Infection fatality risk for SARS-CoV-2 in community dwelling population of Spain: nationwide seroepidemiological study

Roberto Pastor-Barriuso,1,2,* Beatriz Pérez-Gómez,1,2* Miguel A Hernán,3 Mayte Pérez-Olmeda,4 Raquel Yotti,5 Jesús Oteo-Iglesias,4,6 Jose L Sanmartín,7 Inmaculada León-Gómez,1,2 Aurora Fernández-García,2,4 Pablo Fernández-Navarro,1,2 Israel Cruz,8 Mariano Martín,7 Concepción Delgado-Sanz,1,1 Nerea Fernández de Larrea,1,2 Jose León Paniagua,5 Juan F Muñoz-Montalvo,7 Faustino Blanco,2 Amparo Larrauri,1,7 Marina Pollán,1,2† on behalf of the ENE-COVID Study Group

ABSTRACT

OBJECTIVE

To estimate the infection fatality risk for severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2), based on deaths with confirmed coronavirus disease 2019 (covid-19) and excess deaths from all causes.

DESIGN

Nationwide seroepidemiological study.

SETTING

First wave of covid-19 pandemic in Spain.

PARTICIPANTS

Community dwelling individuals of all ages.

MAIN OUTCOME MEASURES

The main outcome measure was overall, and age and sex specific, infection fatality risk for SARS-CoV-2 (the number of covid-19 deaths and excess deaths divided by the estimated number of SARS-CoV-2 infections) in the community dwelling Spanish population. Deaths with laboratory confirmed covid-19 were obtained from the National Epidemiological Surveillance Network (RENAVE) and excess allcause deaths from the Monitoring Mortality System (MoMo), up to 15 July 2020. SARS-CoV-2 infections in Spain were derived from the estimated seroprevalence by a chemiluminescent microparticle immunoassay for IgG antibodies in 61 098 participants in the ENE-COVID nationwide seroepidemiological survey between 27 April and 22 June 2020.1

RESULTS

The overall infection fatality risk was 0.8% (19 228 of 2.3 million infected individuals, 95% confidence interval 0.8% to 0.9%) for confirmed covid-19 deaths and 1.1% (24 778 of 2.3 million infected individuals, 1.0% to 1.2%) for excess deaths. The infection fatality risk was 1.1% (95% confidence interval 1.0% to 1.2%) to 1.4% (1.3% to 1.5%) in men and 0.6% (0.5% to 0.6%) to 0.8% (0.7% to 0.8%) in women. The infection fatality risk increased sharply after age 50, ranging from 11.6% (8.1% to 16.5%) to 16.4% (11.4% to 23.2%) in men aged 80 or more and from 4.6% (3.4% to 6.3%) to 6.5% (4.7% to 8.8%) in women aged 80 or more.

CONCLUSION

The increase in SARS-CoV-2 infection fatality risk after age 50 appeared to be more noticeable in men than in women. Based on the results of this study, fatality from covid-19 was greater than that reported for other common respiratory diseases, such as seasonal influenza.

Introduction

The infection fatality risk—the proportion of infected individuals who die from an infection—is a key indicator to design public health policies to control infectious diseases. Because the magnitude of the infection fatality risk for severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) is still debated,1 2 lockdowns and other forms of social distancing have been questioned as appropriate responses to the coronavirus disease 2019 (covid-19) pandemic.

An accurate estimation of the infection fatality risk of SARS-CoV-2 is difficult. Even if all symptomatic infections were diagnosed, which has not happened so far in most countries, asymptomatic infections cannot be clinically identified. Therefore, estimating the infection fatality risk relies on population based seroepidemiological surveys that provide an estimate of the proportion of individuals infected, regardless of symptoms.3 Also, because determining the number of deaths from covid-19 is often difficult, calculation of the infection fatality risk can be complemented with data on excess mortality.
A recent unpublished review of 24 serological reports, several also unpublished, estimated an overall infection fatality risk of 0.68% (95% confidence interval 0.53% to 0.83%). The methodological quality of many of these studies was questionable, however, with some exceptions. Estimates of infection fatality risk were mostly based on surveillance registered deaths, and substantial heterogeneity was seen between studies, with estimates ranging from 0.16% to 1.60%. Also, because the infection fatality risk for SARS-CoV-2 is expected to increase with age and might differ by sex, overall crude estimates of infection fatality risk cannot be directly compared between populations with different age and sex structures (eg, China and western Europe). Accurate and reliable age and sex specific estimates of infection fatality risk are needed.

We report overall, and age and sex specific, estimates of the infection fatality risk for SARS-CoV-2 from the large nationally representative seroepidemiological survey of SARS-CoV-2 virus infection in the community dwelling Spanish population (Encuesta Seroepidemiológica de la Infección por el Virus SARS-CoV-2 en España; ENE-COVID).

**Methods**

**Estimation of SARS-CoV-2 infections**

We calculated the prevalence of IgG antibodies against SARS-CoV-2 in the community dwelling Spanish population with data from ENE-COVID, a nationwide population based seroepidemiological survey. The design of the survey has been described previously. Briefly, 1500 census tracts, and up to 24 households within each tract, were randomly selected with a two stage sampling method stratified by province and municipality size. All residents in the 35 885 selected households were invited to participate in the study, resulting in a selected sample of 104 605 individuals of all ages. Serial data from epidemiological questionnaires and serology tests were collected for all study participants in three successive follow-up rounds between 27 April and 22 June 2020. Each round was completed in two weeks, with a one week break between rounds. Half of the households were randomly assigned to data collection during the first week of each round and the other half to the second week, so that serum specimens were collected in all participants two to four weeks apart. The study used two immunoassays to detect IgG antibodies: a point-of-care test (Orient Gene Biotech covid-19 IgG/IgM Rapid Test Cassette) and a chemiluminescent microparticle immunoassay (CMIA) that required venepuncture (SARS-CoV-2 IgG for use with ARCHITECT; Abbott Laboratories, Abbott Park, IL; reference 06R8620), with better performance characteristics (supplementary methods and supplementary figure 1) show a summary of reported sensitivity and specificity estimates of the CMIA test.

Of 98 891 individuals who were eligible for the ENE-COVID seroepidemiological survey, 10 238 could not be contacted, 14 926 declined to participate, 15 had missing data for age, and 5421 did not have valid results from the point-of-care test. Of the remaining 68 291 study participants, 61 098 received the CMIA test in at least one round (61.8% of eligible individuals and 68.9% of contacted individuals), with 43 212 participants receiving the CMIA test in all three rounds, 11 618 in two rounds, and 6268 in one round (fig 1). Response to the CMIA test was lower in individuals aged younger than 10 (22.3% of eligible individuals) and older than 80 (51.7%), and in men aged 20-59 compared with women (62.6% v 70.1%).

We calculated the seroprevalence, overall and in groups defined by age and sex, as the proportion of participants who had detectable IgG antibodies against SARS-CoV-2 in any round by the CMIA test (supplementary table 1). To account for the different sampling selection probabilities by province and to adjust for non-response to the CMIA test based on sex, age, and average income in the census tracts, we assigned sampling weights to each study participant. Design based standard errors for seroprevalence were computed taking into account stratification by province and municipality size and the clustering of seropositivity by household and census tract.

In sensitivity analyses, we corrected the seroprevalence estimates of SARS-CoV-2 for the sensitivity and specificity of the CMIA test, which were estimated as 90.6% (95% confidence interval 88.1% to 92.6%) and 99.3% (99.0% to 99.5%), respectively, from a meta-analysis of 23 diagnostic accuracy studies. In these studies, sensitivity ranged from 75% to 100% in 1494 samples from patients with confirmed covid-19 of different severity, and specificity ranged from 97.5% to 100% in 7696 samples obtained before the pandemic, with moderate between study heterogeneity for sensitivity ($I^2=46\%$) and no heterogeneity for specificity ($I^2=0\%$) (supplementary methods and supplementary figure 1).

We calculated the number of seropositive people in Spain by multiplying the age and sex specific prevalence of IgG antibodies by the size of the corresponding community dwelling Spanish population groups as of 15 July 2020.

**Estimation of deaths from covid-19**

Given the practical difficulties in reporting and confirming deaths from covid-19 during the pandemic, we estimated the infection fatality risk separately for deaths with confirmed covid-19 and excess all cause deaths. The two sources of information were the Spanish National Epidemiological Surveillance Network (RENAVE) and the Monitoring Mortality System (MoMo).

RENAVE provided individual data on the 29 137 deaths with laboratory confirmed covid-19 registered in Spain up to 15 July 2020. Personal data were missing for 2 49 deaths (0.9%), which were distributed according to the sex and age group distributions of all other deaths. The median interval between onset of symptoms and death in the RENAVE data was 12 days (interquartile range 7-19).
MoMo collects information on deaths from 3945 municipal civil registries that cover 93% of the Spanish population. With a model described previously, MoMo data are used to quantify excess deaths for a particular period, as the difference between the observed daily deaths corrected for reporting delay and those expected based on historical seasonal variation (centred seven day moving averages) and a non-linear secular trend (annual median daily deaths) from the past 10 years. Between 1 March and 15 July 2020, 44 459 excess all cause deaths were estimated (mainly for 13 March to 22 May 2020). MoMo estimates are of similar magnitude to those reported with a different method.

RENAVE and MoMo do not distinguish between people living in long term care facilities and the community dwelling population. An estimated 9909 deaths with confirmed covid-19 and 19 681 deaths with suspected covid-19 were reported in long term care facilities, mainly nursing homes, during the same period (supplementary table 2). We subtracted these deaths from those identified by RENAVE and MoMo, respectively, in the population aged 60 and older (supplementary methods).

Estimation of infection fatality risks
The infection fatality risk is the number of deaths from covid-19 divided by the number of individuals with SARS-CoV-2 infection. We obtained separate estimates of the overall infection fatality risk based on confirmed covid-19 deaths from RENAVE (lower bound of deaths because of limited confirmation of deaths in the surveillance network) and the excess all cause deaths from MoMo (possible upper bound of deaths because included are deaths that might not result from direct or indirect effects of the pandemic). We then repeated the analyses in each stratum defined by sex and 10 year age groups. We also estimated sex and age standardised infection fatality risk ratios by geographical unit (NUTS1).

We calculated 95% confidence intervals based on delta methods that accounted for the binomial variance in the number of deaths and the estimated design based variance in the number of infections. The supplementary methods provide further details on point and interval estimates of infection fatality risks and standardised infection fatality risk ratios.

Analyses were carried out with survey commands in Stata, version 16, and survey package in R, version 4.

Results
The overall SARS-CoV-2 seroprevalence was 4.9% (95% confidence interval 4.6% to 5.3%), which corresponded to 2.3 million (95% confidence interval 2.2 to 2.5 million) community dwelling individuals in Spain with antibodies against SARS-CoV-2 by 22 June 2020 (table 1). Seroprevalence was similar in men and women and increased with age up to 20-29 (5.7-5.8%), with a smooth decline at older ages.

Up to 15 July 2020, 19 228 laboratory confirmed covid-19 deaths and 24 778 excess all cause deaths were estimated in community dwelling individuals in Spain with antibodies against SARS-CoV-2 by 22 June 2020 (table 1). Seroprevalence was similar in men and women and increased with age up to 20-29 (5.7-5.8%), with a smooth decline at older ages.

Overall, the infection fatality risk estimate was 0.83% (95% confidence interval 0.78% to 0.89%) for deaths with confirmed covid-19 and 1.07% (24 778 of 2.3 million infected individuals, 1.00% to 1.15%) for excess deaths. The corresponding estimates were 1.11% (95% confidence interval 1.02% to 1.21%) and 1.40% (1.29% to 1.52%) for men, and 0.58% (0.53% to 0.63%) for women.
The infection fatality risk estimate varied with age: less than one per 1000 to age 49, with much lower values in younger age groups (<1 per 10000 to age 29), and increased sharply in older age groups (fig 2). In men aged 80 or older, the infection fatality risk estimate was 11.6% (95% confidence interval 8.1% to 16.5%) for deaths with confirmed covid-19 and 16.4% (11.4% to 23.2%) for excess deaths. In women aged 80 or older, the corresponding estimates were 4.6% (3.4% to 6.3%) and 6.5% (4.7% to 8.8%) (table 1). Standardised infection fatality risk ratios for deaths with confirmed covid-19 were consistent between geographical regions and the whole country (supplementary fig 2).

In sensitivity analyses, the estimates of infection fatality risk corrected for imperfect sensitivity and specificity were slightly higher, with a corrected overall infection fatality risk of 0.88% (19 228 of 2.2 million infected individuals, 95% confidence interval 0.80% to 0.97%) for deaths with confirmed covid-19 and 1.14% (24 778 of 2.2 million infected individuals, 1.03% to 1.25%) for excess deaths (supplementary table 3).

Discussion

We estimated an infection fatality risk for SARS-CoV-2 of 0.83-1.07% in Spain up to 15 July 2020. The infection fatality risk was greater in men than in women and increased with age: 11.6-16.4% in men aged 80 or older and 4.6-6.5% in women aged 80 or older. The higher mortality in elderly people might result from a greater number of comorbidities (cardiovascular disease, type 2 diabetes, lung and chronic kidney diseases) and immunological changes that affect the severity of SARS-CoV-2 infections. The higher mortality in men might result from more comorbidities and risk factors (eg, smoking, obesity) than in women, and also differences in cellular immunity between men and women, including poorer T cell activation and an increase in proinflammatory cytokines in men.

Comparison with previous studies

Comparing estimates of infection fatality risk between studies is difficult because of differences in methodology and populations. For example, some studies reported case fatality risks (number of deaths divided by the number of individuals with confirmed covid-19) rather than infection fatality risks, and others estimated infection fatality risks based on modelling assumptions rather than on data from population based seroepidemiological surveys.

Our overall estimate of infection fatality risk was similar to that found in seroepidemiological surveys with a low risk of bias, and our sex and age specific estimates suggest that the heterogeneity in published estimates of infection fatality risk is largely because of the different sex and age structure of the population. Our crude estimates of infection fatality risk, in common with countries with a similar age structure, such as Italy, were greater than those for countries with a younger population. Variations in infection fatality risk values might also be related to the local dynamics of the pandemic (eg, surge in number of new infections of covid-19, spread of the virus among vulnerable communities), combined with the capacity of the health system to cope and treat many patients.

Table 1 | Infection fatality risk for severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) in community dwelling population by sex and age during the first wave of the covid-19 pandemic in Spain 2020

<table>
<thead>
<tr>
<th>Sex, age (years)</th>
<th>No in population (000s)</th>
<th>SARS-CoV-2 seroprevalence (%; 95% CI)*</th>
<th>Individuals with SARS-CoV-2 antibodies (000s; 95% CI)</th>
<th>No of confirmed covid-19 deaths</th>
<th>No of excess all cause deaths</th>
<th>Infection fatality risk (%; 95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Overall</td>
<td>46 887.1</td>
<td>4.9 (4.6 to 5.3)</td>
<td>2305.8 (2152.8 to 2469.1)</td>
<td>19 228</td>
<td>24 778</td>
<td>0.83 (0.78 to 0.89)</td>
</tr>
<tr>
<td>Male</td>
<td>23 006.9</td>
<td>4.8 (4.4 to 5.2)</td>
<td>1105.1 (1016.9 to 1200.6)</td>
<td>12 317</td>
<td>15 480</td>
<td>1.11 (1.02 to 1.21)</td>
</tr>
<tr>
<td>0-9</td>
<td>2205.5</td>
<td>3.2 (1.9 to 5.4)</td>
<td>71.7 (62.5 to 119.7)</td>
<td>3</td>
<td>32</td>
<td>0.00 (0.00 to 0.01)</td>
</tr>
<tr>
<td>10-19</td>
<td>2557.9</td>
<td>3.6 (2.8 to 4.8)</td>
<td>93.3 (71.0 to 122.2)</td>
<td>3</td>
<td>0</td>
<td>0.01 (0.01 to 0.02)</td>
</tr>
<tr>
<td>20-29</td>
<td>2479.1</td>
<td>5.8 (4.7 to 7.3)</td>
<td>142.7 (116.1 to 174.9)</td>
<td>2</td>
<td>0</td>
<td>0.01 (0.01 to 0.02)</td>
</tr>
<tr>
<td>30-39</td>
<td>2978.7</td>
<td>4.7 (3.8 to 5.7)</td>
<td>139.7 (114.0 to 170.9)</td>
<td>48</td>
<td>3</td>
<td>0.03 (0.02 to 0.05)</td>
</tr>
<tr>
<td>40-49</td>
<td>3916.7</td>
<td>5.3 (4.6 to 6.2)</td>
<td>209.0 (180.0 to 242.4)</td>
<td>192</td>
<td>168</td>
<td>0.09 (0.07 to 0.11)</td>
</tr>
<tr>
<td>50-59</td>
<td>3491.8</td>
<td>5.3 (4.6 to 6.1)</td>
<td>184.0 (157.8 to 214.3)</td>
<td>705</td>
<td>601</td>
<td>0.38 (0.32 to 0.45)</td>
</tr>
<tr>
<td>60-69</td>
<td>2598.2</td>
<td>4.9 (4.1 to 5.9)</td>
<td>127.1 (105.3 to 153.2)</td>
<td>1904</td>
<td>2065</td>
<td>1.50 (1.24 to 1.82)</td>
</tr>
<tr>
<td>70-79</td>
<td>1783.7</td>
<td>4.7 (3.7 to 6.0)</td>
<td>83.6 (65.4 to 106.5)</td>
<td>4145</td>
<td>5114</td>
<td>4.96 (3.87 to 6.33)</td>
</tr>
<tr>
<td>≥80</td>
<td>993.3</td>
<td>4.6 (3.2 to 6.5)</td>
<td>45.6 (31.8 to 64.9)</td>
<td>5299</td>
<td>7497</td>
<td>11.6 (8.06 to 16.5)</td>
</tr>
<tr>
<td>Female</td>
<td>23 880.1</td>
<td>5.0 (4.7 to 5.4)</td>
<td>1200.5 (1105.0 to 1297.4)</td>
<td>6911</td>
<td>9298</td>
<td>0.58 (0.53 to 0.62)</td>
</tr>
<tr>
<td>0-9</td>
<td>2078.3</td>
<td>4.2 (2.7 to 6.7)</td>
<td>88.0 (55.1 to 139.0)</td>
<td>2</td>
<td>11</td>
<td>0.00 (0.00 to 0.01)</td>
</tr>
<tr>
<td>10-19</td>
<td>2396.7</td>
<td>4.4 (3.4 to 5.6)</td>
<td>105.1 (81.7 to 134.7)</td>
<td>3</td>
<td>22</td>
<td>0.00 (0.00 to 0.01)</td>
</tr>
<tr>
<td>20-29</td>
<td>2604.1</td>
<td>5.7 (4.6 to 7.0)</td>
<td>137.4 (111.2 to 169.3)</td>
<td>17</td>
<td>10</td>
<td>0.01 (0.01 to 0.02)</td>
</tr>
<tr>
<td>30-39</td>
<td>3012.4</td>
<td>5.2 (4.6 to 6.2)</td>
<td>156.7 (132.0 to 185.8)</td>
<td>29</td>
<td>71</td>
<td>0.02 (0.01 to 0.03)</td>
</tr>
<tr>
<td>40-49</td>
<td>3877.8</td>
<td>5.3 (4.6 to 6.2)</td>
<td>206.8 (177.9 to 240.0)</td>
<td>103</td>
<td>91</td>
<td>0.05 (0.04 to 0.06)</td>
</tr>
<tr>
<td>50-59</td>
<td>3563.5</td>
<td>5.2 (4.6 to 6.0)</td>
<td>184.4 (158.8 to 213.8)</td>
<td>318</td>
<td>369</td>
<td>0.17 (0.14 to 0.21)</td>
</tr>
<tr>
<td>60-69</td>
<td>2803.4</td>
<td>5.0 (4.2 to 6.0)</td>
<td>140.4 (117.2 to 167.9)</td>
<td>749</td>
<td>875</td>
<td>0.53 (0.44 to 0.65)</td>
</tr>
<tr>
<td>70-79</td>
<td>2138.1</td>
<td>4.6 (3.7 to 5.8)</td>
<td>98.9 (79.0 to 123.4)</td>
<td>1986</td>
<td>2666</td>
<td>2.01 (1.60 to 2.52)</td>
</tr>
<tr>
<td>≥80</td>
<td>1605.8</td>
<td>5.0 (3.7 to 6.8)</td>
<td>80.2 (58.7 to 108.9)</td>
<td>3704</td>
<td>5203</td>
<td>4.62 (3.38 to 6.29)</td>
</tr>
</tbody>
</table>

*Proportion of participants with detectable IgG antibodies against SARS-CoV-2 in any round of the ENE-COVID nationwide seroepidemiological survey by the chemiluminescent microparticle immunoassay, 27 April-22 June 2020, Spain.

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Strengths and limitations of this study

Our study has several strengths. Firstly, we used data from ENE-COVID, a nationwide population based seroepidemiological study. The first round of the study started one month after the peak incidence of COVID-19, around 20 March, and the last round ended on 22 June. Thus the study was designed to detect all individuals with antibodies during the first wave of the pandemic because most participants would have been infected one month before their first participation and IgG antibodies are detected two to three weeks after the onset of symptoms in more than 90% of infections. Although IgG antibodies in some infected participants could have decreased over time, particularly in mild infections, a recent study showed that this phenomenon occurs at least three months after infection, so substantial underestimation of the number of infected individuals is unlikely.

Secondly, the tests selected to measure antibodies against SARS-CoV-2 had high sensitivity and high specificity, according to our meta-analysis (supplementary fig 1). We used pooled estimates of sensitivity and specificity from the meta-analysis to calculate corrected ENE-COVID values for seroprevalence; the resulting estimates of infection fatality risk were slightly higher but consistent with the primary results (supplementary table 3), showing the robustness of our estimates.

Thirdly, because confirmation of deaths from COVID-19 is difficult during a large scale pandemic, we used mortality data from two sources: deaths in individuals with laboratory confirmed COVID-19 and excess all cause deaths. The latter included mortality directly because of SARS-CoV-2 infection and net mortality as a result of the societal effects of the pandemic and its control measures, such as delayed care for emergencies and for pre-existing chronic conditions caused by reorganisation of medical care and patients’ reluctance to seek attention, and reduced traffic injuries and other unintentional injuries. To include potentially delayed COVID-19 deaths, we considered all deaths registered up to 15 July (supplementary figure 3 shows the time distribution of individuals with COVID-19 and deaths in Spain).

A potential limitation of the study is that the two mortality surveillance systems (RENAVE and MoMo) did not differentiate between deaths in long term care facilities and deaths in the community. Therefore, we calculated deaths with confirmed and suspected COVID-19 in nursing homes from reports from the Spanish regional authorities, separated the deaths by sex and age group (as described in the supplementary methods), and subtracted these deaths from the total number of deaths. Because of insufficient diagnostic capacity in regions greatly affected by the first wave of the pandemic, individuals with COVID-19 could not be confirmed in many instances whereas deaths in patients with suspected COVID-19 living in nursing homes might have been underreported. If this were the case, the infection fatality risk based on excess deaths in the community dwelling population might be overestimated.

Our estimates of infection fatality risk do not apply to people living in nursing homes in Spain (about 334 000 residents; 76% aged 80 or older) where more than 19 000 people died during the study period. Long term care facilities have clusters of vulnerable populations where the virus could spread rapidly, and estimating the infection fatality risk for SARS-CoV-2 in long term care facilities will require a unique approach. Because nursing homes had limited access to hospital care during the initial outbreak in Spain, our estimates of infection fatality risk cannot be applied to elderly people in long term care facilities.

Policy implications and conclusions

We estimated the infection fatality risk for SARS-CoV-2 by age and sex from one of the largest seroepidemiological surveys in the world, carried out during the initial COVID-19 outbreak. Our overall infection fatality risk estimates (0.83-1.07%) were about 10 times larger than those for seasonal influenza. The high infection fatality risk in the older age groups supports existing measures (eg, social distancing, face masks, and educational campaigns) to shield these groups from infection. But relying exclusively on shielding elderly people might be a high risk strategy for the management of a pandemic. Given the high rate of transmission of the disease and the large proportion of susceptible individuals in the population, even a substantial reduction in
transmission in the elderly population could result in many deaths.

**AUTHOR AFFILIATIONS**

1National Centre for Epidemiology, Institute of Health Carlos III, Monforte de Lemos 5, 28029 Madrid, Spain
2Consortium for Biomedical Research in Epidemiology and Public Health (CIBERESP), Institute of Health Carlos III, Madrid, Spain
3Departments of Epidemiology and Biostatistics, Harvard TH Chan School of Public Health; Harvard-MIT Division of Health Sciences and Technology, Boston, MA, USA
4National Centre for Microbiology, Institute of Health Carlos III, Madrid, Spain
5Institute of Health Carlos III, Madrid, Spain
6Spanish Network for Research in Infectious Diseases (REIPI), Institute of Health Carlos III, Madrid, Spain
7Ministry of Health, Madrid, Spain
8National School of Public Health, Institute of Health Carlos III, Madrid, Spain

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**Ethical approval:** The ENE-COVID study was approved by the institutional review board of the Institute of Health Carlos III (register No PI 39_2020), written informed consent was obtained from participants.

**Data sharing:** The manuscript includes all values needed to replicate the infection fatality risk estimations, and the sources used are indicated. ENE-COVID seroprevalence values are provided here for all searches as aggregated values from the Spanish National Epidemiological Surveillance Network (RENAVE) and the Monitoring Mortality System (MoMo). Anonymised data from these systems are available on request. The specific formulary for this purpose is provided by the Department of Communicable Diseases, National Centre for Epidemiology, Carlos III Institute of Health, Monforte de Lemos 5, 28029 Madrid, Spain (vigilancia.cne@isciii.es and mortalidad@isciii.es). Population figures were provided by the National Institute of Statistics and are publicly available (https://www.ine.es). The code for calculation of the infection fatality risk is available at https://portalcan.isciii.es/enecovid19/.

The lead authors (RP-B, BP-G, and MP) affirm that the manuscript is an honest, accurate, and transparent account of the study being reported, and no important aspects of the study have been omitted.

Dissemination to participants and related patient and public communities: ENE-COVID individual results were provided to participants, a lay summary of the main results of ENE-COVID is available at https://portalcan.isciii.es/enecovid19/

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Web appendix: Supplementary material