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Long term exposure to ambient fine particulate matter and incidence of stroke: prospective cohort study from the China-PAR project

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ABSTRACT

OBJECTIVE

To study the effect of long term exposure to ambient fine particulate matter of diameter ≤ 2.5 μm ($\text{PM}_{2.5}$) on the incidence of total, ischemic, and hemorrhagic stroke among Chinese adults.

DESIGN

Population based prospective cohort study.

SETTING

Prediction for Atherosclerotic Cardiovascular Disease Risk in China (China-PAR) project carried out in 15 provinces across China.

PARTICIPANTS

117 575 Chinese men and women without stroke at baseline in the China-PAR project.

MAIN OUTCOME MEASURES

Incidence of total, ischemic, and hemorrhagic stroke.

RESULTS

The long term average $\text{PM}_{2.5}$ level from 2000 to 2015 at participants' residential addresses was $64.9 \mu\text{g}/\text{m}^3$, ranging from $31.2 \mu\text{g}/\text{m}^3$ to $97.0 \mu\text{g}/\text{m}^3$. During 900 214 person years of follow-up, 3540 cases of incident stroke were identified, of which 63.0% ($n=2230$) were ischemic and 27.5% ($n=973$) were hemorrhagic. Compared with the first quarter of exposure to $\text{PM}_{2.5}$ ($\leq 54.5 \mu\text{g}/\text{m}^3$), participants in the highest quarter ($>78.2 \mu\text{g}/\text{m}^3$) had an increased risk of incident stroke (hazard ratio 1.53, 95% confidence interval 1.34 to 1.74), ischemic stroke (1.82, 1.55 to

2.14), and hemorrhagic stroke (1.50, 1.16 to 1.93).

For each increase of $10 \mu\text{g}/\text{m}^3$ in $\text{PM}_{2.5}$ concentration, the increased risks of incident stroke, ischemic stroke, and hemorrhagic stroke were 13% (1.13, 1.09 to 1.17), 20% (1.20, 1.15 to 1.25), and 12% (1.12, 1.05 to 1.20), respectively. Almost linear exposure-response relations between long term exposure to $\text{PM}_{2.5}$ and incident stroke, overall and by its subtypes, were observed.

CONCLUSIONS

This study provides evidence from China that long term exposure to ambient $\text{PM}_{2.5}$ at relatively high concentrations is positively associated with incident stroke and its major subtypes. These findings are meaningful for both environmental and health policy development related to air pollution and stroke prevention, not only in China, but also in other low and middle income countries.

Introduction

Ambient air pollution, especially fine particulate matter ($\text{PM}_{2.5}$, particles with an aerodynamic diameter of $\leq 2.5 \mu\text{m}$) pollution, is a major public health concern in China and worldwide.¹ More than 90% of the world's population lives in regions where air quality exceeds the World Health Organization's annual mean $\text{PM}_{2.5}$ standard, set at $10 \mu\text{g}/\text{m}^3$.¹ The Global Burden of Disease Study estimated that 2.94 million deaths were attributable to ambient $\text{PM}_{2.5}$ air pollution globally in 2017, and about 48% of these were from ischemic heart diseases and strokes.² Moreover, the estimated disease burden of ambient $\text{PM}_{2.5}$ pollution was much higher in low and middle income countries such as China and India.³

Although previous studies of short term exposure (eg, over a few days) to ambient $\text{PM}_{2.5}$ have reported increased risks of hospital admission for stroke in China, evidence of the accumulated damage on cardiovascular health caused by long term exposure to air pollution was limited. Several cohort studies have investigated the association between long term exposure to ambient $\text{PM}_{2.5}$ over multiple years and incidence of stroke.⁴⁻⁷ Most of these studies, however, were conducted in North America and Europe, where the average ambient $\text{PM}_{2.5}$ levels have historically been lower than $25 \mu\text{g}/\text{m}^3$.³ Consequently, these studies provided limited insight into the health effects of exposure to extreme $\text{PM}_{2.5}$ air pollution, which were generally found in low and middle income countries.³ In

WHAT IS ALREADY KNOWN ON THIS TOPIC

The association between long term exposure to ambient $\text{PM}_{2.5}$ and incidence of stroke has been well documented in North America and Europe, with most studies having relatively low $\text{PM}_{2.5}$ concentrations (typically $\leq 25 \mu\text{g}/\text{m}^3$)

However, evidence is lacking about the association between long term exposure to $\text{PM}_{2.5}$ and incidence of stroke at high ambient concentrations, usually found in low and middle income countries

WHAT THIS STUDY ADDS

Long term exposure to ambient $\text{PM}_{2.5}$ at relatively high concentrations is positively associated with risk of incident stroke and its major subtypes

The results suggest almost linear exposure-response relations between exposure to $\text{PM}_{2.5}$ and incidence of stroke, covering a wide range of $\text{PM}_{2.5}$ levels (31.2 – $97.0 \mu\text{g}/\text{m}^3$)

These findings are important for environmental and health policies related to air pollution and stroke prevention in China, as well as for other low and middle income countries

addition, the cause and risk profiles of ischemic stroke versus hemorrhagic stroke are different, highlighting the importance of considering these events separately.⁴ However, few studies have distinguished between the effects of long term exposure to PM_{2.5} on ischemic and hemorrhagic stroke.

Epidemiological studies focusing on the chronic health effects of exposure to ambient PM_{2.5} in China have been hindered owing to the absence of routine air quality monitoring networks before 2013.⁸ To support such studies, several spatiotemporal statistical models based on satellite retrieved aerosol optical depth have been developed to provide long term estimates of PM_{2.5}.⁸⁻⁹ Combining monitoring data and satellite based PM_{2.5} estimations has now made it possible to provide accurate PM_{2.5} exposure assessment for epidemiological studies. Existing studies pertaining to the adverse health impacts of PM_{2.5} in China have, however, primarily assessed exposure to PM_{2.5} at relatively coarse spatial resolution (eg, 10×10 km),¹⁰⁻¹¹ which may have failed to capture fine scale PM_{2.5} gradients thereby leading to misclassification of exposure.

We investigated the association between long term exposure to ambient PM_{2.5} and incidence of stroke, overall and by subtypes, using data from satellite based high quality PM_{2.5} estimates at 1×1 km spatial resolution combined with the Prediction for Atherosclerotic Cardiovascular Disease Risk in China (China-PAR)—an established prospective project.

Methods

Study population

The China-PAR project was established to investigate the distribution patterns of cardiovascular diseases and associated risk factors in the general Chinese population. A detailed description of the design is published elsewhere.¹²⁻¹³ Briefly, the China-PAR project was carried out in 15 Chinese provinces and included four subcohorts: China MUCA (1992-94) (China Multi-Center Collaborative Study of Cardiovascular Epidemiology),¹⁴ China MUCA (1998),¹⁵ InterASIA (International Collaborative Study of Cardiovascular Disease in Asia),¹⁶ and CIMIC (Community Intervention of Metabolic Syndrome in China and Chinese Family Health Study).¹² From 1992 to 2008, a total of 127 840 Chinese adults (≥18 years) were initially enrolled for the baseline examinations. The subcohorts were last followed-up between 2012 and 2015 (see supplementary file for details).

Data collection

For all subcohorts we used identical methods at baseline and follow-up surveys. Trained healthcare staff used a standardized questionnaire under strict quality control to collect information on personal characteristics, medical history, and lifestyle risk factors. Smoking was defined as having consumed at least 400 cigarettes or 500 g of tobacco leaves throughout the lifetime, or at least one cigarette daily for one year or more. We categorized participants into

never, former, and current smokers. Current smokers were defined as those who were smoking at the time of survey, and former smokers were those who ever smoked but had quit at the time of survey. Alcohol consumption was defined as drinking alcohol at least once weekly during the past year. Education was classified as high school or above (≥10 years of education received) or less than high school (<10 years). Work related physical activity was categorized into vigorous or moderate activity, light or sedentary activity, or no job or retirement based on participants' occupation. Body weight and height were measured to the nearest 0.5 kg and 0.5 cm, respectively, with participants wearing only lightweight clothing and no shoes. Body mass index (BMI) was calculated as weight (kg)/(height (m)²).

Trained healthcare practitioners measured participants' blood pressure during clinic or home visits in accordance with the protocol recommended by the American Heart Association.¹⁷ Three measurements were obtained after participants had rested for five minutes. For this analysis we used the average of the three readings. In addition, blood samples were obtained from participants after a prescribed 10 hours of fasting to measure both serum glucose and lipid levels.

Information on stroke incidence during the follow-up period was collected by interviewing study participants or their proxies, and further checking hospital records or death certificates for verification. Local investigators initially recorded fatal and non-fatal stroke events. The central adjudication committee at Fuwai Hospital (Beijing, China) reviewed all medical and death records and determined the final diagnosis. Two adjudication committee members verified events independently, and discrepancies were resolved by discussion involving additional committee members. Causes of death were coded according to ICD-10 (international classification of diseases, 10th revision). For this analysis, incident stroke was defined as a confirmed diagnosis of first ever fatal or non-fatal stroke event during follow-up (I60-I69). We also classified subtypes of stroke: ischemic stroke (I63), hemorrhagic stroke (I60-I62), and unspecified stroke (I64-I69).

Written informed consent was obtained from each participant before data collection.

Exposure assessment

We used a satellite based spatiotemporal model to estimate ambient PM_{2.5} levels at 1×1 km spatial resolution. A detailed methodology of this model is published elsewhere.¹⁸⁻¹⁹ Briefly, we used machine learning algorithms to estimate ground level PM_{2.5} concentrations in China based on satellite retrieved aerosol optical depth retrievals from US National Aeronautics and Space Administration, land use information, roads, meteorology, and population density data. To ensure a high spatiotemporal coverage and minimize the bias of PM_{2.5} estimations, we used a multiple imputation method to fill

missing aerosol optical depth values. Compared with our previous $PM_{2.5}$ prediction model,¹⁸ we improved the model for the current study by using the 1×1 km resolution aerosol optical depth data retrieved through a Multi-Angle Implementation of Atmospheric Correction algorithm²⁰ and estimated monthly $PM_{2.5}$ concentrations at 1×1 km resolution. We then validated this model using ground $PM_{2.5}$ measurements from the China Environmental Monitoring Center (www.cnemc.cn/). We obtained an overall 10-fold cross validation R^2 of 0.93 at the monthly level and 0.95 at the annual level. To assess the accuracy of predicting historical $PM_{2.5}$ concentrations (before 2013), we compared model predictions with available monitoring data from Hong Kong, Taiwan, and the US Embassy in Beijing, Shanghai, and Guangzhou, and the prediction R^2 and root mean squared error were 0.80 and $8.90 \mu\text{g}/\text{m}^3$, respectively at the annual level.

Based on this model, we obtained monthly $PM_{2.5}$ levels from 2000 to 2015 in China. Each participant's residential address was geocoded into latitude and longitude data. To account for participants who changed address, in the main analyses we calculated a time weighted average exposure to $PM_{2.5}$ from 2000 to 2015 for each participant, with weights defined as the duration spent at each residence. In addition, we also calculated the annual $PM_{2.5}$ level as a time varying exposure in a sensitivity analysis, taking into account the temporal variations of $PM_{2.5}$ during our study period.

Statistical analysis

We present the baseline characteristics of the included participants as means with corresponding standard deviations for continuous variables and as percentages for categorical variables.

Person years of follow-up were calculated from the date of initial examination or 1 January 2000 (for participants enrolled before 2000) until the date of stroke occurrence, death, or the date of most recent follow-up, whichever occurred first. Stratified Cox proportional hazards regression models with strata defined as subcohorts were used to examine the association between long term exposure to $PM_{2.5}$ and stroke, which were commonly adopted regression models in previous epidemiological studies.^{21–22} We assessed the proportional hazards assumption by evaluating the weighted Schoenfeld residuals,²³ and no violations were observed ($P > 0.05$). Participants were analyzed in one dataset to ensure sufficient statistical power and wider gradients of exposure to $PM_{2.5}$. We categorized participants into four groups according to quarters of exposure to $PM_{2.5}$ (ie, 31.2–54.5 $\mu\text{g}/\text{m}^3$, 54.6–59.6 $\mu\text{g}/\text{m}^3$, 59.7–78.2 $\mu\text{g}/\text{m}^3$, and 78.3–97.0 $\mu\text{g}/\text{m}^3$). When $PM_{2.5}$ was considered a continuous variable, we reported the hazard ratio of incident stroke for each $10 \mu\text{g}/\text{m}^3$ increment in $PM_{2.5}$ level. An exposure-response curve between $PM_{2.5}$ concentration and incident stroke was plotted using a penalized spline with two degrees of freedom. We

evaluated the Akaike Information Criterion to decide on degrees of freedom.²⁴

In the multivariate adjusted Cox models, we included covariates that we hypothesized a priori could potentially confound the relation between $PM_{2.5}$ concentration and stroke. We included three models using covariates collected at baseline.^{6 11 25 26} In model 1, we adjusted for age and sex. In model 2, we additionally adjusted for urbanization (urban or rural area) and geographical region. To account for the residual confounders related to spatial variation in stroke onset, we included seven geographic regions—north, east, north eastern, south, central, south western, and north western. In model 3, we further adjusted for smoking status (never, former, or current smoker), alcohol intake (yes or no), education level (less than high school, or high school or above), work related physical activity (vigorous or moderate, light or sedentary, no job or retirement), body mass index, and hypertension (yes or no). Model 3 was used as the main model in further analyses.

Based on previous evidence of possible effect modifications,^{5 6} we conducted subgroup analyses by age, sex, cigarette smoking, alcohol intake, BMI, abdominal obesity, hypertension, diabetes, education level, and urbanization using model 3. We fitted separate Cox models for each subgroup and obtained subgroup specific effect estimates of $PM_{2.5}$. Then we performed a two sample z test to assess whether the effect estimates between subgroups were statistically different.^{21 27 28}

Based on model 3, we performed several sensitivity analyses by further adjusting for baseline diabetes and total cholesterol levels, and alternately by excluding those who had coronary heart disease at the baseline examination, or simply by excluding incident stroke events that occurred during the first year of follow-up. To assess the confounding role of neighborhood socioeconomic status, we collected the information about county level averaged years of education as a surrogate for socioeconomic status, and we further adjusted for it in the Cox model. Since the baseline information was collected several years before 2000 for the China MUCA (1992–94) subcohort, we conducted a sensitivity analysis by excluding this subcohort. Moreover, to investigate the potential confounding effect of smoking dose, we conducted a sensitivity analysis by further adjusting for smoking pack year history in three of the subcohorts; the China MUCA (1992–94) subcohort had no information on smoking dose.

To consider the temporal variation of $PM_{2.5}$ concentrations, we conducted a sensitivity analysis by using annual $PM_{2.5}$ level as a time varying exposure in the survival model. In each annual scale of this model, all covariates were updated using the latest follow-up information. We also conducted an additional sensitivity analysis by using baseline two year average $PM_{2.5}$ levels as the long term exposure (ie, average $PM_{2.5}$ exposures from 2000 to 2001 for the China MUCA (1992–94), China MUCA (1998), and InterASIA

Fig 1 | Flowchart of study participants included and excluded in analyses. China-PAR=Prediction for Atherosclerotic Cardiovascular Disease Risk in China project

cohorts; averages from 2007 to 2008 for the CIMIC cohort), and we re-ran the Cox models.

Statistical analyses were performed using SAS 9.4 (SAS Institute, Cary, NC) and R software, version 3.4.2 (R Foundation for Statistical Computing, Vienna, Austria). Tests were two sided with statistical significance set at $P < 0.05$.

Patient and public involvement

Patients and members of the public were not involved in setting the research question or the outcome measures, nor were they involved in design and implementation of the study. The results of physical examination and biochemical test were disseminated to each study participant.

Results

Of the original participants in China-PAR project, 119 388 (93.4%) were followed-up successfully. Because the $PM_{2.5}$ exposure data were available since 2000, we used the follow-up information after that year. Therefore, we excluded 1548 participants with previous stroke and 165 deceased participants before 2000. A further 100 participants were excluded because of missing residential details, leaving 117 575 participants for the current analysis (fig 1).

Figure 2 shows the cities in which participants resided at baseline. The average $PM_{2.5}$ level from 2000 until 2015 at the participants' residential address was $64.9 \mu g/m^3$, ranging from $31.2 \mu g/m^3$ to $97.0 \mu g/m^3$. Table 1 presents the baseline characteristics, overall and according to quarters of exposure to $PM_{2.5}$. At baseline the mean age of participants was 50.9 years, 41.0% were men, and 23.9% were current smokers. Participants with higher exposure to $PM_{2.5}$ were less often classified as current smokers and tended to have higher BMI, systolic blood pressure, and diastolic blood pressure levels.

During 900 214 person years of follow-up, 3540 cases of incident stroke were identified, with an incidence rate of 393 cases per 100 000 person years. Of the total number of stroke cases, 63.0% ($n=2230$) were classified as ischemic stroke, 27.5% ($n=973$) were classified as hemorrhagic stroke, 8.2% ($n=291$) were considered of unknown subtype, and 1.3% ($n=46$) were classified as both ischemic and hemorrhagic stroke subtypes. Table 2 presents the crude and multivariate adjusted hazard ratios and 95% confidence intervals for the association of incident stroke with long term exposure to $PM_{2.5}$. In the age and sex adjusted model, higher exposure to $PM_{2.5}$ was significantly associated with an increased risk of stroke ($P < 0.001$). After multivariate adjustment, we observed similar results,

Fig 2 | Average fine particulate matter of diameter $\leq 2.5 \mu m$ ($PM_{2.5}$) levels in China, 2000-15. Red solid circles denote cities where study participants resided at baseline. AOD=aerosol optical depth

Table 1 | Baseline characteristics of participants in Prediction for Atherosclerotic Cardiovascular Disease Risk in China (China-PAR) project and according to quarters of PM_{2.5}* concentrations. Values are numbers (percentages) unless stated otherwise

Characteristics	Total cohort	Quarters of PM _{2.5} levels			
		First (31.2-54.5 µg/m ³)	Second (54.6-59.6 µg/m ³)	Third (59.7-78.2 µg/m ³)	Fourth (78.3-97.0 µg/m ³)
No of participants	117 575	28 540	30 751	28 951	29 333
Mean (SD) PM _{2.5} level (µg/m ³)	64.9 (14.2)	48.3 (5.8)	57.0 (1.6)	71.1 (7.7)	83.3 (1.8)
Mean (SD) age at baseline (years)	50.9 (11.8)	50.6 (10.3)	51.2 (12.9)	50.4 (12.3)	51.4 (11.6)
Men	48 225 (41.0)	12 279 (43.0)	12 110 (39.4)	11 997 (41.4)	11 839 (40.4)
Smoking status:					
Never	84 797 (72.4)	19 867 (69.9)	22 156 (72.7)	20 519 (71.1)	22 255 (76.0)
Former	4 358 (3.7)	1 148 (4.0)	979 (3.2)	1 472 (5.1)	759 (2.6)
Current	27 921 (23.9)	7 416 (26.1)	7 334 (24.1)	6 889 (23.9)	6 282 (21.4)
Education:					
Less than high school	100 364 (85.8)	23 252 (82.3)	27 300 (89.3)	24 294 (84.1)	25 518 (87.1)
High school or above	16 650 (14.2)	5 012 (17.7)	3 285 (10.7)	4 589 (15.9)	3 764 (12.9)
Work related physical activity:					
Vigorous or moderate level	65 201 (55.6)	17 100 (60.2)	16 414 (53.6)	15 910 (55.1)	15 777 (53.9)
Light or sedentary level	33 159 (28.3)	8 628 (30.4)	9 152 (29.9)	7 772 (26.9)	7 607 (26.0)
No job or retirement	18 829 (16.1)	2 673 (9.4)	5 035 (16.5)	5 218 (18.1)	5 903 (20.2)
Urban residence	14 456 (12.3)	6 698 (23.5)	972 (3.2)	3 473 (12.0)	3 313 (11.3)
Alcohol drinker	22 924 (19.5)	7 300 (25.6)	6 299 (20.5)	3 931 (13.6)	5 394 (18.4)
Mean (SD) body mass index	23.6 (3.6)	22.8 (3.5)	23.0 (3.4)	23.8 (3.5)	24.6 (3.7)
Mean (SD) systolic blood pressure (mm Hg)	127.6 (21.2)	121.4 (18.5)	126.9 (19.2)	126.7 (21.0)	135.2 (23.3)
Mean (SD) diastolic blood pressure (mm Hg)	78.8 (11.7)	75.8 (11.2)	77.6 (11.1)	78.6 (11.4)	83.2 (11.8)
Diabetes	6 137 (5.5)	967 (3.8)	1 289 (4.5)	1 947 (7.0)	1 934 (6.7)
Hypercholesterolemia	5 014 (4.5)	1 649 (6.5)	766 (2.7)	1 257 (4.4)	1 342 (4.7)

*Particles of diameter ≤ 2.5 µm.

with hazard ratios of 1.11 (95% confidence interval 0.98 to 1.26), 1.30 (1.13 to 1.49), and 1.53 (1.34 to 1.74) for participants exposed to the second, third, and fourth quarters of PM_{2.5} concentrations, respectively.

Each 10 µg/m³ increase in PM_{2.5} was significantly associated with a 13% increase in stroke incidence

(1.13, 1.09 to 1.17) after multivariate adjustment using model 3. In the analyses for stroke subtypes, incidence of both ischemic stroke and hemorrhagic stroke were significantly associated with long term exposure to PM_{2.5}. Each 10 µg/m³ increase in PM_{2.5} was associated with a 20% increase in incidence of ischemic stroke

Table 2 | Stroke incidence associated with long term exposure to fine particulate matter (PM_{2.5})*. Values are hazard ratios (95% confidence intervals) unless stated otherwise

Characteristics	Quarters of PM _{2.5} levels				P trend	PM _{2.5} per 10 µg/m ³ increment
	First (31.2-54.5 µg/m ³)	Second (54.6-59.6 µg/m ³)	Third (59.7-78.2 µg/m ³)	Fourth (78.3-97.0 µg/m ³)		
Stroke incidence						
Person years	270 110	212 760	216 536	200 808	—	—
No of cases	825	718	995	1002	—	—
Incidence rate†	305.4	337.5	459.5	499.0	—	—
Model 1‡	1.00	1.06 (0.95 to 1.17)	1.80 (1.63 to 1.98)	2.02 (1.83 to 2.22)	<0.001	1.26 (1.23 to 1.29)
Model 2§	1.00	1.20 (1.06 to 1.35)	1.54 (1.34 to 1.76)	1.95 (1.72 to 2.22)	<0.001	1.22 (1.18 to 1.27)
Model 3¶	1.00	1.11 (0.98 to 1.26)	1.30 (1.13 to 1.49)	1.53 (1.34 to 1.74)	<0.001	1.13 (1.09 to 1.17)
Ischemic stroke						
Person years	270 799	213 025	217 093	201 306	—	—
No of cases	495	435	627	719	—	—
Incidence rate†	182.8	204.2	288.8	357.2	—	—
Model 1‡	1.00	1.10 (0.97 to 1.26)	2.09 (1.85 to 2.37)	2.62 (2.32 to 2.95)	<0.001	1.35 (1.31 to 1.39)
Model 2§	1.00	1.18 (1.01 to 1.38)	1.77 (1.49 to 2.11)	2.31 (1.97 to 2.71)	<0.001	1.29 (1.24 to 1.35)
Model 3¶	1.00	1.10 (0.94 to 1.29)	1.49 (1.25 to 1.77)	1.82 (1.55 to 2.14)	<0.001	1.20 (1.15 to 1.25)
Hemorrhagic stroke						
Person years	271 437	213 425	218 300	203 168	—	—
No of cases	264	216	290	249	—	—
Incidence rate†	97.3	101.2	132.8	122.6	—	—
Model 1‡	1.00	0.97 (0.81 to 1.18)	1.41 (1.19 to 1.68)	1.39 (1.16 to 1.66)	<0.001	1.14 (1.09 to 1.19)
Model 2§	1.00	1.31 (1.04 to 1.66)	1.64 (1.27 to 2.12)	1.99 (1.56 to 2.55)	<0.001	1.23 (1.15 to 1.31)
Model 3¶	1.00	1.21 (0.94 to 1.54)	1.40 (1.07 to 1.82)	1.50 (1.16 to 1.93)	<0.001	1.12 (1.05 to 1.20)

*Particles of diameter ≤ 2.5 µm.

†Incident rate per 100 000 person years.

‡Model 1: stratified Cox proportional hazard model, adjusted for age and sex.

§Model 2: model 1+adjusted for geographic region and urbanization.

¶Model 3: Model 2+adjusted for education, smoking status, alcohol consumption, physical activity, body mass index, and hypertension.

(1.20, 1.15 to 1.25) and a 12% increase in incidence of hemorrhagic stroke (1.12, 1.05 to 1.20).

Supplementary figure S1 shows the association of exposure to $PM_{2.5}$ with incident stroke and related subtypes stratified by selected characteristics. The effect of $PM_{2.5}$ on stroke incidence did not vary by individual risk factors. For ischemic stroke, however, stronger associations seemed to exist among elderly people than among younger people ($P=0.05$ for effect modification). In addition, for ischemic stroke stronger effects of $PM_{2.5}$ were observed in normal weight participants compared with overweight participants ($P=0.02$ for effect modification), whereas no effect modification was observed for abdominal obesity.

Figure 3 illustrates the exposure-response curve for the association between long term exposure to $PM_{2.5}$

and incident risk of all stroke, ischemic stroke, and hemorrhagic stroke. The exposure-response curves were observed to be monotonic and almost linear across the range of $PM_{2.5}$ concentrations in the study population.

Table 3 shows the results of several sensitivity analyses based on model 3. Overall, no substantial changes occurred in the effect estimates of $PM_{2.5}$ across the various analyses. When further adjustment was made for the county level averaged years of education, the effect estimates were 1.15 (1.11 to 1.19) for total stroke, 1.24 (1.18 to 1.30) for ischemic stroke, and 1.13 (1.06 to 1.21) for hemorrhagic stroke. When using annual mean $PM_{2.5}$ concentrations as time-varying exposures or baseline two year average $PM_{2.5}$ levels as long term exposures in the Cox model, similar associations with the main analyses were obtained. Exclusion of the China MUCA (1992-94) subcohort did not substantially influence the results. Furthermore, in the analysis with smoking pack year history considered in the three subcohorts with available information on smoking dose, the associations between $PM_{2.5}$ level and incident stroke were observed to be not significantly different from those fitted using model 3. Adjustment for baseline diabetes and total cholesterol levels did not substantially alter the results either. In addition, the results remained stable after excluding those with a diagnosis of coronary heart disease at baseline examination or when excluding those with incident stroke events occurring within the first year of follow-up.

Discussion

In this prospective cohort study in China we found significantly higher risks of incident stroke, ischemic stroke, and hemorrhagic stroke associated with long term exposure to high $PM_{2.5}$ levels. We identified almost linear exposure-response relations between $PM_{2.5}$ level and incident stroke and both subtypes separately.

Comparison with other studies

The results of previous studies on the long term effects of exposure to $PM_{2.5}$ on stroke incidence have been inconclusive.^{4-6 29 30} Our findings highlighted a positive effect between long term exposure to $PM_{2.5}$ and incident stroke that were consistent with findings from the European Study of Cohorts for Air Pollution Effects project and the Women's Health Initiative.^{5 6} However, analyses of the Health Professionals cohort and other cohorts from UK and Sweden failed to identify statistically significant effects of air pollution on incidence of stroke.²⁹⁻³¹

The magnitude of effects of $PM_{2.5}$ on stroke incidence in the current study (hazard ratio 1.13, 95% confidence interval 1.09 to 1.17, per 10 $\mu g/m^3$ increment) was similar to that found in a recent meta-analysis of separate cohort studies in North America and Europe (ie, hazard ratio 1.064, 95% confidence interval 1.021 to 1.109) per 5 $\mu g/m^3$ increase in $PM_{2.5}$.³² In addition, we found an association between exposure to $PM_{2.5}$ and an increase in the risk of both incident ischemic

Fig 3 | Exposure-response relations between long term exposure to fine particulate matter of diameter $\leq 2.5 \mu m$ ($PM_{2.5}$) and incidence of total stroke, ischemic stroke, and hemorrhagic stroke among Chinese adults, using a penalized spline function with two degrees of freedom. Hazard ratios were estimated by comparing to a reference value of $31.2 \mu g/m^3$. Solid line represents point estimate and dashed lines represent 95% confidence intervals. Boxplot represents distribution of $PM_{2.5}$ exposure levels for participants

Table 3 | Sensitivity analysis for association of stroke incidence with 10 µg/m³ increase in fine particulate matter (PM_{2.5})*

Sensitivity analysis	No of cases	Hazard ratio (95% CI)	P value
Further adjusting for county level averaged years of education			
Total stroke	3509	1.15 (1.11 to 1.19)	<0.001
Ischemic stroke	2262	1.24 (1.18 to 1.30)	<0.001
Hemorrhagic stroke	1004	1.13 (1.06 to 1.21)	<0.001
Annual mean PM_{2.5} as a time varying exposure during follow-up			
Total stroke	3511	1.17 (1.13 to 1.21)	<0.001
Ischemic stroke	2262	1.23 (1.18 to 1.29)	<0.001
Hemorrhagic stroke	1006	1.14 (1.07 to 1.21)	<0.001
Baseline two year average PM_{2.5} exposure			
Total stroke	3509	1.15 (1.11 to 1.20)	<0.001
Ischemic stroke	2262	1.23 (1.17 to 1.29)	<0.001
Hemorrhagic stroke	1004	1.13 (1.05 to 1.21)	<0.001
Further adjusting for smoking pack year†			
Total stroke	2769	1.18 (1.13 to 1.23)	<0.001
Ischemic stroke	1758	1.28 (1.21 to 1.34)	<0.001
Hemorrhagic stroke	798	1.16 (1.07 to 1.24)	<0.001
Exclusion of China MUCA (1992-94) study			
Total stroke	2977	1.17 (1.13 to 1.22)	<0.001
Ischemic stroke	1876	1.28 (1.22 to 1.35)	<0.001
Hemorrhagic stroke	868	1.14 (1.06 to 1.22)	<0.001
Further adjusting for baseline diabetes and total cholesterol level			
Total stroke	3326	1.12 (1.08 to 1.17)	<0.001
Ischemic stroke	2155	1.19 (1.14 to 1.25)	<0.001
Hemorrhagic stroke	940	1.10 (1.03 to 1.18)	0.005
Excluding those with coronary heart disease at baseline‡			
Total stroke	3459	1.14 (1.10 to 1.18)	<0.001
Ischemic stroke	2224	1.20 (1.15 to 1.26)	<0.001
Hemorrhagic stroke	996	1.12 (1.05 to 1.20)	<0.001
Excluding incident stroke occurring within first year of follow-up§			
Total stroke	3290	1.12 (1.08 to 1.16)	<0.001
Ischemic stroke	2128	1.17 (1.12 to 1.23)	<0.001
Hemorrhagic stroke	924	1.13 (1.06 to 1.21)	<0.001

MUCA=Multi-Center Collaborative Study of Cardiovascular Epidemiology.

*Particles of diameter ≤2.5 µm.

†Data on smoking pack year history were collected in three of four subcohorts: China MUCA (1998), InterASIA (International Collaborative Study of Cardiovascular Disease in Asia), and CIMIC (Community Intervention of Metabolic Syndrome in China and Chinese Family Health Study).

‡Coronary heart disease before baseline interview or 1 January 2000 (if date of baseline examination was earlier than 2000) was excluded.

§Incident stroke events within first year after baseline interview or 1 January 2000 (if date of baseline examination was earlier than 2000) was excluded.

stroke and incident hemorrhagic stroke, which was rarely reported in previous cohort studies.⁴ Our results are in line with previous time series studies or case crossover studies that have reported an increased risk of both ischemic stroke and hemorrhagic stroke after short term exposure to particulate matter.^{33 34} In contrast, a smaller cohort study with short follow-up, conducted in Hong Kong, China found only statistically significant associations of long term exposure to PM_{2.5} at low concentrations (mean 35.8 (SD 2.4) µg/m³) with ischemic stroke but not with hemorrhagic stroke.⁴ These differences may have resulted from a relatively small sample size, variation of follow-up duration, narrow range of exposure to PM_{2.5}, misclassification of stroke subtypes, and population characteristics. In subgroup analysis, although the P values for effect modification were not statistically significant in most subgroups, we found suggestive evidence of stronger associations between PM_{2.5} and incidence of ischemic stroke among elderly people and participants of normal body weight. These results were consistent with findings from the European Study of Cohorts for

Air Pollution Effects project.⁵ Further experimental studies were needed to understand the interaction effects between exposure to ambient PM_{2.5} and other risk factors in the development process of stroke.

Several researchers and practitioners have hypothesized differences between the mechanisms that link long term exposure to PM_{2.5} with ischemic and hemorrhagic stroke.^{33 35 36} It was hypothesized that exposure to PM_{2.5} might result in platelet activation, promoting blood coagulation and thrombosis.³⁷ In addition, long term exposure to PM_{2.5} might induce systemic inflammatory responses through increased release of plasma cytokines, leading to acceleration of atherosclerosis.³⁷ Exposure to ambient PM_{2.5} could increase the risk of ischemic stroke through these pathophysiologic changes, whereas the mechanisms by which exposure to PM_{2.5} increase the risk of hemorrhagic stroke might include a triggering of arterial vasoconstriction and increases in blood pressure.³⁸ Alternatively, endothelial dysfunction caused by air pollution might increase the vulnerability of brain vessels to rupture.³³ Further basic medical research would provide insight into this casual pathway.

Estimating exposure-response relations between long term exposure to PM_{2.5} and risk of stroke is critical for assessing the burden of diseases and making public health policies. In the current study, we found almost linear exposure-response relations between PM_{2.5} concentration and stroke, overall and according to each subtype. The GBD Study, however, suggested a non-linear relationship between exposure to PM_{2.5} and stroke mortality, with steeper slopes at lower exposures, which level off at higher exposures.³ To consider the shape of exposure-response curves at high ambient concentrations, the Global Burden of Disease Study integrated risk estimates from studies in high income countries with lower ambient PM_{2.5} levels, as well as effect estimates for high concentrations derived from studies of household solid cooking fuel, secondhand smoke, and active smoking.³⁹ This method was proposed without considering the variations in PM_{2.5} toxicity across different pollution sources (eg, ambient air pollution versus smoking), which may have influenced the effect estimates.^{39 40} With a broader range of PM_{2.5} exposure (ie, 31.2-97.0 µg/m³), our study filled the data gap at high ambient PM_{2.5} levels and raised the reliability of generalizations to China and other low and middle income countries, in terms of estimation of the burden of PM_{2.5} related disease.

Strengths and limitations of this study

This study has several advantages over previous research. First, we used data from large population based prospective cohorts with long follow-up and high quality outcome assessments. Second, our study participants were widely distributed across China, covering a wide range of geographic PM_{2.5} levels, from 31.2 µg/m³ to 97.0 µg/m³, which presented a valuable opportunity to assess exposure-response relations. Third, we used an advanced satellite based spatiotemporal model with high accuracy to estimate

ambient PM_{2.5} levels at 1×1 km resolution, which ensured accurate exposure assessment. Fourth, we collected data on multiple cardiovascular risk factors at the individual level, which allowed us to better control for potential confounders.

Some limitations of our study should be noted. First, we did not consider other air pollutants, meteorological factors, and environmental noise in our statistical models, because high quality and high resolution exposure estimates were unavailable. Second, we did not collect data around indoor air pollution sources such as household use of solid fuels. Third, we estimated exposures to ambient PM_{2.5} at the participants' residential address without considering their daily activities and location, which could result in misclassification of exposure. Indoor sources of air pollution and individual time-activity patterns are currently being collected and we will consider them in future analyses. Fourth, considering the limited statistical power, we did not perform analyses on subtypes of ischemic stroke.

Conclusions and policy implications

Long term exposure to ambient PM_{2.5} at relatively high ambient concentrations in China is positively and almost linearly associated with an increased risk of stroke, overall and by subtype. These findings are meaningful for both air quality related policy development and stroke prevention, not only in China, but also in other low and middle income countries.

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Contributors: KH and FLiang are joint first authors. DG contributed to the study design, implementation, manuscript discussion and critical revision. YLiu mainly contributed to PM_{2.5} exposure assessment protocol and methodology. JH contributed to critical revision of the manuscript. KH, FLiang, XY, FLiu, and JL analyzed the data and KH drafted the manuscript. FLiang and QX contributed to the PM_{2.5} exposure assessment. JL, QX, JChen, XLiu, JCao, CS, LY, FLu, XPW, LZ, XGW, YLi, DH, JH, and XLu collected and interpreted the data. All authors have read and approved the final manuscript. YLiu and DG are the guarantors. The corresponding author attests that all listed authors meet authorship criteria and that no others meeting the criteria have been omitted.

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Data sharing: No additional data available.

The lead author (DG) affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned have been explained.

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Supplementary information: details of cohort and additional table and figure