with sulphonylureas. However, the magnitude of this risk has not yet been measured.

What this study adds

We found about a 50% increase in risk of hypoglycaemia and a supplementary case for every 10 patients treated with DPP4 inhibitors and sulphonylureas in comparison with patients treated only with SU. Thus, the recommendations for a decrease in SU dose when initiating DPP4 inhibitors must be followed, even though the effectiveness of this risk minimization strategy has not yet been assessed.

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INTRODUCTION

Hypoglycaemia is a potentially life-threatening event associated with an increased risk of hospital admission,[1] cardiovascular disease and mortality.[2, 3] This is illustrated in the ACCORD (Action to Control Cardiovascular Disease in Diabetes) trial that evaluated intensive glucose lowering in Type II diabetic patients in whom a 2.5-fold increase in hypoglycaemic events was found. That trial was prematurely stopped for reasons of increased mortality possibly related to the unfavourable effect of hypoglycaemia in susceptible patients, such as those with underlying coronary diseases.[4, 5]

Hypoglycaemia has emerged as a leading complication of diabetes in older adults with a longer history of disease. It is the second cause of hospitalisation in type II diabetic patients,[6] it can cause falls and fractures in the elderly,[7] and it accounts for 20%-25% of hospital admissions for adverse drug reactions.[1, 8] More generally, hypoglycaemia has a negative impact on patient quality of life,[9, 10] and, in the long-term, may impair the maintenance of euglycaemia and the full benefit of treatments.[11] Moreover, the importance of mild-to-moderate (iatrogenic) hypoglycaemia should not be neglected as this may lead to hypoglycaemia unawareness (through altered adrenergic response to hypoglycaemia).[12] This may compromise behavioural defences (hunger resulting in carbohydrate ingestion), and increase the risk of recurrent episodes and severe hypoglycaemia.[13, 14] Therefore, hypoglycaemia is a serious adverse event that must be considered when studying the safety of glucose-lowering drugs.

Dipeptidyl peptidase 4-inhibitors (DPP4-i) are a recently marketed class of oral anti-diabetic drugs indicated as a second line treatment in patients with Type II diabetes mellitus not adequately responsive or intolerant to metformin, or in whom treatment with other glucose-lowering drugs (such as sulphonylureas, SU, or thiazolidinediones) is insufficient to achieve glycaemic control. The mechanisms of action of these anti-diabetic drugs are different. For

instance, target tissue sensitivity to insulin is increased by thiazolidinediones,[15] hepatic gluconeogenesis is suppressed by metformin,[16] and insulin secretion is increased indirectly by DPP4-i (via the inhibition of incretin catabolism[17]) and directly by SU.[18, 19] A number of randomised clinical trials (RCTs) have studied DPP4-i both in monotherapy and, more frequently, in patients treated with other glucose-lowering drugs, metformin in particular, but also thiazolidinediones and SU.[20] When used in monotherapy, DPP4-i has shown an incidence of hypoglycaemia comparable to that related to placebo or metformin (around 5%),[21, 22] and a number of RCTs indicate that this risk is not increased when DPP4-i are used in patients treated with metformin or thiazolidinediones, thus confirming their acceptable safety profile.[20, 23]

Conversely, when DPP4-i are used in association with SU, an increased frequency of hypoglycaemia has been noted.[24, 25] This could be related to the higher frequency of hypoglycaemia among SU-treated patients (about 20% and increases as a function of treatment duration)[26] that is further increased when patients are treated by a second drug acting on insulin secretion. While the summaries of the product characteristics (SmPCs) of DPP4-i acknowledge the increased risk of hypoglycaemia due to this association,[27-31] this risk remains insufficiently assessed and has yet to be quantified. Thus, a meta-analysis to quantify the risk of hypoglycaemia associated with the use of DPP4-i and SU in patients with Type II diabetes mellitus was performed.

METHODS

Eligibility criteria

Clinical trials eligible for this meta-analysis were those: i) that studied the effect of adding one DPP4-i to SU, with or without other oral antidiabetic drug(s), in Type II diabetics; ii) that studied one DPP4-i used at daily doses approved in clinical practice, namely alogliptin (trade

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name Nesina[®] in the US, and Vipidia[®] in Europe), linagliptin (trade name Tradjenta[®] or Trajenta[®]), saxagliptin (trade names Onlgyza[®], or Kombiglyze[®] when in fixed combination with metformin), sitagliptin (Januvia[®], Ristaben[®], Tesavel[®], Xelevia[®], and Efficib[®], or Janumet[®], Ristfor[®] and Velmetia[®] when in fixed combination with metformin), and vildagliptin (Jalra[®], Xiliarx[®], Galvus[®], and Eucreas[®], or Icandra[®] and Zomarist[®] when in fixed combination with metformin); iii) that were randomized; iv) that were placebo-controlled; v) that included at least 50 patients treated with DPP4-i. Reports concerning RCT extension phases were not eligible.

Patient involvement

No patients were involved in setting the research question or the outcome measures, nor were they involved in developing plans for design or implementation of the study. No patients were asked to advise on interpretation or writing up of results. There are no plans to disseminate the results of the research to study participants or the relevant patient community.

Search strategy

Medline, ISI Web of Science, SCOPUS and Cochrane Central Register of Controlled Trials databases were searched on 15 October 2013 using keywords related to *DPP-4 inhibitors* and *randomised controlled trials*. The detailed list of keywords used to search the Medline database is provided in the Supplement. In addition, articles in the "Related citations in PubMed" were screened and a snowballing procedure was conducted to examine the references cited in systematic reviews and meta-analyses retrieved through the systematic search. *Clinicaltrials.gov* was also periodically investigated in order to identify and include hitherto unpublished but eligible RCTs. The last search in *clinicaltrials.gov* was performed in

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November 2014. No time or language restriction was applied to the searches. EndNote X6 for Macintosh (Thomson Reuters) was used to compile the bibliography.

Study selection

Two authors (FS and AP) independently reviewed and screened the title and abstract of potentially relevant RCTs and determined final eligibility through examination of full texts. Disagreements were resolved through discussion. Each eligible RCT was checked for the presence of the number of patients treated with DPP4-i + SU, with PBO + SU, and for the number of patients with at least one episode of hypoglycaemia in each treatment group. If part of these data were unavailable in the full text, missing information was requested by email to the study authors or study contacts.

Data extraction

Two authors (FS and AP) independently extracted the following information: i) methods: study design, study duration, and allowed use of other glucose-lowering drugs; ii) participants: age, gender, country, setting, and baseline mean glycated haemoglobin A_{1C} (HbA_{1c}); iii) intervention: DPP4-i and SU international non-proprietary names, daily doses, and number of treated patients; iv) hypoglycaemia: definition of hypoglycaemia used in the study. Disagreements were resolved through discussion and/or revision of the full text.

Quality assessment and evidence quality

Study quality assessment was performed using the Cochrane Collaboration tool for assessing risk of bias in randomized trials through examination of the full text or the original study protocol (as published or reported in *clinicaltrial.gov*) of the included studies.[32] The quality assessment considered the following items: i) random sequence generation; ii) allocation

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concealment; iii) blinding of participants, personnel, and outcome assessors; iv) incomplete outcome data; v) selective outcome reporting; vi) other potential biases. The risk of bias for each of these items was assessed as high, low or unknown. The GRADE framework was used to determine the strength of evidence of the meta-analysis.[33] This approach is used to contextualize or justify recommendations; it grades the quality of evidence resulting from a meta-analysis from very low to high, which corresponds to how likely further research might alter conclusions drawn from the current evidence. "High quality" suggests that it is very unlikely for conclusions about effect estimates to change, whereas "very low quality" means very likely for conclusions about effect estimates to change.[34]

The study was performed in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) statement (see research checklist supplement).[35]

Statistical analysis

The risk of hypoglycaemia in patients treated with DPP4-i + SU was estimated in comparison with that in patients treated with PBO + SU. All studies meeting the inclusion criteria were included in the quantitative analysis, irrespective of their quality.[32] The risk ratio (RR) of hypoglycaemia and its 95% confidence of intervals (95%CI) were computed for each study. The pooled RR was computed using fixed-effect models (Mantel-Haenszel method)[36] or, in the event of significant heterogeneity between estimates, using random-effect models.[37] Mantel-Haenszel method was used as it has been shown to have better statistical properties than inverse variance methods when included studies report few events [38], which is the case in a meta-analysis investigating the risk of hypoglycaemia in RCTs investigating primarily the efficacy of glucose-lowering drugs. Statistical heterogeneity among studies was evaluated using the Q-statistic (p<0.10 considered significant), and the

proportion of total variation contributed by between-study variance was estimated using the I² index.[39] All P values were two-sided.

The primary analysis concerned all studies meeting the inclusion criteria; secondary analyses were performed classifying the DPP4-i doses into full and low daily dose (as mentioned in the corresponding SmPC, the latter are mostly recommended in patients with renal impairment; see eTable 1 in Supplement), and according to the presence of a clear definition of hypoglycaemia. The forest plot of each analysis presents the subgroups which were compared using the Cochrane Q test and the I₂ index.[38] Moreover, sensitivity analyses were conducted by excluding studies with a high risk of bias (i.e. at least one item), studies allowing the use of insulin, or studies for which one or more patients characteristics were imbalanced among groups.

Publication bias was evaluated by using a funnel plot and Egger's regression test (p<0.05 considered significant).[40] The number of patients needed to be treated to observe a harmful outcome (Number Needed to Treat for one person to be Harmed, NNT(H)) was estimated according to the Cochrane recommendations.[41] The Assumed Control Risk (ACR) of hypoglycaemia in SU-treated patients was calculated from a meta-analysis reported by Hemmingsen *et al.* that included 27 clinical trials from which the incidence of hypoglycaemia was calculated.[26] On the assumption that the prevalence of hypoglycaemia is related to the length of follow-up, different follow-up scenarios were created: any duration (ACR 19.9%, 23 studies), ≤ 6 months (ACR 11.6%, 7 studies), from 6.1 to 12 months (ACR: 13.3%, 9 studies), more than 12 months (ACR 22.8%, 11 studies) (see eTable 2 in the Supplement for study details).

The analyses were conducted with Review Manager software (RevMan version 5.3, The Nordic Cochrane Centre, The Cochrane Collaboration) and R software (version 2.15.3).

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All relevant aspects related to search strategy, study selection, data extraction and quality assessment, and data analysis were specified in a synopsis protocol detailing the metaanalysis objective and context, and the principles and modalities of the literature search and the data analysis were developed.

RESULTS

Study selection

The literature search identified 2,379 records from the literature databases used, 687 of which were duplicates and were thus removed. Eleven records were retrieved through other sources. Thus, the title and abstract of 1,708 individual study records were assessed, 1,650 of which were found to be irrelevant and were excluded. The remaining 57 records underwent full text examination (results detailed in the Supplement); 10 were finally included in this meta-analysis (Figure 1).[24, 42-50]

Study characteristics

The 10 selected RCTs included a total of 6,456 patients of whom 4,020 received DPP4-i + SU, and 2,526 PBO + SU. All studies were randomized and used double-blind procedures. The study reported by Barnett *et al.* included only patients aged \geq 70 years.[24] The planned follow-up of the included studies ranged from 12 to 76 weeks. The associated SU varied across the selected RCTs (Table 1). Drug therapy also included metformin in four RCTs.[24, 44, 47, 50] Use of insulin was allowed in two RCTs.[24, 50] Baseline key patient characteristics (namely mean glycated haemoglobin A_{1C}, mean age, and gender) were well balanced among the patients included in each group of included RCTs, with the exception of two studies[24, 46] in which there was a notable difference in sex ratio between the groups (Table 1).

Three RCTs studied linagliptin 5 mg/day for a total of 1,038 patients.[24, 46, 47] Vildagliptin 100 mg/day was studied in two RCTs,[43, 45] and vildagliptin 50 mg/day in one[43] for a total of 271 patients with 100 mg/day, and 170 with 50 mg/day. Alogliptin was studied once at 12.5 mg/day and once at 25 mg/day[48, 49] for a total of 308 patients with 12.5 mg/day and 302 with 25 mg/day. White *et al.* studied alogliptin at different doses (from 6.5 mg/day to 25 mg/day) in 1,198 patients receiving SU.[50] Saxagliptin (248 patients with 2.5 mg/day, and 253 with 5 mg/day)[42] and sitagliptin 100 mg/day (222 patients)[44] were each studied once. Overall, a total of 2,526 patients receiving PBO + SU were identified in the included RCTs (Table 1).

Six of the ten included RCTs did not clearly report the definition of hypoglycaemia (Table 1).[42, 44, 46, 47, 49, 50] There was a high risk of reporting bias in three of the included studies.[43-45] One RCT also presented a high risk of detection bias (Figure 2).[44] Overall, 4,020 patients received DPP4-i (2,096 at full dose, 726 at low dose, and 1,198 at undefined dose) + SU, of whom 479 patients developed hypoglycaemia (311 at full dose, 67 at low dose, and 101 at undefined dose) corresponding to an absolute risk of 11.9%; 2,526 received PBO + SU, of whom 169 developed hypoglycaemia, corresponding to an absolute risk of 6.7%.

Meta-analysis

The RR of hypoglycaemia for DPP4-i any dose + SU versus PBO + SU was 1.52 (95%CI 1.29 to 1.80), with no evidence of heterogeneity across RCTs (Q = 11.2, p = 0.26, $I^2 = 20\%$; Figure 3). For any DPP4-i +SU duration of use, the corresponding NNT(H) was 10 (6 to 17); it was 17 (11 to 30) for a treatment duration ≤ 6 months, 15 (9 to 26) for 6.1 to 12 months, and 8 (5 to 15) for a treatment duration longer than one year.

The pooled RR did not markedly change when RCTs with a high risk of detection bias and reporting bias (1.40; 1.18 to 1.67; eFigure 1 in the Supplement), or when the RCTs which allowed the use of insulin (1.61; 1.30 to 2.00), were excluded from the analysis. The RR was similar to that of the principal analysis when RCTs in which a notable imbalance in sex ratio were excluded (1.52; 1.27 to 1.81; Q = 10.70, p = 0.15; $I^2 = 35\%$; eFigure 2 in the supplement). The pooled RR was also similar for RCTs in which a definition of hypoglycaemia was reported (1.54; 0.99 to 2.42; Q = 2.1, p = 0.5, $I^2 = 0\%$), and in those in which a definition was not reported (1.52; 1.27 to 1.82; Q = 9.1, p = 0.10, $I^2 = 45\%$), without any evidence of heterogeneity between these two groups (Q = 0.0, p = 0.95, $I^2 = 0\%$; eFigure 3 in the supplement).

According to the dose of DPP4-i evaluated, the subgroup analysis showed no difference between low and full DPP4-i dose with regard to the risk of hypoglycaemia (Q = 0.99, p = 0.32, I² = 0%; Figure 4). The risk remained significantly increased for DPP4-i full dose (1.66; 1.34 to 2.06) but was not significantly increased for DPP4-i low doses (1.33; 0.92 to 1.94; Figure 5). For DPP4-i full dose+SU, the NNT(H) was 8 (5 to 15) for any treatment duration; it was 13 (8 to 25) for a treatment duration ≤ 6 months, 11 (7 to 22) for a treatment duration between 6.1 to 12 months, and 7 (4 to 13) for a treatment duration longer than one year. Visual inspection of the funnel plot did not show any clear evidence of publication bias (Figure 5), and the Egger test did not find any asymmetry (z=1.3; p=0.2). The strength of evidence of this meta-analysis was evaluated as high with regards to the GRADE framework (Table 2).

DISCUSSION

Principal findings

This meta-analysis found about a 50% increase in the risk of hypoglycaemia when DPP4-i and SU were associated in Type II diabetic patients, leading to one supplementary case of hypoglycaemia for every 10 treated patients. This risk was confirmed for full doses of DPP4i, while it could not be excluded for lower doses.

DPP4-i act indirectly on insulin levels by enforcing the incretin effect, which is a response to high oral intake of carbohydrates and fatty acids.[17] Such drugs should therefore act on glycaemia only in response to such intakes, thereby protecting patients from hypoglycaemia. However, in patients treated with SU, insulin secretion is already stimulated independently of glycaemia and the addition of a reinforced incretin effect on insulin levels leads to an increase in the risk of hypoglycaemia. Given the frequency of this event in Type II diabetic patients treated with SU, the risk associated with the addition of DPP4-i would lead to a huge number of cases of induced hypoglycaemia, some of which could be severe.[51] The present meta-analysis did not allow investigation of the threshold of dose combination (DPP4-i + SU) associated with an increased risk of hypoglycaemia; an individual patient meta-analysis could be helpful in this regard.

The risk of hypoglycaemia related to the addition of a DPP4-i to SU is acknowledged in the SmPCs for DPP4-i; most recommend using full-dose DPP4-i but a reduced SU dose in patients taking such combinations, although the magnitude of reduction is not stated.[27-31] Currently, to what extent this recommendation would lower the number of excess cases of induced hypoglycaemia is unknown. It is also of note that the suggested individual patient meta-analysis would not fill this knowledge gap as the effect of SU dose reduction has not been investigated in trials studying DPP4-i.

For low doses of DPP4-i (half the full dose when applicable), the increase in hypoglycaemia risk was not significant. However, the existence of this risk cannot be fully ruled out by the present results and a larger sample would be required to increase the precision of the

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estimates. Furthermore, although the point estimate was lower (RR 1.33 vs. 1.66 for fulldoses), which suggests a potential dose-effect, no heterogeneity was found between low and full doses of DPP4-i, yet this could result from a lack of power in the heterogeneity test (lowdose group was half the size of the high-dose group).

Strengths and limitations of study

The present analysis has important strengths. Firstly, it is based on a large sample of patients; over 4,000 treated with a combination of DPP4-i and SU, and over 2,500 treated with PBO and SU. Secondly, the overall quality of the included studies seems high according to the Cochrane Collaboration tool for risk of bias assessment, which was confirmed by the GRADE framework evaluation of the meta-analysis that considers that the strength of evidence provided is high. The present meta-analysis used data concerning all currently marketed DPP4-i (alogliptin, linagliptin, saxagliptin, sitagliptin, and vildagliptin), and results were consistent within studies with no heterogeneity being found among estimates. Thirdly, there was no evidence of publication bias; the funnel plot was balanced and the Egger test was not significant.

Nevertheless, the meta-analysis does have certain limitations. Firstly, certain studies that presented a high risk for detection and reporting bias were included in the main analysis,[43-45] but exclusion of these studies did not change the estimates significantly. Secondly, three studies could not be included as data were not available for the risk of hypoglycaemia in patients receiving SU.[52-54] However, in view of the GRADE framework, including results from these studies would be unlikely to change the results significantly owing to the size of the present meta-analysis, the high number of hypoglycaemia cases, and the confidence intervals of the pooled RR that clearly do not cross the line of no effect.[33] The absence of heterogeneity in estimates found from the 10 included studies further supports this hypothesis.

Thirdly, the results of this meta-analysis are dominated by the results of three studies that account for more than 80% of the pooled results of the principal analysis;[42, 47, 50] a sensitivity analysis without these studies did not substantially change the results of the meta-analysis (data not shown). Fourthly, the definition of hypoglycaemia varied among the included RCTs and was not reported in five. Other authors did not perform a meta-analysis on hypoglycaemia risk on the basis of this lack of homogeneity in its definition across the RCTs;[20] nevertheless, this could be considered as a minor limitation, as in the present analysis the risk did not differ between RCTs with or without a clear definition of hypoglycaemia. The incidence of hypoglycaemia also differed among studies, mainly because of different durations of follow-up. However, this did not have any impact on the estimation of the pooled risk (no statistical heterogeneity was found) nor on the NNT(H) calculation, which was based on an external Assumed Control Risk of hypoglycaemia retrieved from 27 clinical studies included in a meta-analysis of the Cochrane library.[26]

Clinical importance

It is important to underline that hypoglycaemia is the most frequent adverse reaction related to anti-diabetic treatments and that, even when not directly life-threatening, it is associated with an increased risk of all-cause mortality, cardiovascular disease, and cardiovascular mortality and hospital admission.[2, 3, 6] In addition, it should not be neglected that hypoglycaemia and its related symptoms (e.g. nervousness, sweating, trembling, weakness, palpitations) impact negatively on patient quality of life and disrupt many daily activities such as driving, work performance and leisure activities.[9, 10] More importantly, mild-to-moderate iatrogenic hypoglycaemia can decrease the usual adrenergic response to hypoglycaemia.[12] This may cause hypoglycaemia unawareness and compromise behavioural defences (hunger resulting in sugar ingestion), which in turn can lead to severe hypoglycaemia.[13, 14] It is thus an

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important to lower the risk of mild-to-moderate hypoglycaemia, which remains a serious adverse event. Adequate information regarding the risk of hypoglycaemia, whatever its severity, should thus be considered of primary importance for patients and all health professionals involved in the management of diabetic patients. Reaching good glycated haemoglobin levels should not be at the expense of hypoglycaemic events, which could outweigh the benefit of preventing risks associated with elevated blood glucose concentrations. Thus, the risk demonstrated herein for all-type hypoglycaemia should not be minimized by considering that only severe episodes would be of clinical concern.

Conclusions

In conclusion, this meta-analysis found about a 50% increase in the risk of hypoglycaemia associated with the addition of DPP4-i to SU in patients with type II diabetes. For this adverse event commonly experienced by treated diabetic patients, this would lead to the occurrence of one supplementary hypoglycaemic event in every 10 treated patients. This potentially represents a huge number of attributable cases worldwide. These results clearly highlight the need to respect existing recommendations for SU dose reduction when initiating a DPP4-i treatment, and the urgency to determine the efficacy of this measure in minimizing the risk of hypoglycaemia.

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The corresponding author had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Francesco Salvo (corresponding author) and Antoine Pariente (manuscript's guarantor) affirm that the manuscript is an honest, accurate, and transparent, and that no important aspects of the study have been omitted.

Authors' contributions

FS: conception and design; acquisition, analysis and interpretation of data; drafting the article and revising it critically for important intellectual content; final approval of the version to be published. FS gives agreement to be accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

NM: conception and design; analysis and interpretation of data; revising the article critically for important intellectual content; final approval of the version to be published. NM gives agreement to be accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved. MA: analysis and interpretation of data; revising the article critically for important intellectual content; final approval of the version to be published. MA gives agreement to be accountable

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for all aspects of the work by ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

PR: analysis and interpretation of data; revising the article critically for important intellectual content; final approval of the version to be published. PR agrees to be accountable for all aspects of the work by ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

ER: conception and design; interpretation of data; revising the article critically for important intellectual content; final approval of the version to be published. ER agrees to be accountable for all aspects of the work by ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

FDP: conception and design; interpretation of data; revising the article critically for important intellectual content; final approval of the version to be published. FDP agrees to be accountable for all aspects of the work by ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

BB: conception and design; analysis and interpretation of data; revising the article critically for important intellectual content; final approval of the version to be published. BB agrees to be accountable for all aspects of the work by ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

AP conception and design; acquisition and interpretation of data; drafting the article and revising it critically for important intellectual content; investigating accuracy and integrity of any part of the work; final approval of the version to be published. AP agrees to be accountable for all aspects of the work by ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved. AP is the guarantor.

Declaration of interests

The corresponding author ensures that the manuscript is complete and that the conflict of interest disclosures are accurate, up-to-date, and consistent with the information provided in each author's ICMJE Form for Disclosure of Potential Conflicts of Interest. All authors have read and understood BMJ policy on declaration of interests; all authors have completed the Unified Competing Interest form at www.icmje.org/coi_disclosure.pdf (available on request from the corresponding author) and declare no support from any organisation for the submitted work. FS, MA, ER, FdP, and BB, and have no with any organisations that might have an interest in the submitted work in the previous three years. NM and PR have had specified relationships on other matters with Novartis and Takeda, which might have an interest in the submitted work. AP has had specified relationships on other matters with public regulatory agencies and with health care insurance systems that might have an interest in the submitted work. All authors declare no other relationships or activities that could appear to have influenced the submitted work.

Ethical approval

This type of study does not require ethical approval.

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Data sharing

No additional data available.

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Figure legends

<text><text><text> Figure 1. Flow diagram of study identification, selection, and inclusion. The search strategy identified 2,379 records of which 687 were duplicates and removed. Fifteen references were retrieved by other sources, thus a total of 1,707 individual titles and abstracts were assessed, leading to the exclusion of 1,650 records. After evaluation of 57 full texts, 13 studies were eligible for this meta-analysis. Data from three studies were not available so 10 studies were included.

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Figure 3. Forest plot of the risk of hypoglycaemia in patients treated with DPP4-i + SU in comparison with those treated with PBO + SU. Risk ratios (RR) calculated for individual randomized controlled trials (RCTs) with 95% confidence intervals (CI) are presented. Arrows indicate the CI exceeding the limits of the graph. Overall RR is also presented (black <text><text><text><text><text> diamond). An estimate of the weight of each RCT on overall RR is reported as a percentage and graphically (blue square size). Statistical heterogeneity among studies was evaluated with the Q statistic (p < 0.10 considered significant), and the proportion of total variation contributed by between-study variance was estimated by using the I² index. The risk of bias for each study included is presented as different coloured circles: green represents a low risk of bias, red represents a high risk of bias, yellow represents an unclear risk of bias.

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Figure 4. Forest plot of the risk of hypoglycaemia in patients treated with full or low DPP4-i doses + SU in comparison with those treated with PBO + SU. Risk Ratios (RR) calculated for individual randomized controlled trials (RCTs) with 95% confidence interval (CI) are presented. Arrows indicate the CI exceeding the limits of the graph. For each subgroup, an estimate of the weight of each RCT on pooled RRs is reported as a percentage and graphically (black square size). Pooled RRs for low and full doses are also presented (black diamonds). Statistical heterogeneity among studies was evaluated with the Q statistic (p < 0.10 considered significant), and the proportion of total variation contributed by between-study variance was estimated by using the I^2 index. The risk of bias for each included study is presented as different coloured circles: green represents a low risk of bias, red represents a high risk of bias, yellow represents an unclear risk of bias.

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Table 1. Study characteristics

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Table 1. Study characteristics							
	Study duration, weeks	Intervention, daily dose (n)	Associated SU	Mean HbA _{1c} at baseline, %	Mean age of participants, years	Male, %	Definition of hypoglycaemia
Barnett et al.[24]	24	Linagliptin 5 mg (95 pts) or PBO (43 pts)	SU, not specified	DPP4-i: 7.8 PBO: 7.7*	DPP4-i: 75 PBO: 75*	DPP4-i: 72 PBO: 62*	PG of 3.9 mmol/l or less, with or without symptoms
Chacra et al.[42]	24	Saxagliptin 2.5 mg (248 pts), saxagliptin 5 mg (253 pts), or PBO (267 pts)	Glyburide	DPP4-i: 8.4-8.5 PBO: 8.4	DPP4-i: 55 PBO: 55	DPP4-i: 45 PBO: 46	Not reported
Garber <i>et al.</i> [43]	24	Vildagliptin 50 mg (170 pts) or 100 mg (169 pts), or PBO (176 pts)	Glimepiride	DPP4-i: 8.5-8.6 PBO: 8.5	DPP4-i: 58-59 PBO: 58	DPP4-i: 59 PBO: 58	Symptomatic hypoglycaemia confirmed by self-monitored BG <3.1 mmol/l
Hermansen et al. [44]	24	Sitagliptin 100 mg (222 pts) or PBO (219 pts)	Glimepiride	DPP4-i: 8.3 PBO: 8.3	DPP4-i: 56 PBO: 56.5	DPP4-i: 53 PBO: 53	Not reported, but hypoglycaemia is included in the AEs of special interest
Kikuchi et al.[45]	12	Vildagliptin 100 mg (102 pts) or PBO (100 pts)	Glimepiride	DPP4-i: 7.8 PBO: 8.0	DPP4-i: 59 PBO: 60	DPP4-i: 73.5 PBO: 69	Symptomatic hypoglycaemia, confirmed by self-monitored BG <3.1 mmol/l
Lewin <i>et al.</i> [46]	18	Linagliptin 5 mg (161 pts) or PBO (84 pts)	SU, not specified	DPP4-i: 8.6 PBO: 8.6	DPP4-i: 57 PBO: 56	DPP4-i: 48 PBO: 62	Not reported, but hypoglycaemia were recorded and analyzed separately from other AEs.
Owens et al.[47]	24	Linagliptin 5 mg (792 pts) or PBO (263 pts)	SU, not specified	DPP4-i: 8.1 PBO: 8.1	DPP4-i: 58 PBO: 58	DPP4-i: 48 PBO: 47	Not reported
Pratley et al. [48]	26	Alogliptin 12.5 mg (203 pts), alogliptin 25 mg (198 pts), or PBO (99 pts)	Glyburide	NR	DPP4-i: 56.5 PBO: 57	DPP4-i: 52 PBO: 51.5	Symptomatic hypoglycaemia with BG <3.3 mmol/l or BG <2.8 mmol/l without symptoms
Seino et al.[49]	12	Alogliptin 12.5 mg (105 pts), alogliptin 25 mg (104 pts), or PBO (103 pts)	Glimepiride	DPP4-i: 8.5% PBO: 8.6%	DPP4-i: 60 PBO: 60	DPP4-i: 66 PBO: 69	Not reported
White <i>et al.</i> [50]	76**	Alogliptin any doses (1,198), or PBO (1,172 pts)	SU, not specified	DPP4-i: 8.0 PBO: 8.0*	DPP4-i: 61 PBO: 61***	DPP4-i: 68 PBO: 69*	Not reported

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HbA_{1c} Glycated hemoglobin A_{1C}; NR: not reported; PBO: placebo; Pts: patients; SU: sulphonylureas; y=years old. PG: Plasma Glucose; mmol/l: millimols/liter BG: Blood Glucose; AEs: adverse events.

* Data refer to overall study population, not only to SU treated patients.

** Median exposure weeks for alogliptin treated patients.

*** Median age (years).

Quality assessment				Quality assessment № of patients				atients	Ef	ffect		
№ of studies	Study design	Risk of bias	Inconsistency	Indirectness	Imprecision	Other considerations	All studies	Placebo	Relative (95% CI)	Absolute (95% CI)	Quality	1
Hypogly	caemia	I	9	0.	1		1	1	I			
10	randomized trials	not serious ¹	not serious ²	not serious	not serious $\frac{3}{2}$	dose response gradient	479/4,020 (11.9%)	169/2,526 (6.7%)	RR 1.52 (1.29 to 1.80)	35 more per 1,000 (from 19 more to 54 more)	⊕⊕⊕⊕ HIGH	CI
	2. No het 3. The sa (lower 4. Hypog	were exclud erogeneity as mple size is bound of 95 lycaemia is t	led from the analy mong estimates w large (n=6,526), tl %CI = 1.29) he most frequent	vsis the result die as found. he number of the adverse reaction	d not change sub e events high (64 n related to anti-c ss, sweating, trer	Among them, a high stantially. 18), and the confider liabetic treatment. It nbling, weakness, p	nce intervals of t increases the r alpitations) red	the pooled RR	clearly do no	ot cross the line	of no effec ar events.	et

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Importance

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	Risk of hypoglycaemia related to the addition of DPP-4 inhibitors plusto
	sulphonylureas: systematic review and meta-analysis- Francesco Salvo, clinical pharmacologist (1,2), Nicholas Moore, professor of pharmacology
	(1,2,3), Mickael Arnaud, statistician (1), Philip Robinson, medical writer (3,4), Emanuel
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19	ABSTRACT
20	Objective: Quantify the risk of hypoglycaemia associated with the concomitant use of
21	dipeptidyl peptidase-4 inhibitors (DPP4-i) and sulphonylureas (SU) in comparison with those
22	treated with placebo (PBO) and SU.
23	Design: Systematic review and meta-analysis. The Cochrane Collaboration's Collaboration
24	tool for assessing risk of bias in randomized trials was used for quality assessment. The Risk
25	Ratio (RR) of hypoglycaemia with 95% confidence intervals (95%CI) was computed for each
26	study and thethen pooled. The number of patients needed to be treated to observe a harmful
20 27	outcome (Number Needed to Harm, NNH)Treat for one person to be Harmed, NNT(H), was
28	estimated-and presented in forest plot.
	Data source: Medline, ISI Web of Science, SCOPUS, Cochrane Central Register of
29	Controlled Trials, and clinicaltrial.gov were searched without any language restriction.
30	Eligibility criteria for selecting studies: PBO-controlled randomized trials with at least 50
31	Type II diabetic patients treated with DPP4-i + SU.
32	Results: The ten10 studies included represented a total of 6,546 patients (4,020 received
33	DPP4-i + SU, 2,526 PBO + SU). The RR of hypoglycaemia was 1.52 (95% confidence
34	interval 1.29 to 1.80) with a corresponding NNHNT(H) of 26.9 (19.5 to 43.3). The10 (6 to
35	17). The subgroup analysis by dose did not reveal any difference between full and low DPP4-i
36	doses: the RR related to full dose DPP4-i was 1.66 (1.34 to 2.06), with a corresponding
37	NNHNNT(H) of 19.4 (13.98 (5 to 32.215). The increased RR related to low dose DPP4-i did
38	not reach significance (RR 1.33; 0.92 to 1.94).
39	Conclusions: AssociatingAddition of DPP4-i withto SU in patients with type II diabetes
40	would lead to about a 50% increase in risk of hypoglycaemia and to a supplementary case of
40 41	this for every 27 treated 10 patients treated. This highlights the need to strictly respect
	uns foi every an active patients <u>incared</u> . This inginigits the need to strengt espect
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$1 \\ 2 \\ 3 \\ 4 \\ 5 \\ 6 \\ 7 \\ 8 \\ 9 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1$	recommendations for a decrease in SU dose when initiating DPP4-i ₂ and to urgoinly assess the effectiveness of this risk minimization strategy.	
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19	What this paper adds
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21	What is already known on this subject
22	Hypoglycaemia is a serious event that could be related to an-increased morbidity and
	mortality in Type II diabetic patients. It is known the The risk of hypoglycaemia is
23	increasedknown to increase when DPP4-i inhibitors are used concomitantly with
24	SU-sulphonylureas. However, itsthe magnitude of this risk has not yet been measured.
25	
26	What this study adds

26	What this study adds
27	AWe found about a 50% of increase in risk of hypoglycaemia and a supplementary case for
28	every 2710 patients treated with DPP4-i inhibitors and SUsulphonylureas in comparison with
29	patients treated only with SU-was found. Thus, the recommendations for a decrease in SU
30	dose when initiating DPP44 inhibitors must be followed, even though the effectiveness of this
30 31	risk minimization strategy has not yet been assessed.
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INTRODUCTION

Hypoglycaemia is a potentially life-threatening event associated with an increased risk of hospital admission.[1] cardiovascular disease; and mortality [2, 3] 4-This is illustrated in the ACCORD (Action to Control Cardiovascular Disease in Diabetes) trial evaluatingthat evaluated intensive glucose lowering in Type II diabetic patients, in whom a 2.5-fold increase in hypoglycaemic events was noted. This found. That trial was prematurely stopped for reasons of increased mortality, possibly related to the unfavourable effect of hypoglycaemia in susceptible patients, such as those with underlying coronary diseases.[4, 5] Studies conducted within US homitals found that hypoglycaemia accounted for 20% of homital

sions attributed to adverse drug reactions.[1] with a median four days of hospital

Hypoglycaemia has emerged as a leading complication of diabetes in older adults with a lenger history of disease. It is the second cause of hospitalisation in type II diabetic patients.[6] it can cause falls and fractures in the elderly,[7] and it accounts for 20%-25% of hospital admissions for adverse drug reactions.[1.8] More generally, hypoglycaemia has a negative impact on patient quality of life.[9.10] and, in the long-term, may impair the maintenance of euglycaemia and the full benefit of treatments.[11] Moreover, the importance of mild-to-moderate (iatrogenic) hypoglycaemia should not be neglected as this may lead to hypoglycaemia unawareness (through altered adrenergic response to hypoglycaemia.).[12] This may compromise behavioural defences (hunger resulting in carbohydrate ingestion), and increase the risk of recurrent episodes and severe hypoglycaemia.[13,14] Therefore, hypoglycaemia is a serious adverse event that must be considered when studying the safety of

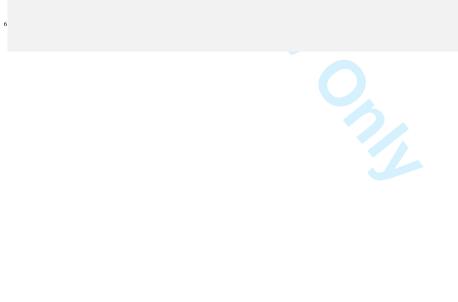
Dipeptidyl peptidase 4-inhibitors (DPP4-i) are a recently marketed class of oral antidiabeticanti-diabetic drugs indicated as a second line treatment in patients with Type II

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18	diabetes mellitus not adequately responsive or intolerant to metformin, or in whom treatment	
19	with other glucose-lowering-drugs (such as sulphonylureas, SU, or thiazolidinediones) is	
20	insufficient to achieve glycaemic control. Netably, the The mechanisms of action of these	
21	hypoglycaemicanti-diabetic drugs are different. For instance, target tissue sensitivity to	
22	insulin is increased by thiazolidinediones.[715] hepatic gluconeogenesis is suppressed by	
23	metformin,[%]6] and insulin secretion is increased indirectly by DPP4-i (via the inhibition of	
24	incretin catabolism[<u>917</u>]) and directly by SU.[<u>40, 1118, 19</u>]	
25	A number of randomised clinical trials (RCTs) have studied DPP4-i both in monotherapy and	I,
26	more frequently, in patients treated with other glucose-lowering drugs- metformin in	
27	particular, but also thiazolidinediones or and SU.[12] These 20] When used in monother apy,	
28	DPP4-i has shown an incidence of hypoglycaemia comparable to that related to placebo or	
29	metformin (around 5%).[21, 22] and a number of RCTs indicate an acceptable safety profile	
30	that this risk is not increased when DPP4-i are used in patients treated with metformin or	
31	thiazolidinediones.[12, 13], thus confirming their acceptable safety profile.[20, 23]	
32	Conversely, when DPP4-i are used in association with SU ₂ an increased frequency of	
33	hypoglycaemia washas been noted.[14, 15] The24, 25] This could be related to the higher	
34	frequency of hypoglycaemia among SU-treated patients (about 20% and increases as a	
35	function of treatment duration)[26] that is further increased when patients are treated by a	
36	second drug acting on insulin secretion. While the summaries of the product characteristics	
37	(SmPCs) of DPP4-i acknowledge the increased risk of hypoglycaemia due to this	
38	association;[16-20] however, [27-31] this risk remains insufficiently assessed and it was	
39	neverhas yet to be quantified. Thus, a meta-analysis to quantify the risk of hypoglycaemia	
40	associated with the use of DPP4-i and SU in patients with Type II diabetes mellitus was	
41	performed.	
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METHODS

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Patient involvement

Search strategy

Eligibility criteria

Placebo (PBO)-controlled RCTs that studied the effect of adding DPP4-i to SU for the

RCTsClinical trials eligible for this meta-analysis were those: i) that were perfor

at of Type II diabetes mellitus were selected thorough a systematic review

other oral antidiabetic drug(s), in Type II diabetics; ii) that studied the effect of one DPP4-i

used at daily doses approved in clinical practice, in addition to SU, with or without o

linagliptin (trade name Tradjenta® or Trajenta®), saxagliptin (trade names Onlgyza®, or

Tesavel[®], Xelevia[®], and Efficib[®], or Janumet[®], Ristfor[®] and Velmetia[®] when in fixed

Icandra® and Zomarist® when in fixed combination with metformin); iii) that were

DPP4-i. Reports concerning RCT extension phases were not eligible.

Kombiglyze® when in fixed combination with metformin), sitagliptin (Januvia®, Ristaben®,

combination with metformin), and vildagliptin (Jalra®, Xiliarx®, Galvus®, and Eucreas®, or

randomized; iv) that were placebo-controlled; v) that included at least 50 patients treated with

No patients were involved in setting the research question or the outcome measures, nor were

they involved in developing plans for design or implementation of the study. No patients were

asked to advise on interpretation or writing up of results. There are no plans to disseminate

the results of the research to study participants or the relevant patient community.

-2 diabetes mellitusstudied the effect of adding one DPP4-i to SU, with or without

edrug(snamely alogliptin (trade name Nesina[®] in the US, and Vipidia[®] in Europe),

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19	Medline, ISI Web of Science, SCOPUS, and Cochrane Central Register of Controlled Trials databases were searched in 15 October 2013 using keywords related to DPP-4 inhibitors	
20	and randomised controlled trials. The detailed list of keywords used to search the Medline	
21	database is provided in supplement. the Supplement. In addition, articles in the "Related	
22 23 24 25 26 27	citations in PubMed" were screened, and a snowballing procedure was conducted to examine	
23	the references cited in systematic reviews and meta-analyses retrieved through the systematic	
24	search. Clinicaltrials.gov was also periodically investigated in order to identify and include	
25	not-yet publishedhitherto unpublished but eligible RCTs. The last search in <i>clinicaltrials.gov</i>	
26	was performed in November 2014. No time or language restriction was applied to the	
27	searches. EndNote X6 for Macintosh (Thomson Reuters) was used to compile the	
28	bibliography.	
29		
30	Study selection	
31	Two authors (FS and AP) independently reviewed and screened independently the title and	
32	abstract of the potentially relevant RCTs, and performed their <u>determined</u> final eligibility	
33	through examination of full-texts. Disagreements were solvedresolved through discussion.	
34	Each eligible RCT was checked for the presence of the number of patients treated with DPP4-	
35	i+SU, with PBO + SU, and for the number of patients with at least one episode of	
36	hypoglycaemia in each treatment group. If part of these data were unavailable in the full-text,	
37	missing information was requested by email to the principalstudy authors or study contacts.	
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39	Data extraction	
40	Two authors (FS and AP) extracted independently extracted the following information: i)	
41	methods: study design, study duration, and allowed use of metformin and dosesother glucose-	
42	lowering drugs; ii) participants: age, gender, country, setting, and baseline mean glycated	
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solvedresolved through discussion and/or revision of the full-text. Quality assessment and evidence quality Study quality assessment was performed using the Cochrane Collaboration's Collaboration tool for assessing risk of bias in randomized trials through the examination of the full-text or the original study protocol (as published or reported in clinicaltrial.gov) of the included studies.[2432] The quality assessment considered the following items: i) random sequence generation; ii) allocation concealment; iii) blinding of participants, personnel, and outcome ssessors; iv) incomplete outcome data; v) selective outcome reporting; vi) other potential biasbiases. The risk of bias for each of these items was assessed as high, low; or unknown. The GRADE framework was used to determine the strength of evidence of the metaanalysis.[2233] This approach is used to contextualize or justify recommendations; it grades the quality of evidence resulting from a meta-analysis from very low to high, which corresponds to how likely further research might to-alter conclusions drawn from the current evidence. ""High quality" quality" suggests that it is very unlikely for conclusions about effect estimates to change, whereas ""very low quality" quality" means very likely for conclusions about effect estimates to change.[2334] This The study was performed in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) statement (see research checklist supplement).[2435] Statistical analysis

haemoglobin A1C (HbA1c); iii) intervention: DPP4-i and SU international non-proprietary

names, daily doses, and number of treated patients; iv) hypoglycaemia: definition of

hypoglycaemia used in the study; v) allowed insulin treatment. Disagreements were



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10 19 The risk of hypoglycaemia in patients treated with DPP4-i + SU was estimated in corr	nnarison
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25 to have better statistical properties then inverse variance methods when included stud	
26 report few events [38] which is the case in a meta-analysis investigating the risk of	
27 hypoglycaemia in RCTs investigating primarily the efficacy of glucose-lowering drug	75
28 Statistical heterogeneity among studies was evaluated using the Q-statistic (p<0.10 cd	
29 significant), and the proportion of total variation contributed by between-study variation	
30 estimated using the l^2 index. [2739] All P values were two-sided.	
31 The primary analysis concerned all studies meeting the inclusion criteria; secondary a	inalyses
32 were performed classifying the DPP4-i doses into full and low daily dose (as mention	
33 corresponding SmPC, the latter <u>are</u> mostly recommended in patients with renal impai	
34 see eTable 1 in supplement), Supplement), and according to the presence of a clear de	
35 of hypoglycaemia. <u>The forest plot of each analysis presents the subgroups which were</u>	<u>e</u>
36 compared using the Cochrane O test and the I ₂ index.[38] Moreover, sensitivity analy	ses were
37 conducted by excluding studies with a high risk of bias (i.e. at least one item), or studies	lies
 allowing the use of insulin, or studies for which one or more patients characteristics w 	vere
 39 Imbalanced among groups. Publication bias was evaluated <u>by</u> using a funnel plot and Egger's regression test. 	(p<0.05
40	<u>harmful</u>
considered significant).[40] The number of patients needed to be treated to observe a	
41 autoome (Number Needed to Harm (NNH) Treat for one parson to be Harmed NNT()	H)) was
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19	estimated according to the Cochrane recommendations.[41] The Assumed Control Risk (ACR) of hypoglycaemia in SU-treated patients was calculated for eachfrom a meta-analysis
20	reported by Hemmingsen <i>et al.</i> that included 27 clinical trials from which the incidence of
21	hypoglycaemia was calculated.[26] On the assumption that the prevalence of hypoglycaemia
22	is related to the length of follow-up, different follow-up scenarios were created: any duration
23	(ACR 19.9%, 23 studies), <6 months (ACR 11.6%, 7 studies), from 6.1 to 12 months (ACR:
24	13.3%, 9 studies), more than 12 months (ACR 22.8%, 11 studies) (see eTable 2 in the
25	Supplement for study and pooled in a forest plot.[29] The details).
26	The analyses were conducted using with Review Manager software (RevMan version 5.3, The
27	Nordic Cochrane Centre, The Cochrane Collaboration) and R software (version 2.15.3).
28	All relevant aspects related to search strategy, study selection, data extraction and quality
29	assessment, and data analysis were specified in a synopsis protocol specifyingdetailing the
30	meta-analysis objective and context, and the principles and modalities of the literature search
31	and the data analysis waswere developed.
32 33	
	RESULTS Study selection
34 35	The literature search identified 2,379 records from the literature databases used, 687 of which
36	were duplicates and were thus removed. Eleven records were retrieved through other sources.
30 37	Thus, the title and abstract of 1,708 individual study records were assessed, 1,650 of which
38	were found to be irrelevant and were excluded. The remaining 57 records underwent full-text
30 39	examination (results detailed in the supplement); tenSupplement); 10 were finally included in
39 40	this meta-analysis (Figure 1).[14, 30-38<u>24, 42-50]</u>
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41	Study characteristics
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19	The 10 selected RCTs included a total of 6,456 patients, of whom 4,020 received DPP4-i +	
20	SU, and 2,526 PBO + SU. All studies were randomized, and used double-blind procedures;	
20	the. The study reported by Barnett <i>et al.</i> included only patients aged \geq 70 years.[1424] The	
22	planned follow-up of the included studies ranged from 12 to 76 weeks. The associated SU	
23	varied across the selected RCTs (Table 1); drug). Drug therapy also included metformin in	
23 24	four RCTs; [14, 32, 35, 38] use.[24, 44, 47, 50] Use of insulin was allowed in two RCTs.[14 ,	
24 25	3824, 50] Baseline key patient characteristics (namely mean (standard deviation) glycated	
25 26	haemoglobin A1C (HbA1c) of, mean age, and gender) were well balanced among the patients	
20 27	included in theseeach group of included RCTs-ranged from 7.8% (0.8) to 8.6% (0.8, with the	
	exception of two studies[24, 46] in which there was a notable difference in sex ratio between	
28	the groups (Table 1).	
29	Three RCTs studied in-linagliptin 5 mg/day; for a total of 1,038 patients.[14, 34, 3524, 46, 47]	
30	Vildagliptin 100 mg/day was studied in two RCTs, [31, 3343, 45] and vildagliptin 50 mg/day	
31	in one,[31[43] for a total of 271 patients with 100 mg/day, and 170 with 50 mg/day.	
32	Alogliptin was studied once at 12.5 mg/day and once at 25 mg/day; [36, 37[48, 49] for a total	
33	of 308 patients with 12.5 mg/day; and 302 with 25 mg/day. White et al. studied alogliptin at	
34	different doses (from 6.5 mg/day to 25 mg/day) in 1,198 patients receiving SU.[3850]	
35	Saxagliptin (248 patients with 2.5 mg/day, and 253 with 5 mg/day)[3042] and sitagliptin 100	
36	mg/day (222 patients)[3244] were each studied once. Overall, a total of 2,526 patients	
37	receiving PBO + SU were identified in the included RCTs (Table 1).	
38	Six of the ten included RCTs did not clearly report the definition of hypoglycaemia (Table	
39	2):[30, 32, 34, 35, 37, 381).[42, 44, 46, 47, 49, 50] There was a high risk of reporting bias in	
40	three of the included studies, [31-33] one. [43-45] One RCT also presented a high risk of	
41	detection bias (Figure 2).[3244]	
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Overall, 4,020 patients received DPP4-i (2,096 at full dose, 726 at low dose, and 1,198 at undefined dose) + SU, of whom 479 patients developed hypoglycaemia (311 at full dose, 67 at low dose, and 101 at undefined dose) corresponding to an absolute risk of 11.9%; 2,526 received PBO + SU, of whom 169 developed hypoglycaemia, corresponding to an absolute risk of 6.7%.

Meta-analysis

The RR of hypoglycaemia for DPP4-i any dose + SU versus PBO + SU was 1.52 (95%CI 1.29 to 1.80), with no evidence of heterogeneity across RCTs (Q = 11.2, p = 0.26, $I^2 = 20\%$; Figure 3). The correspondent NNHFor any DPP4-i +SU duration of use, the corresponding NNT(H) was 10 (6 to 17); it was 17 (11 to 30) for a treatment duration <6 months, 15 (9 to 26-9 (19-) for 6.1 to 12 months, and 8 (5 to 43.3; Figure 4)-15) for a treatment duration longer than one year. The pooled RR did not markedly change when RCTs with a high risk of detection bias and reporting bias (1.40; 1.18 to 1.67; eFigure 1 in the supplementSupplement), or when the RCTs which allowed the use of insulin (1.61; 1.30 to 2.00), were excluded from the analysis. The RR was similar to that of the principal analysis when RCTs in which a notable imbalance in sex ratio were excluded (1.52; 1.27 to 1.81; Q = 10.70, p = 0.15; $I^2 = 35\%$; eFigure 2 in the supplement). The pooled RR was also similar for RCTs in which a definition of hypoglycaemia was reported (1.54; 0.99 to 2.42; Q = 2.1, p = 0.5, $I^2 = 0\%$), and for in those in which a definition was not reported (1.52; 1.27 to 1.82; Q = 9.1, p = 0.10, I² = 45%), without any evidence of heterogeneity between these two groups (Q = 0.0, p = 0.95, $I^2 = 0\%$; eFigure ent). According to the dose of DPP4-i evaluated, the RR of hypoglycaemia nificantly increased for DPP4-i full dose (1.66; 1.34 to 2.06), but not for DPP4-i tes (1.33; 0.92 to 1.94; Figure 5). The NNH for DPP4-i full dose was 19.4 (13.9 to

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19	32-2; Figure 6). Funnel plot did not show clear evidence of publication bias, (Figure 7) and the Egger test did not find asymmetry (z=1.3; p=0.2). The strength of evidence of this meta-	
20 21	analysis was evaluated as high with regards to the GRADE framework (Table 33 in the	
22	supplement).	Formatted: Font: Not Bold
23	According to the dose of DPP4-i evaluated, the subgroup analysis showed no difference between low and full DPP4-i dose with regard to the risk of hypoglycaemia ($Q = 0.99$, $p =$	
24 25	0.32 , $I^2 = 0\%$; Figure 4). The risk remained significantly increased for DPP4-i full dose (1.66)	
26	1.34 to 2.06) but was not significantly increased for DPP4-i low doses (1.33; 0.92 to 1.94; Figure 5). For DPP4-i full dose+SU, the NNT(H) was 8 (5 to 15) for any treatment duration;	
27	rigue 5, rol Dre4-1 un dosers0, the NN1(1) was (310-13) for any treatment duration, it was 13 (8 to 25) for a treatment duration <6 months, 11 (7 to 22) for a treatment duration	
28 29	between 6.1 to 12 months, and 7 (4 to 13) for a treatment duration longer than one year.	
29 30	Visual inspection of the funnel plot did not show any clear evidence of publication bias (Figure 5), and the Egger test did not find any asymmetry (z=1.3; p=0.2). The strength of	
31	evidence of this meta-analysis was evaluated as high with regards to the GRADE framework	
32 33	(Table 2).	
34	DISCUSSION	
35	Principal findings	
36 37	This meta-analysis found <u>about</u> a 50% increase in the risk of hypoglycaemia when DPP4-i and SU were associated in Type II diabetic patients, leading to one supplementary case of	
38	hypoglycaemia for every 2710 treated patients. This risk was confirmed for full doses of	
39	DPP4-i, while it emmetcould not be excluded for lower doses. DPP4-i act indirectly on insulin levels by enforcing the incretin effect, which is a response to	
40 41	high oral intake of carbohydrates and fatty acids. [917] Such drugs should therefore act on	
42	glycaemia only in response to such intakes, thereby protecting patients from hypoglycaemia.	
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However, in patients treated with SU, insulin secretion is already stimulated independently of

glycaemia and the addition of a reinforced incretin effect on insulin levels leads to an increase

ashuge number of cases of induced hypoglycaemiashypoglycaemia, some of which

in the risk of hypoglycaemia. Given the frequency of this event in Type II diabetic patients

could be severe. [3951] The present meta-analysis did not allow the investigation as toof the

treated with SU, the risk associated with the addition of DPP4-i would lead to a

threshold of dose combination (DPP4-i + SU) associated with an increased risk of

The risk of hypoglycaemia related to the addition of a DPP4-i to SU is, however,

SU dose reduction has not been investigated in trials studying DPP4-i.

hypoglycaemia; an individual patient meta-analysis could be helpful in this regard. This

acknowledged in the SmPCs for DPP4-i; most recommend using full-dose DPP4-i but a

not stated.[16-2027-31] Currently, to what extent this recommendation would lower the

reduced SU dose in patients taking such combinations, although the magnitude of reduction is

number of excess cases of induced hypoglycaemia is unknown. Thelt is also of note that the

suggested individual patient meta-analysis would not fill this knowledge gap as the effect of

For low doses of DPP4-i (half the full-dose when applicable), the increase in hypoglycaemia

risk was not significant. The However, the existence of this risk cannot however be fully

ruled_out by the present results; and a larger sample would be required to increase the precision of the estimates. Furthermore, although the point estimate was lower (RR 1.33 vs.

1.66 for full-doses), which suggests a potential dose-effect, no heterogeneity was found

between low- and full-doses of DPP4-i-, yet this could result from a lack of power in the

heterogeneity test (low-dose group was half the size of the high-dose group).

Strengths and limitations of study

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19	The present analysis has important strengths. Firstly, it is based on <u>a</u> large sample of patients;
20	over 4,000 treated with a combination of DPP4-i and SU, and over 2,500 treated with PBO
21	and SU. Secondly, the overall quality of the included studies appeared seems high overall
22	according to the Cochrane Collaboration's Collaboration tool for risk of bias assessment,
23	which was confirmed by the GRADE framework evaluation of the meta-analysis that
24	considers that the strength of evidence provided is high. The present meta-analysis used data
25	concerning all currently marketed DPP4-i (alogliptin, linagliptin, saxagliptin, sitagliptin, and
26	vildagliptin), and results were consistent within studies with no heterogeneity being found
27	among estimates. Thirdly, there was no evidence of publication bias; the funnel plot was
28	balanced and the Egger test was not significant. The
29	Nevertheless, the meta-analysis does, however, have certain limitations. Firstly, certain
30	studies that presented witha high risk for detection and reporting bias risk of bias were
31	included in the main analysis, [31-3343-45] but exclusion of these fromstudies did not change
32	the estimates significantly. Secondly, three studies could not be included as data were not available for the risk of hypoglycaemia in patients underreceiving SU[40-4252-54] However,
33	in view of the GRADE framework, including results from these studies would be unlikely to
33 34	significantly change the results significantly owing to the size of the present meta-analysis,
34 35	the high number of hypoglycaemia cases, and the confidence intervals of the pooled RR that
36	clearly do not cross the line of no effect. [2233] The absence of heterogeneity in estimates
30 37	found from the ten10 included studies further supports this hypothesis. Thirdly, the results of
37 38	this meta-analysis are dominated by the results of three studies that account for more than
	80% of the pooled results of the principal analysis: [42, 47, 50] a sensitivity analysis without
39 40	these studies did not substantially change the results of the meta-analysis (data not shown).
40 41	Fourthly, the definition of hypoglycaemia varied among the included RCTs; and was not
	reported in five. Other authors havedid not performedperform a meta-analysis on
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19	hypoglycaemia risk on the basis of this lack of homogeneity in its definition across the
20	RCTs;[1220] nevertheless, this could be considered as a minor limitation, as in the present
20	analysis the risk did not differ between RCTs with or without a clear definition of
22	hypoglycaemia. The incidence of hypoglycaemia also differed among studies, mainly because
22	of different durations of follow-up. However, this did not have any impact on the estimation
23 24	of the pooled risk (no statistical heterogeneity was found) nor on the NNT(H) calculation.
24 25	which was based on an external Assumed Control Risk of hypoglycaemia retrieved from 27
	clinical studies included in a meta-analysis of the Cochrane library.[26]
26	
27	Clinical importance
28	It is important to underline that hypoglycaemia is the most frequent adverse reaction related to
29	anti-diabetic treatments and that, even when not directly life-threatening, it is associated with
30	an increased risk of all-cause mortality, cardiovascular disease, and cardiovascular mortality
31	and hospital admission.[2, 3, 436] In addition, it should not be neglected that these frequent
32	eventshypoglycaemia and theirits related symptoms (e.g. nervousness, sweating, trembling,
33	weakness, palpitations) impact negatively on patient quality of life, and disrupt many daily
34	activities such as driving, work performance, and leisure pursuits.[44, 45] The risk herein
35	demonstrated for all-type hypoglycaemia should thus and disrupt many daily activities such as
36	driving, work performance and leisure activities.[9, 10] More importantly, mild-to-moderate
37	iatrogenic hypoglycaemia can decrease the usual adrenergic response to hypoglycaemia.[12]
38	This may cause hypoglycaemia unawareness and compromise behavioural defences (hunger
39	resulting in sugar ingestion), which in turn can lead to severe hypoglycaemia.[13, 14] It is
40	thus an important to lower the risk of mild-to-moderate hypoglycaemia, which remains a
41	serious adverse event. Adequate information regarding the risk of hypoglycaemia, whatever its severity, should thus be considered of primary importance for patients and all health
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19	professionals involved in the management of diabetic patients. Reaching good glycated	
20	haemoglobin levels should not be at the expense of hypoglycaemic events, which could	
20	outweigh the benefit of preventing risks associated with elevated blood glucose	
21	concentrations. Thus, the risk demonstrated herein for all-type hypoglycaemia should not be	
22	minimized by considering that only severe episodes would be of clinical concern.	
23		
24	Conclusions	
25	In conclusion, this meta-analysis found about a 50% increase in the risk of hypoglycaemia	
26	associated with the addition of DPP4-i to SU in patients with type II diabetes. For this adverse	
27	event commonly experienced by treated diabetic patients, this would lead to the occurrence of	
28	one supplementary hypoglycaemic event of hypoglycaemia forin every 2710 treated patients.	
29	This has the potential to represent potentially represents a tremendous huge number of	
30	attributable cases worldwide. These results clearly highlight the need to respect existing	
31	recommendations for SU dose reduction when initiating a DPP4-i treatment, and the urgency	
32	to determine the efficacy of this measure in minimizing the risk of hypoglycaemia.	
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to the EXAMINE trial.

the study have been omitted.

Authors' contributions

investigated and resolved.

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The corresponding author had full access to all the data in the study and takes responsibility

Francesco Salvo (corresponding author) and Antoine Pariente (manuscript's guarantor) affirm

that the manuscript is an honest, accurate, and transparent, and that no important aspects of

FS: conception and design; acquisition, analysis and interpretation of data; drafting the article

and revising it critically for important intellectual content; final approval of the version to be

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Declaration of interests	
The corresponding author ensures that the manuscript is complete and that the conflict	of

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have noany financial relationships that may be relevant to the submitted work;, and that they

Ethical approval

This type of study does not require ethical approval.

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completed the Unified Competing Interest form at www.icmje.org/coi_disclosure.pdf

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Data sharing

No additional data available.



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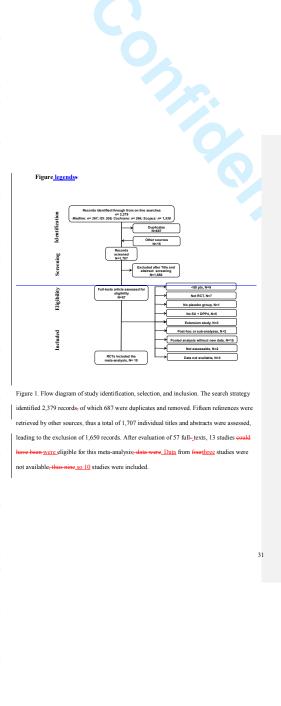
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Random sequence generation (selection bias) Allocation concealment (selection bias) Blinding of participants and personnel (performance bias) Blinding of outcome assessment (detection bias) Incomplete outcome data (attrition bias) Selective reporting (reporting bias) Other bias 50% 0% Low risk of bias Unclear risk of bias High risk of bias Figure 2. Risk of bias graph. Review authors' judgments for each 'Risk of bias' item presented as percentages across all included studies. The risk of bias of the included studies is presented in different colorscolours: green represents a low risk of bias, red represents a high risk of bias, yellow representrepresents an unclear risk of bias.

75%



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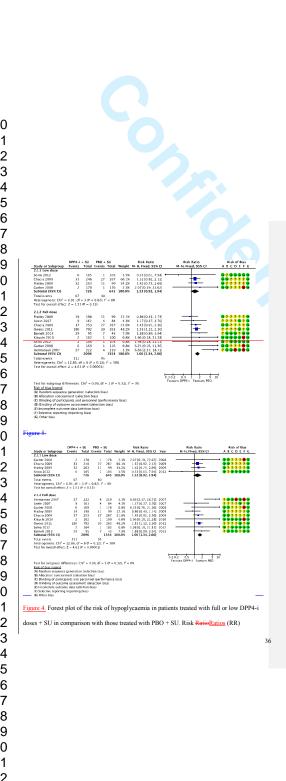
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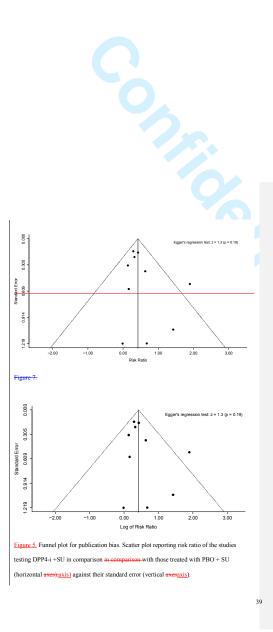




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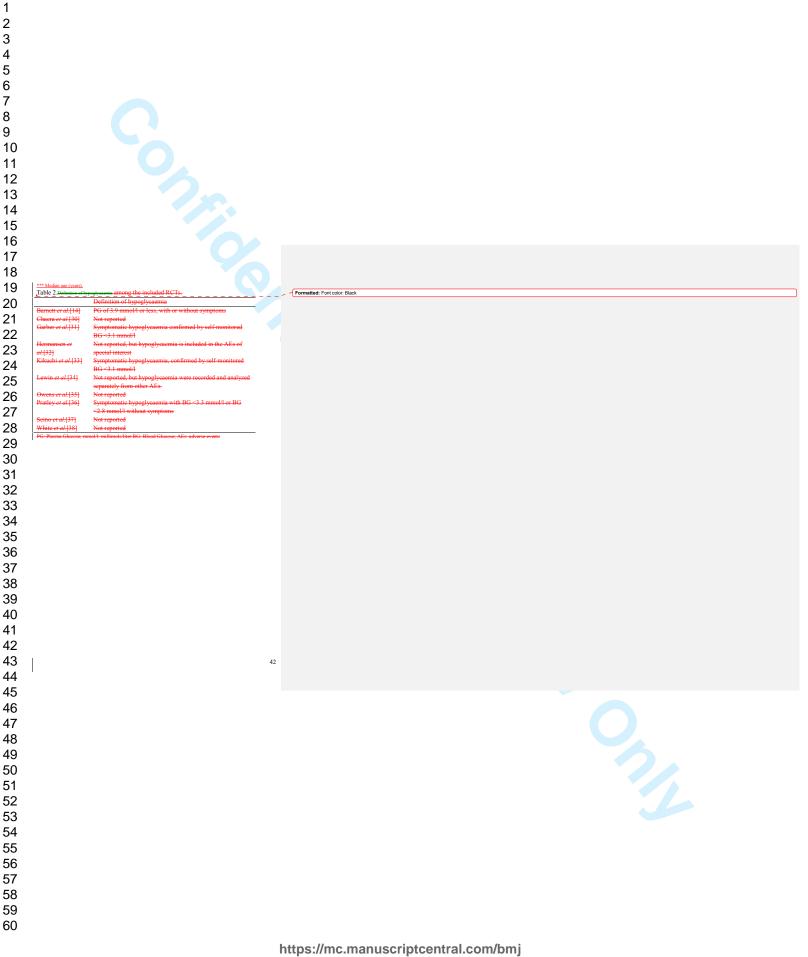
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19	STUDY NNH (85%CI) Weight, %
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21	Chara 2009 222 (WH 770.9) 15.7
22	Total (85%Ci) 19.4 (13.9; 32.2) 100.0
23 24	NNT NNH
24 25	Figure 6. Forest plot of the Number Needed to Harm (NNH) of hypoglycaemia in patients
26	treated with full DPP4-i dose + SU in comparison with those treated with PBO + SU. NNH
27	calculated for individual randomized controlled trials (RCTs) with 95% confidence intervals
28	(CI) are presented; arrows indicate the CI exceeding the limits of the graph. Protective
29	estimates (or CI) are reported as Number Needed to Treat (NNT, left side of the forest plot). An estimate of the weight of each RCT on overall NNH is reported as a percentage and
30	graphically (black square size). Overall NNH is also presented (black diamond).
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14		Table 1. Stu	udy characteristi	cs				le l		[5]
15				Intervention, daily dose (n) Associated SU	Mean HbA _{1c} at baseline, <u>%</u>	Mean age of	Male, %	Definition of hypoglycaemia		
16						participants-(, years)		/		[7] [9]
17 ^B	arnett	t <i>et al.</i> [<mark>1424]</mark>	24	Linagliptin 5 mg (95 pts) or PBO (43 pts) SU, not specified	<u>DPP4-i: 78%</u>	DPP4-i: 75 PBO: 75*	DPP4-i: 72 PBO: 62*	PG of 3.9 mmol/l or less, with or without symptoms	Formatted	[10]
				Saxagliptin 2.5 mg (248 pts), saxagliptin 5 Glyburide mg (253 pts), or PBO (267 pts)	PBO: 7.7* DPP4-i: 8.4%*-8.5	DPP4-i: 55	DPP4-i: 45	Not reported		[11]
10				mg (253 pts), or PBO (267 pts)	<u>8.4%*-8.5</u> PBO: 8.4			==========		[12]
20 6	iarber	et al.[<mark>3143]</mark>	24	Vildagliptin 50 mg (170 pts) or 100 mg Glimepiride (169 pts), or PBO (176 pts)	<u>DPP4-i-</u> 8.5%	<u>5800004-i: 58-59</u>	DPP4-i: 59 PBO: 58	Symptomatic hypoglycaemia	Formatted	[14]
20					<u>8.6</u> PBO: 8.5	<u>PBO: 58</u>		<u>BG <3.1 mmol/1</u>		[15]
	lerman 1.[<mark>32</mark>	nsen_et	24	Sitagliptin 100 mg (222 pts) or PBO (219 Glimepiride pts)	<u>DPP4-i: 8 3%</u> PBO: 8 3	NRDPP4-i: 56	DPP4-i: 53 PBO: 53	Not reported, but hypoglycaemia is included in		[17]
				Vildagliptin 100 mg (102 pts) or PBO Glimepiride	7.0%DPP4.i			the AEs of special interest Symptomatic hypoglycaemia,		[18]
23 6	IKUGI			Vildagliptin 100 mg (102 pts) or PBO Glimepiride (100 pts)	7.8	DPP4-i: 59 PBO: 60	<u>PBO: 69</u>			[19]
24	ewin	et al.[<mark>34<u>46]</u></mark>	18	(100 pts) Linagliptin 5 mg(161 pts) or PBO (84SU, not specified pts)	<u>PBO: 8.0</u> <u>DPP4-i:</u> 8.6%	DPP4-i: 57	DPP4-i: 48	Not reported, but	Formatted	[21]
25				pts)	<u>PBO: 8.6</u>	<u>PBO: 56</u>	<u>PBO: 62</u>	hypoglycaemia were recorded and analyzed separately from		[22]
26	wens	et al [35 47]	24	Linagliptin 5 mg (792 pts) or PBO (263 SU, not specified	DPP4-i: 8 1%	DPP4-i: 58	DPP4-i: 48	other AEs. Not reported		[23]
27 *	- 1		26	Lingliptin 5 mg (792 pts) or PBO (263SU, not specified pts)	PBO: 8.1	PBO: 58	PBO: 47		Formatted	[25]
28 P	ratley	et al. [<mark>3648]</mark>	20	Alogliptin 12.5 mg (203 pts), alogliptin 25 Glyburide mg (198 pts), or PBO (99 pts)	NK	<u>+ + DPP4_i: 56.5</u> <u>PBO: 57</u>	PBO: 51.5	Symptomatic hypoglycaemia with BG <3.3 mmol/1 or BG <2.8 mmol/1 without symptoms		[26]
29 30 ^{&}	eino e	et al.[<mark>37<u>49</u>]</mark>	12	Alogliptin 12.5 mg (105 pts), alogliptin 25 Glimepiride mg (104 pts), or PBO (103 pts)	DPP4-i:		<u>DPP4-i: 66</u>	<2.8 mmol/l without symptoms Not reported		[28]
				mg (104 pts), or PBO (103 pts)	<u>DPP4-i:</u> <u>8.5%**%</u> PBO: 8.6%	<u>PBO: 60</u>	_ <u>PBO: 69</u>			[29]
31	/hite e	et_al.[<mark>38</mark> 50]	76****	Alogliptin any doses (1,198), or PBOSU_ not specified (1,172 pts)	<u>DPP4-i:</u> 80%	<u>61_DPP4=i: 61</u>	DPP4-i: 68 PBO: 69*	<u>Not reported</u>		[30]
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13	PBQ: 61*** [HbAic Glycated hemoglobin Air; NR: not reported; PBQ: placebo; Pts; patients; SU; sulphonylureas; y=years old, PG: Plasma Glucose; mmol/1: millimols/liter BG: Blood Glucose; AEs:	
14	Adverse events_ *Data for group with saxagliptin 2.5 mg/day and in placebo group; in saxagliptin 5 mg/day HbA _L was 8-5%.	Formatted: Font: 9 pt
15	** Data for group with alogliptin; in placebo group HbA _{1c} was 8-6%.	
16	**** Data refer to overall study population, not only to SU treated patients. **_Median exposure weeks for alogliptin treated patients.	- Formatted: Font: 9 pt
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14		-	Quality asso	-			Nº of pa	atients	Ei	fect		
15 _{№ of} 16 ^{studies}	Study design	Risk of bias	Inconsistency	Indirectness	Imprecision	Other considerations	All studies	Placebo	Relative (95% CI)	Absolute (95% CI)	Quality	Importance
17Hypogly	vcaemia	I										
18 ₁₀ 19 20 21	randomized trials	not serious ¹	not serious ²	not serious	not serious ³	dose response gradient	479/4,020 (11.9%)	169/2,526 (6.7%)	RR 1.52 (1.29 to 1.80)	35 more per 1,000 (from 19 more to 54 more)	⊕⊕⊕⊕ HIGH	CRITICAL ⁴
22 23 24 25 26 27 28 29 30 31 32 33 34 35 36 37 38 39 40 41	2. No he 3. The sa (lower 4. Hypog	were exclus terogeneity a mple size is bound of 95 lycaemia is	led from the analy mong estimates w large (n=6,526), tl %CI = 1.29) the most frequent :	visis the result die as found. he number of the adverse reaction	d not change sub e events high (64 n related to anti-c	Among them, a high stantially. 18), and the confider liabetic treatment. It nbling, weakness, p	nce intervals of t	the pooled RR	t clearly <u>do</u> no tescause morta	ot cross the line lity and-of card	of no effect	

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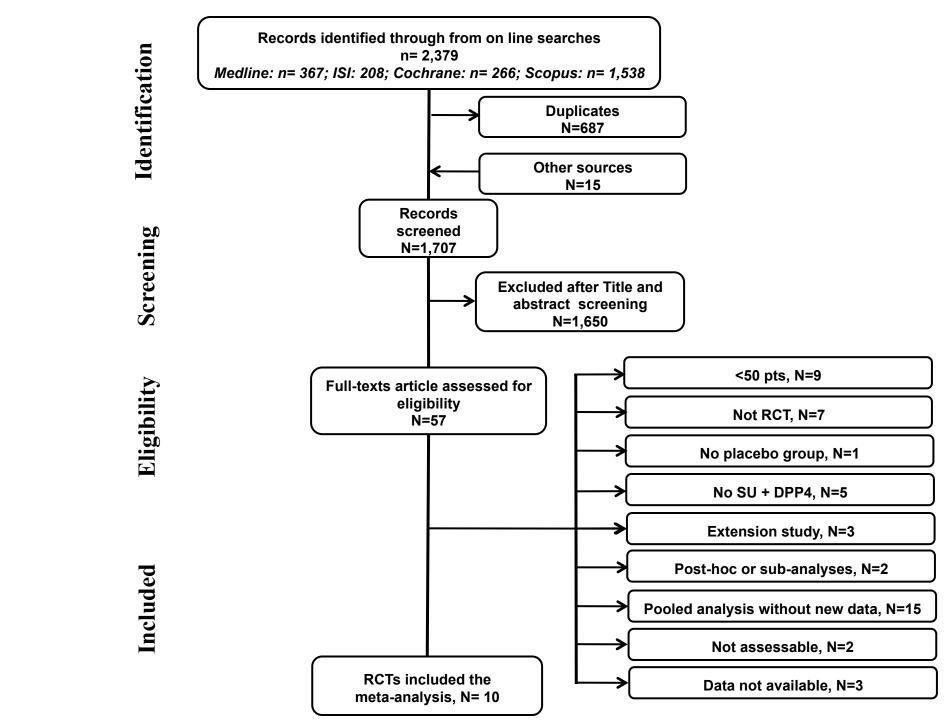
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22 23	Random sequence generation (selection bias)					
24 25	Allocation concealment (selection bias))				
26 27	Blinding of participants and personnel (performance bias)					
28 29	Blinding of outcome assessment (detection bias)					
30 31	Incomplete outcome data (attrition bias)					
32	Selective reporting (reporting bias)					
33 34	Other bias					
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19 20		DPP4-i		PBO +			Risk Ratio		Risk Ratio	Risk of Bias
20	Study or Subgroup	Events					M-H, Fixed, 95% Cl			ABCDEFG
21	Hermansen 2007	27	222	4	219	1.9%	6.66 [2.37, 18.71]			?
	Lewin 2007	9	161	4	84	2.5%	1.17 [0.37, 3.70]			???? ?+ ? +
23	Garber 2008 Chacra 2009	8 70	339 501	1 27	176 267	0.6% 16.9%	4.15 [0.52, 32.94]			
24	Pratley 2009	51	401	11	207	8.4%	1.38 [0.91, 2.10] 1.14 [0.62, 2.11]			
25	Kikuchi 2010	2	102	1	100	0.5%	1.96 [0.18, 21.28]			
26	Owens 2011	180	792	39	263	28.0%	1.53 [1.12, 2.10]			????+?+
27	Seino 2012	2	209	1	103	0.6%	0.99 [0.09, 10.74]		← →	
28	Barnett 2013	29	95	7	43	4.6%	1.88 [0.89, 3.94]			? • • • ? • •
29	White 2013	101	1198	74	1172	35.8%	1.34 [1.00, 1.78]	2013	⊢∎ -	+ + + + + ? +
30 31	Total (95% CI)		4020		2526	100.0%	1.52 [1.29, 1.80]		•	
32	Total events	479		169						
	Heterogeneity: Chi ² =				$1^2 = 20$	%			0.10.2 0.5 1 2 5 10	
33	Test for overall effect:	: Z = 4.90	(P < 0.0)	00001)					Favours DPP4-i Favours PBO	
34	Dist. (11)									
35	Risk of bias legend (A) Random sequence	aonoratio	on (color	tion his	-)					
36	(B) Allocation conceal				5)					
37	(C) Blinding of particip				ormanc	e bias)				
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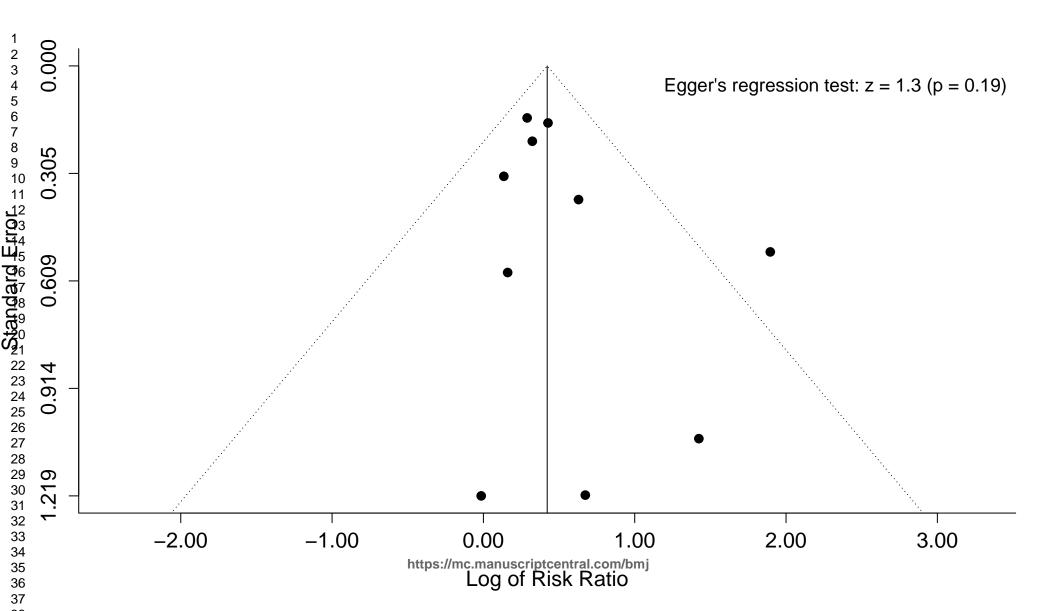
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15	Study or Subgroup	DPP4-i Events		PBO +		Waight	Risk Ratio M-H, Fixed, 95% Cl	Voar	Risk Ratio M-H, Fixed, 95% Cl
16-	2.1.1 Low dose	Events	TOLAT	Events	TOLAI	weight	M-H, Fixed, 95% CI	rear	
17	Garber 2008	2	170	1	176	2.3%	2.07 [0.19, 22.62]		
18 19	Chacra 2009 Pratley 2009	33 32	248 203	27 11	267 99	60.1% 34.2%	1.32 [0.82, 2.12] 1.42 [0.75, 2.69]		
20	Seino 2012	0	105	1	103	3.5%	0.33 [0.01, 7.94]		← − − − − − − − − − − − − − − − − − − −
21	Subtotal (95% CI) Total events	67	726	40	645	100.0%	1.33 [0.92, 1.94]		
22	Heterogeneity: $Chi^2 =$		= 3 (P =		$^{2} = 0\%$				
23	Test for overall effect	: Z = 1.51	(P=0.	13)					
24 25	2.1.2 Full dose								
25 26	Hermansen 2007	27	222	4	219	3.3%	6.66 [2.37, 18.71]		·
27	Lewin 2007 Garber 2008	9 6	161 169	4 1	84 176	4.3% 0.8%	1.17 [0.37, 3.70] 6.25 [0.76, 51.36]		
28	Pratley 2009	19	198	11	99	12.1%	0.86 [0.43, 1.74]		
29	Chacra 2009	37	253	27	267	21.6%	1.45 [0.91, 2.30]		
30	Kikuchi 2010 Owens 2011	2 180	102 792	1 39	100 263	0.8% 48.2%	1.96 [0.18, 21.28] 1.53 [1.12, 2.10]		
31 32	Seino 2012	2	104	1	103	0.8%	1.98 [0.18, 21.51]	2012	
33	Barnett 2013 Subtotal (95% CI)	29	95 2096	7		7.9% 100.0%	1.88 [0.89, 3.94] 1.66 [1.34, 2.06]	2013	•
34	Total events	311		95			,		-
35	Heterogeneity: Chi ² = Test for overall effect	12.86, df	F = 8 (P)	= 0.12);	$l^2 = 38$	3%			
36	rest for overall effect	. 2 – 4.01	(r < 0.	00001)					
37 38									0.1 0.2 0.5 1 2 5 10
	Test for subgroup dif	ferences: (Chi ² = ().99, df :	= 1 (P =	= 0.32), I ²	= 0%		Favours DPP4-i Favours PBO
40	Risk of bias legend								
41	(A) Random sequence (B) Allocation conceal				s)				
42	(C) Blinding of particip				ormanc	e bias)			
43 44	(D) Blinding of outcor(E) Incomplete outcon				bias)				
45	(F) Selective reporting			DIdS)					
46	(G) Other bias								
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20 21		DPP-4i		PBO +			Risk Ratio		Risk Ratio	Risk of Bias
22-	Study or Subgroup	Events	Total	Events	Total	-	M-H, Fixed, 95% Cl			ABCDEFG
23 24	Lewin 2007 Chacra 2009	9 70	161 501	4 27	84 267	2.6% 17.4%	1.17 [0.37, 3.70] 1.38 [0.91, 2.10]	2009		????? ? ? + ? + ++ ????? +
24 25	Pratley 2009 Owens 2011	51 180	401 792	11 39	99 263	8.7% 28.9%	1.14 [0.62, 2.11] 1.53 [1.12, 2.10]			
26	Seino 2012	2	209	1	103	0.7%	0.99 [0.09, 10.74]	2012		•?••••
27 28	White 2013 Barnett 2013	101 29	1198 95	74 7	1172 43	36.9% 4.8%	1.34 [1.00, 1.78] 1.88 [0.89, 3.94]			
28 29	Total (95% CI)									
30	Total events	442	3357	163		100.0%	1.40 [1.18, 1.67]			
31	Heterogeneity: Chi ² =	1.60, df		0.95); l ²	² = 0%					
32 33	Test for overall effect:	Z = 3.82	P = 0.0)001)					Favours DPP4-i Favours PBO	
34	Risk of bias legend									
35	(A) Random sequence(B) Allocation conceal				5)					
36 37	(C) Blinding of particip (D) Blinding of outcon					e bias)				
38	(E) Incomplete outcom	ne data (a	ttrition l		DIAS)					
39	(F) Selective reporting(G) Other bias	(reportin	g bias)							
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18 19									
20		DPP4-i		PBO +			Risk Ratio		Risk Ratio Risk of Bias
21	Study or Subgroup	Events					M-H, Fixed, 95% Cl		
22 ⁻ 23	Hermansen 2007	27	222	4	219	2.1%	6.66 [2.37, 18.71]		
	Garber 2008 Chacra 2009	8 70	339 501	1 27	176 267	0.7% 18.2%	4.15 [0.52, 32.94]		
24	Pratley 2009	51	401	11	207	9.1%	1.38 [0.91, 2.10] 1.14 [0.62, 2.11]		
25	Kikuchi 2010	2	102	1	100	0.5%	1.96 [0.18, 21.28]		
26	Owens 2011	180	792	39	263	30.2%	1.53 [1.12, 2.10]		
27	Seino 2012	2	209	1	103	0.7%	0.99 [0.09, 10.74]		$\longleftarrow \qquad \qquad$
28	White 2013	101	1198	74	1172	38.6%	1.34 [1.00, 1.78]	2013	╺┓┓┓┓
29	Total (95% CI)		3764		2200	100.0%	1.52 [1.27, 1.81]		
30	Total events	441	5704	158	2333	100.070	1.52 [1.27, 1.01]		•
31	Heterogeneity: $Chi^2 =$		= 7 (P)		$1^2 = 35$	%			
32	Test for overall effect				1 - 55	/0			0.1 0.2 0.5 1 2 5 10 5 5 10
33			(, , , , , , , , , , , , , , , , , , ,	,					Favours DPP4-i Favours PBO
34	<u>Risk of bias legend</u>								
35	(A) Random sequence				5)				
36	(B) Allocation conceal								
37	(C) Blinding of partici					e bias)			
~~	(D) Blinding of outcor(E) Incomplete outcon				bias)				
	(F) Selective reporting			DIAS)					
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16		DPP4-i		PBO +			Risk Ratio		Risk Ratio	Risk of Bias
17-	Study or Subgroup 3.1.1 Hypoglycaemia	Events defined	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	Year	M–H, Fixed, 95% Cl	ABCDEFG
18	Garber 2008	8	339	1	176	4.4%	4.15 [0.52, 32.94]			••???••
19 20	Pratley 2009 Kikuchi 2010	51 2	401 102	11 1	99 100	59.6% 3.4%	1.14 [0.62, 2.11] 1.96 [0.18, 21.28]			
21	Barnett 2013	29	95	7	43	32.6%	1.88 [0.89, 3.94]	2013		? + + + ? + +
22	Subtotal (95% CI) Total events	90	937	20	418	100.0%	1.54 [0.99, 2.42]			
23	Heterogeneity: Chi ² =		= 3 (P =		= 0%					
24 25	Test for overall effect:	Z = 1.90	(P = 0.)	06)						
26	3.1.2 Hypoglycaemia	not defi	ned							
27	Hermansen 2007	27	222	4	219	2.2%	6.66 [2.37, 18.71]			
28	Lewin 2007 Chacra 2009	9 70	161 501	4 27	84 267	2.9% 19.7%	1.17 [0.37, 3.70] 1.38 [0.91, 2.10]			
29 30	Owens 2011	180	792	39	263	32.7%	1.53 [1.12, 2.10]	2011		????+?+
31	Seino 2012 White 2013	2 101	209 1198	1 74	103 1172	0.7% 41.7%	0.99 [0.09, 10.74] 1.34 [1.00, 1.78]			
32	Subtotal (95% CI)		3083			100.0%	1.52 [1.27, 1.82]		•	
33	Total events Heterogeneity: Chi ² =	389 9 15 df	= 5 (P =	149 0 10) [,] 1 ²	= 45%	Ś				
34	Test for overall effect:				15/					
35 36										
37									0.1 0.2 0.5 1 2 5 10 Favours DPP4-i Favours PBO	
38	Test for subgroup diff <u>Risk of bias legend</u>	erences:	$Chi^2 = C$).00, df =	= 1 (P =	= 0.95), l ²	= 0%			
39 40	(A) Random sequence				5)					
40	(B) Allocation concealr (C) Blinding of particip				rmanc	e bias)				
42	(D) Blinding of outcom					e blas)				
43	(E) Incomplete outcom			bias)						
44 45	(F) Selective reporting(G) Other bias	(reportin	y Dias)							
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