FTO genotype and weight loss: systematic review and meta-analysis of 9563 individual participant data from eight randomised controlled trials

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ABSTRACT

OBJECTIVE
To assess the effect of the FTO genotype on weight loss after dietary, physical activity, or drug based interventions in randomised controlled trials.

DESIGN
Systematic review and random effects meta-analysis of individual participant data from randomised controlled trials.

DATA SOURCES
Ovid Medline, Scopus, Embase, and Cochrane from inception to November 2015.

ELIGIBILITY CRITERIA FOR STUDY SELECTION
Randomised controlled trials in overweight or obese adults reporting reduction in body mass index, body weight, or waist circumference by FTO genotype (rs9939609 or a proxy) after dietary, physical activity, or drug based interventions. Gene by treatment interaction models were fitted to individual participant data from all studies included in this review, using allele dose coding for genetic effects and a common set of covariates. Study level interactions were combined using random effect models. Metaregression and subgroup analysis were used to assess sources of study heterogeneity.

RESULTS
We identified eight eligible randomised controlled trials for the systematic review and meta-analysis (n=9563). Overall, differential changes in body mass index, body weight, and waist circumference in response to weight loss intervention were not significantly different between FTO genotypes. Sensitivity analyses indicated that differential changes in body mass index, body weight, and waist circumference by FTO genotype did not differ by intervention type, intervention length, ethnicity, sample size, sex, and baseline body mass index and age category.

CONCLUSIONS
We have observed that carriage of the FTO minor allele was not associated with differential change in adiposity after weight loss interventions. These findings show that individuals carrying the minor allele respond equally well to dietary, physical activity, or drug based weight loss interventions and thus genetic predisposition to obesity associated with the FTO minor allele can be at least partly counteracted through such interventions.

SYSTEMATIC REVIEW REGISTRATION
PROSPERO CRD42015015969.

Introduction

The epidemic of obesity, together with its associated health burden, continues to spread globally.1 With an estimated 2.1 billion adults now overweight or obese,2 there is an urgent need to develop more effective strategies for preventing and managing obesity.3 Genotype plays an important role in the development of obesity,4 and recent genome wide association studies have identified multiple loci associated with body mass index5 and distribution of body fat.6 Specifically, 97 loci have been identified as accounting for about 2.7% of variation in body mass index,8 of which the fat mass and obesity associated (FTO) gene,9 melanocortin 4 receptor (MC4R) gene,10 and transmembrane protein 18 (TMEM18) gene11 have shown the strongest associations. Supplementary appendix 1 summarises the details of key genes associated with body mass index. However, so far the FTO gene explains the largest amount of the genetic variance in obesity traits over the lifespan.9 Those homozygous for the FTO (rs9939609) minor allele weigh on average 3 kg more and have a 1.7-fold increased odds of being obese compared with those homozygous for the lower risk allele.12

To date, several intervention studies have explored the interaction between FTO genotype and lifestyle changes on adiposity.13-19 A study in 742 obese adults found that after a two year dietary intervention, those with the FTO minor allele had a 1.5 kg greater weight loss on a high protein diet but not on a low protein diet.16 Similarly, a six month moderate intensity
exercise intervention in 105 obese women resulted in up to twofold greater weight loss in carriers of the FTO minor allele compared with non-carriers. In contrast, a one year lifestyle intervention in 3548 adults showed smaller changes in adiposity in those with the FTO minor allele compared with those without, and a lifestyle intervention in 502 participants found no effect of the FTO genotype on weight loss. Moreover, an analysis of data from the Diabetes Prevention Programme (DPP) (n=1824) and the Action for Health in Diabetes study (Look AHEAD; n=3906), showed no effect of FTO genotype on weight loss.

A recent meta-analysis by Xiang et al assessed the effect of the FTO genotype on weight change across several randomised and non-randomised intervention studies. In the randomised controlled trials the authors averaged FTO effects on weight change across both intervention and control arms of the studies under investigation and concluded that homozygous carriers of the FTO minor allele lost 0.44 kg (95% confidence interval 0.09 to 0.79 kg); P=0.015) more weight than non-carriers. However, such averaging across study arms can be justified only in the absence of single nucleotide polymorphisms by treatment interactions, which were not tested by Xiang et al. This limitation, combined with the authors' use of summary level data to assess the influence of participant level characteristics—e.g., age and body mass index at baseline, means that the findings are potentially misleading and subject to aggregation bias.

We carried out a systematic review and meta-analysis to provide a critical analysis of the evidence that the FTO genotype influences changes in obesity related outcomes in randomised weight loss interventions. In particular, we investigated whether the FTO genotype (rs9939609 or a proxy) predicted the magnitude of weight loss in response to a randomised weight loss programme. This analysis also employed individual participant data analyses to assess the role of participant level covariates.

Methods

Our systematic review was conducted according to Cochrane and the Centre for Reviews and Dissemination guidelines and is reported according to the preferred reporting items for systematic reviews and meta-analyses guidelines. The protocol was registered with PROSPERO, the International Prospective Register of Systematic Reviews (CRD42015015969). An initial systematic literature search was undertaken to identify eligible studies for inclusion in the meta-analysis. Once eligible studies were identified, we invited the corresponding authors of those studies to contribute individual participant data or to undertake a prescribed statistical plan (if individual participant data were not available for sharing). If authors did not reply, we excluded the corresponding studies from the meta-analysis but reported the study characteristics.

Search strategy

We undertook an electronic search to identify intervention studies reporting weight loss by FTO genotype (rs9939609 or a proxy) after a dietary, physical activity, or drug based intervention. Ovid Medline (www.nlm.nih.gov/bsd/pmrresources.html), Embase (www.embase.com), and Scopus (www.scopus.com) were searched systematically for studies published from inception to November 2015. To determine other studies potentially eligible for inclusion we hand searched the reference lists of identified publications and previously published related systematic reviews. The search strategy involved combining words from the concepts of genes, weight loss, and highly sensitive search filters for identifying randomised controlled trials for Ovid Medline and Embase (see supplementary appendix 2 for the detailed search strategy).

Study selection criteria

Included studies were randomised intervention studies in overweight or obese (body mass index ≥25) participants aged 18 years or more designed to induce weight loss (either as a primary or a secondary outcome) and that reported change in adiposity indices (body mass index, body weight, or waist circumference) by FTO genotype (rs9939609 or a proxy). Studies in non-overweight or non-obese participants (body mass index <25) and children (<18 years) were excluded. Only publications with an English language abstract were included. We included studies on men and women with or without health risk factors (such as raised blood pressure, abnormal lipid levels, and metabolic syndrome).

Data extraction

Two reviewers (KML and JL) independently assessed publications for eligibility. The decision to include studies was hierarchical and made initially on the basis of the study title and abstract; when a study could not be excluded with certainty, the full text was obtained for evaluation. Discrepancies between reviewers were resolved through discussion with a third reviewer (CCM), and a consensus approach was used. Information extracted from studies included study design (intervention type (dietary, physical activity, or drug intervention), length of follow-up, and country); participant characteristics (age, sex, and ethnicity); description of measurement methods; and information to assess the risk of bias. Study quality was assessed using the Cochrane risk of bias tool (see supplementary appendix 3). Two reviewers extracted data, one independently and the second confirming or completing information required.

Statistical analysis

Statistical analyses were conducted using Stata 16.0 software (Stata, College Station, TX) and Review Manager (RevMan Version 5.1 for Windows Copenhagen; Nordic Cochrane Centre, Cochrane Collaboration, 2015). A two step individual participant data analysis was undertaken, whereby individual data were first analysed separately by study and subsequently aggregated using weighted meta-analyses. A common analytical plan was used for all studies under investigation, with covariate coding matched across studies as closely as possible. We checked summary level data in the log
files returned by researchers for consistency with the prescribed statistical plan. To ensure consistency in the pattern of results we compared all individual participant data analyses with the published studies.

Outcomes of interest were change in body mass index, body weight, or waist circumference between baseline and follow-up (change calculated as follow-up measurement minus baseline measurement). Linear regression analyses assessed change in the dependent variable (change in body weight, body mass index, or waist circumference) by FTO genotype (independent variable) separately by study arm, with FTO genotype coded using an allele-dose model (ie, number of copies of the minor allele). We adjusted arm-specific models for age (continuous), sex, baseline outcome (body mass index, body weight, or waist circumference; continuous), ethnicity (categorical), country or centre (categorical), socioeconomic status (categorical), physical activity (continuous where possible), and smoking (categorical; see supplementary appendix 4). Participants were excluded from the analyses if they did not have complete data for all outcomes and covariates.

Table 2 presents two sets of regression coefficients. The first set of coefficients is arm specific and captures within arm change in the outcome during the course of the study for each copy of the FTO minor allele. Both treatment and control arms are included in these calculations. The second set of coefficients is intervention specific and captures mean differences in FTO allelic effects between treatment and control arms, where a negative coefficient indicates that individuals carrying a minor allele had a greater reduction in the outcomes of interest (weight, body mass index, or waist circumference) after the intervention than those without the minor allele. So, for studies with K treatment arms and a single control arm, K+1 coefficients are used to capture FTO effects on within arm change scores (set 1), and K coefficients are used to capture FTO effects on treatment versus control differences in change scores (set 2). Importantly, the first set of coefficients is estimated from separate groups of study participants, rendering them statistically independent and facilitating standard error calculations for the second set of coefficients that capture the interaction effects of gene x treatment. However, the second set of coefficients uses a common control group for each study, which introduces dependence at the study level.

This second set of regression coefficients from table 2 was used to evaluate differences in FTO allelic effects on the outcomes of interest between intervention and control groups. We used random effects models to estimate the pooled effect sizes and to account for both sampling error and between study variation in population. Meta-estimates were weighted by the inverse of the variance of the effect size (that is, 1 divided by variance), where variance took into account the two potential sources of variation (ie, variance within and between studies).

For studies with more than one active treatment arm and a single control arm, including all differences in treatment versus control in the meta-analysis shown in figure 2 would have invalidated the standard error calculations programmed into our meta-analytical packages, as these packages assume independent intervention effects, not ones based on comparing different treatment arms with a shared control group. For such multi-arm studies, we used a fixed effect, inverse variance meta-analysis to produce an overall estimate of FTO effects across all treatment arms, in the absence of within study heterogeneity. We then compared that overall FTO effect across treatment arms with the FTO effect on the placebo arm, thus creating a single intervention versus control comparison for each study under consideration that satisfied statistical independence assumptions underlying standard error calculations. This method is recommended by Higgins et al30 to avoid excessive weightings from “double counts” originating from the control group shared by the multiple treatment arms.

A further complication arose because our meta-analytical packages required that intervention effects be entered in terms of mean differences, their standard deviation, and sample size. As the meta-analyses returned point estimates and their standard errors instead, we arbitrarily fixed our effective sample size for the combined intervention effect to the sum of the sample sizes of all K treatment groups in each study, and then back calculated the standard deviation required for the combined intervention effect to have the right level of precision. As the forest plot depends solely on the product of the standard deviation with the square root of the reported sample size, the arbitrariness in these calculations did not affect the validity of the forest plot itself.

Heterogeneity between studies was evaluated using the I2 test and Galbraith plots31 and the 95% confidence interval for I2 calculated using the method of Higgins et al.32 33 Small study effects were appraised by visual inspection of funnel plots of effect size against the standard error, with asymmetry assessed formally with Egger’s test, chosen over Begg’s test for its greater specificity and power,34 where a P value less than 0.1 was considered as significant.35

To explore potential sources of heterogeneity, we conducted moderation testing using intervention type (diet and diet and/or exercise), intervention length (≤6 months and >6 months), age (<50 years and ≥50 years), sex (binary), body mass index (<30 and ≥30), and race/ethnicity (white, black or African-American, and Hispanic) as putative categorical moderators. Between stratum comparisons in table 3 were used to test the hypothesis that the relation between FTO genotype and intervention effects on change in obesity related outcomes differed across levels of the moderator, whereas within stratum tests helped elucidate the nature of the interaction, if present. Variables employed as stratification factors were removed as covariates from the corresponding regression model. In addition, we investigated intervention length as a continuous variable in moderation analyses.

As a secondary analysis, we recoded FTO genotypes using a dominant model, where participants with one (AT) and two (AA) copies of the minor allele were grouped together and compared with those with no copies of the minor allele (TT).
Given that only one study included drug based interventions, we run a further sensitivity analysis where only the lifestyle intervention treatment arm was included. In a final set of sensitivity analysis, we removed from the analyses the study that contributed the most weight to the analyses (Look AHEAD).

To investigate whether changes in obesity outcomes were driven by associations between obesity outcomes at baseline and FTO genotype, additional linear regression analyses (presented in supplementary appendix 8) assessed relations between baseline values of the dependent variable (body weight, body mass index, or waist circumference) and FTO genotype (independent variable), where the FTO genotype was coded using an allele-dose model. These analyses were adjusted for age, sex, ethnicity, country or centre, socioeconomic status, physical activity, and smoking and aggregated using a fixed effects, inverse variance meta-analysis.

### Patient involvement

No patients were involved in setting the research question or the outcome measures, nor were they involved in developing plans for design or implementation of the study. No patients were asked to advise on interpretation or writing up of results. There are no plans to disseminate the results of the research to study participants or the relevant patient community.

### Results

Searches yielded 2247 publications after removal of duplicates. One additional record was identified through other sources. Following screening (n=2248), 2220 publications were excluded based on the titles and abstracts and 28 full text articles were assessed for eligibility (fig 1). Full text articles were excluded from the qualitative synthesis because either they presented data from the same study (n=5) or they did not report results for FTO genotype (n=12). A total of 11 studies were eligible for inclusion in the qualitative synthesis. Three studies were excluded from the quantitative synthesis (meta-analysis) because the authors did not reply to our request for access to data. A total of eight studies were included in the present meta-analysis. Individual participant data were analysed from all eight studies: five studies provided individual participant data and three studies performed our prescribed statistical analysis and provided summary level data.

### Study characteristics

The pooled population of the 11 studies eligible for inclusion in the qualitative synthesis comprised 10 000 adults (table 1). After exclusion of participants without full data for outcomes and covariates (see supplementary appendix 5), the pooled population of adults included in the quantitative synthesis (meta-analysis) was 9563. The mean age at baseline was 51.6 (range 28-74) years and the mean body mass index was 32.2 (23.8-43.2). All studies provided data for single nucleotide polymorphism rs9939609, with the exception of the Dose Response to Exercise in Women aged 45-75 years (DREW) study, where rs8050136 was studied and is known to be in high linkage disequilibrium with rs9939609 (HapMap US residents of European ancestry: $r^2$>0.84). Nine studies were in men and women and two studies in women only. Five studies were of dietary interventions, three of dietary and exercise based interventions, one of an exercise only intervention, and two of a drug and/or lifestyle intervention. The duration of follow-up ranged from eight weeks to three years. Studies were conducted in North America, South America, and Europe; six studies were in white participants and four in mixed populations (table 1).

### Study quality and small study effects

No studies were excluded from the meta-analysis based on quality assessment (see supplementary appendix 3). Visual inspection of funnel plots (see supplementary appendix 6) and Galbraith plots (see supplementary appendix 7) did not identify any systematic small study effects. Similarly, no significant bias was observed for body mass index (P=0.55), body weight (P=0.69), or waist circumference (P=0.39).

### Assessment of gene x treatment interaction effects on change in obesity related outcomes

Table 2 summarises the minor allele effects (coefficient and standard error) on obesity outcomes after the weight loss intervention for each study and study arm. Table 2 also presents differences in minor allele effects between treatment and control arms, indicative of gene by treatment interactions. Mean treatment versus control differences in body mass index change for each copy of the FTO minor allele ranged from −0.33 (95% confidence interval −1.13 to 0.47) after six month follow-up in DREW to 0.14 (−0.06 to 0.35) after one year in Look AHEAD. Body mass index was reduced by an
additional −0.02 (−0.13 to 0.09); \( \beta_0 \) (95% confidence interval 0 to 68; \( P = 0.69 \), body weight by −0.04 (−0.34 to 0.26) kg; \( \beta_0 \) (0 to 68; \( P = 0.78 \)), and waist circumference by −0.06 (−0.43 to 0.31) cm; \( \beta_0 \) (0 to 68; \( P = 0.75 \)) for each copy of the FTO minor allele after weight loss intervention compared with that expected under naturalistic change in the control group (fig 2). When using a dominant FTO genotype model (AA/AT vs TT), body mass index, body weight, and waist circumference were reduced by 0.05 (95% confidence interval −0.21 to 0.11); \( P = 0.558 \), 0.15 (−0.60 to 0.30) kg; \( P = 0.524 \), and 0.22 (−0.77 to 0.33) cm; \( P = 0.437 \) after intervention in carriers of the minor allele compared with non-carriers.

**Stratified and sensitivity analyses**

The relation between FTO genotype and differences between treatment and control in change in obesity related outcomes (body mass index, body weight, and waist circumference) after weight loss interventions was not influenced by study type, study length, sex, race/ethnicity, or body mass index category (table 3). Change in body mass index for each copy of the FTO minor allele was significant in participants aged 50 or more years (\( \beta = -0.23 \) (95% confidence interval −0.44 to –0.22); \( \beta_0 \) (95% confidence interval 0 to 75); \( P = 0.03 \)), but not in participants aged less than 50 years (\( \beta = -0.04 \) (−0.26 to 0.18); \( \beta_0 \) (0 to 75); \( P = 0.75 \)); however, the corresponding moderation test failed to attain significance (\( P = 0.21 \)). In addition, intervention length modelled as a continuous variable did not significantly affect the relation between FTO genotype and differences between treatment and control in change in obesity related outcomes (body mass index, body weight, and waist circumference) after weight loss interventions.

Given the presence of a gene x treatment interaction on waist circumference in table 3 between metformin and control arms in study DPP, we investigated exclusion of the metformin and troglitazone arms from DPP. This did not change the pattern of results (data not shown).

After removal of Look AHEAD, the largest study, overall effect sizes for body mass index, body weight, and waist circumference in figure 2 changed to −0.09 (−0.22 to 0.04; \( P = 0.17 \), −0.21 (−0.56 to 0.14) kg; \( P = 0.25 \), and −0.01 (−0.45 to 0.42) cm; \( P = 0.96 \), respectively, but failed to attain statistical significance.

**Baseline association testing**

To facilitate better interpretation of any effect of FTO genotype on weight loss after an intervention, we investigated the association between FTO genotype and obesity measures at baseline. FTO genotype differences in body mass index, body weight, and waist circumference at baseline for each study are presented in supplementary appendix 8. Collectively, for each copy of the FTO minor allele, body mass index, body weight, and waist circumference increased by 0.31 (0.14 to 0.47); \( P < 0.001 \), 0.89 (0.45 to 1.32) kg; \( P < 0.001 \), and 0.63 (0.28 to 0.98) cm; \( P < 0.001 \), respectively.

**Discussion**

The present systematic review and meta-analysis used individual participant data to investigate the differential effect of FTO genotype on response to weight loss intervention in randomised controlled trials. Our meta-analysis of eight studies involving 9563 adults showed that carriage of the FTO minor allele does not influence change in adiposity measures in response to weight loss intervention, compared with treatment controls. Since we observed a strong association of the FTO minor allele with greater adiposity at baseline, this neutral effect of FTO minor allele on weight loss is an important finding for the development of effective weight loss interventions in the context of the global epidemic of obesity. Specifically, people who carry obesity risk FTO genotypes respond equally well to weight loss treatment.

Our finding of the lack of a FTO genotype by treatment interaction on change in obesity related outcomes expands on a recent meta-analysis on the effect of FTO genotype on weight loss,\(^{21}\) which analysed FTO genotypic effects on weight change averaged across highly effective treatment arms and control arms that were designed to produce minimal to no weight change. Unlike Xiang et al, we included only randomised controlled trials, which provide stronger evidence than

**Table 1 Characteristics of studies included in qualitative synthesis (n=10 000) and meta-analysis (n=9563)**

<table>
<thead>
<tr>
<th>Study</th>
<th>No of participants</th>
<th>SNP</th>
<th>MAF</th>
<th>Type</th>
<th>Length</th>
<th>Region</th>
<th>Ethnicity</th>
<th>Mean (SD)</th>
</tr>
</thead>
<tbody>
<tr>
<td>DPP(^{20})</td>
<td>2835</td>
<td>962</td>
<td>1873</td>
<td>rs9939609</td>
<td>40.9</td>
<td>Drug and lifestyle</td>
<td>1 year</td>
<td>North America</td>
</tr>
<tr>
<td>DREW</td>
<td>278</td>
<td>278</td>
<td>rs8050136(^a)</td>
<td>41.0</td>
<td>Exercise</td>
<td>6 months</td>
<td>North America</td>
<td>Mixed</td>
</tr>
<tr>
<td>Finnish DPS(^{21})</td>
<td>264</td>
<td>97</td>
<td>167</td>
<td>rs9939609</td>
<td>41.5</td>
<td>Diet and exercise</td>
<td>3 years</td>
<td>Europe</td>
</tr>
<tr>
<td>Food4Me(^{26})</td>
<td>671</td>
<td>313</td>
<td>358</td>
<td>rs9939609</td>
<td>44.3</td>
<td>Diet and exercise</td>
<td>6 months</td>
<td>Europe</td>
</tr>
<tr>
<td>Look AHEAD(^{25})</td>
<td>3637</td>
<td>1601</td>
<td>2036</td>
<td>rs9939609</td>
<td>44.5</td>
<td>Diet</td>
<td>1 year</td>
<td>North America</td>
</tr>
<tr>
<td>POUNDS LOST(^{26})</td>
<td>600</td>
<td>240</td>
<td>360</td>
<td>rs9939609</td>
<td>45.1</td>
<td>Diet</td>
<td>2 years</td>
<td>North America</td>
</tr>
<tr>
<td>PREDIMED(^{21})</td>
<td>735</td>
<td>335</td>
<td>400</td>
<td>rs9939609</td>
<td>41.6</td>
<td>Diet</td>
<td>3 years</td>
<td>Europe</td>
</tr>
<tr>
<td>NUGENOB</td>
<td>543</td>
<td>136</td>
<td>407</td>
<td>rs9939609</td>
<td>41.5</td>
<td>Diet</td>
<td>10 weeks</td>
<td>Europe</td>
</tr>
<tr>
<td>Ramos et al(^{27})</td>
<td>86</td>
<td>86</td>
<td>rs9939609</td>
<td>NE</td>
<td>Drug</td>
<td>6 months</td>
<td>South America</td>
<td>White</td>
</tr>
<tr>
<td>De Luis et al(^{28})</td>
<td>305</td>
<td>80</td>
<td>225</td>
<td>rs9939609</td>
<td>44.1</td>
<td>Diet</td>
<td>3 months</td>
<td>Europe</td>
</tr>
<tr>
<td>MOVE(^{29})</td>
<td>46</td>
<td>33</td>
<td>13</td>
<td>rs9939609</td>
<td>NE</td>
<td>Diet</td>
<td>8 weeks</td>
<td>North America</td>
</tr>
</tbody>
</table>

\( SNP = \) single nucleotide polymorphism; \( MAF = \) minor allele frequency for all randomised participants who provided genetic consent and whose DNA data passed quality control procedures; \( BMI = \) body mass index; \( NE = \) not estimable based on data in published manuscript.

\(^a\)SNP rs8050136 was in high linkage disequilibrium with rs9939609 (HapMap US residents of European ancestry: \( r^2 > 0.84 \).
<table>
<thead>
<tr>
<th>Study and study arm</th>
<th>No</th>
<th>Minor allele effects*</th>
<th>Treatment v control differences in minor allele effects†</th>
<th>Minor allele effects*</th>
<th>Treatment v control differences in minor allele effect†</th>
<th>Minor allele effects*</th>
<th>Treatment v control differences in minor allele effect†</th>
</tr>
</thead>
<tbody>
<tr>
<td>Diet (T1): lifestyle</td>
<td>1144</td>
<td>-0.11 (0.22)</td>
<td>-0.10 (0.57 to 0.37)</td>
<td>-0.59 (0.51)</td>
<td>-0.19 (1.51 to 1.13)</td>
<td>0.77 (0.70)</td>
<td>1.07 (0.47 to 2.61)</td>
</tr>
<tr>
<td>Diet (T2): PN: diet</td>
<td>1384</td>
<td>-0.11 (0.17)</td>
<td>-0.11 (0.49 to 0.28)</td>
<td>-0.11 (0.47)</td>
<td>-0.14 (1.21 to 0.93)</td>
<td>-0.19 (0.51)</td>
<td>0.12 (0.18 to 1.42)</td>
</tr>
<tr>
<td>Diet (T3): PN: diet+phenotype</td>
<td>1474</td>
<td>-0.25 (0.18)</td>
<td>-0.25 (0.65 to 0.15)</td>
<td>-0.17 (0.62)</td>
<td>-0.61 (1.75 to 0.53)</td>
<td>-0.14 (0.63)</td>
<td>-0.03 (-0.24 to 0.57)</td>
</tr>
<tr>
<td>Combined treatment arm</td>
<td>277</td>
<td>-0.16 (0.14)</td>
<td>-0.15 (0.44 to 0.14)</td>
<td>-0.29 (0.52)</td>
<td>-0.32 (-1.12 to 0.48)</td>
<td>-0.25 (0.70)</td>
<td>0.06 (-0.93 to 1.09)</td>
</tr>
<tr>
<td>Combined treatment arm</td>
<td>278</td>
<td>0.01 (0.10)</td>
<td>-0.02 (0.09 to 0.28)</td>
<td>0.05 (0.32)</td>
<td>-0.03 (0.87)</td>
<td>-0.70 (0.51)</td>
<td>-0.11 (1.51 to 1.13)</td>
</tr>
<tr>
<td>Look AHEAD*</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Diet: low fat</td>
<td>1806</td>
<td>-0.11 (0.06)</td>
<td>-0.27 (0.17)</td>
<td>-0.59 (0.51)</td>
<td>-0.19 (1.51 to 1.13)</td>
<td>0.77 (0.70)</td>
<td>1.07 (0.47 to 2.61)</td>
</tr>
<tr>
<td>Control</td>
<td>278</td>
<td>-0.16 (0.11)</td>
<td>-0.15 (0.35 to 0.15)</td>
<td>-0.41 (0.38)</td>
<td>-0.05 (-0.85 to 0.76)</td>
<td>0.01 (0.78)</td>
<td>0.68 (-0.52 to 1.88)</td>
</tr>
<tr>
<td>Combined treatment arm</td>
<td>278</td>
<td>0.01 (0.10)</td>
<td>-0.02 (0.09 to 0.28)</td>
<td>0.05 (0.32)</td>
<td>-0.03 (0.87)</td>
<td>-0.70 (0.51)</td>
<td>-0.11 (1.51 to 1.13)</td>
</tr>
<tr>
<td>Combined treatment arm</td>
<td>278</td>
<td>0.01 (0.11)</td>
<td>-0.02 (0.09 to 0.28)</td>
<td>0.05 (0.32)</td>
<td>-0.03 (0.87)</td>
<td>-0.70 (0.51)</td>
<td>-0.11 (1.51 to 1.13)</td>
</tr>
</tbody>
</table>

*Values represent coefficient and standard error for change in BMI, body weight, and waist circumference by FTO minor allele (allele-dose model was employed and coded in terms of copies of minor allele: 0, 1, 2). Minor allele effects for treatment (combined across treatment arms if applicable) and control were used in meta-analysis in figure 2.

†Values represent coefficient and 95% confidence interval for differential change in BMI, body weight, and waist circumference between treatment and control arms by FTO minor allele (allele-dose model was employed and coded in terms of copies of minor allele: 0, 1, 2). Treatment (combined across treatment arms if applicable) versus control differences in minor allele effects for each study correspond to mean difference in figure 2.
one treatment arm was present, values represent combined effects across treatment arms. A meta-analysis of 9563 adults. Covariates across all studies and to perform analyses such as body mass. We also used individual participant data, and a large sample size—are used, terms of demographics and baseline characteristics such as body mass. We also used individual participant data, which allowed us to adjust for the same set of covariates across all studies and to perform analyses stratified by participant characteristics and not study level summaries. This also facilitated comparisons between additive and dominant FTO genotype models. Moreover, our sample size was larger, which improved the power of our meta-analysis. Our findings suggest that when robust intervention design and analysis—including gene by treatment arm interactions, individual participant data, and a large sample size—are used, there is no evidence that FTO genotype affects weight loss in response to lifestyle or dietary interventions.

The present meta-analysis included two studies, Look AHEAD and DPP, which together contributed.
### Table 3: Stratified analysis to identify potential moderators of the relation between FTO genotype and weight loss

<table>
<thead>
<tr>
<th>Variables</th>
<th>BM mean difference*</th>
<th>P value</th>
<th>Adjusted mean difference (95% C)</th>
<th>P value</th>
<th>Intercept</th>
<th>95% C</th>
<th>P value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Diet</td>
<td>0.15 (−0.21, 0.51)</td>
<td>0.83</td>
<td>0.06 (−0.16, 0.28)</td>
<td>0.86</td>
<td>0.08</td>
<td>0.28</td>
<td>0.42</td>
</tr>
<tr>
<td>Age (years)</td>
<td>≥6 months (n=7)</td>
<td>0.09 (−0.28 to 0.33)</td>
<td>0.40</td>
<td>0.01 (−0.14 to 0.16)</td>
<td>0.95</td>
<td>0.01</td>
<td>0.95</td>
</tr>
<tr>
<td></td>
<td>&lt;6 months (n=7)</td>
<td>0.06 (−0.10 to 0.02)</td>
<td>0.19</td>
<td>0.02 (−0.18 to 0.03)</td>
<td>0.03</td>
<td>0.02</td>
<td>0.12</td>
</tr>
<tr>
<td>Race/ethnicity††: White (n=7)</td>
<td>0.18 (−0.32 to 0.69)</td>
<td>0.48</td>
<td>0.10 (−0.43 to 0.65)</td>
<td>0.94</td>
<td>0.10</td>
<td>0.94</td>
<td></td>
</tr>
<tr>
<td></td>
<td>Black/African-American (n=7)</td>
<td>0.01 (−0.38 to 0.40)</td>
<td>0.78</td>
<td>0.23 (−1.39 to 0.92)</td>
<td>0.69</td>
<td>0.23</td>
<td>0.69</td>
</tr>
</tbody>
</table>

67% (n=6472) of the total sample. Independently, neither study identified any association between FTO genotype and weight loss, which was consistent with parallel and pooled analysis of both studies and with our overall findings. In our sensitivity analyses, omission of the largest study (Look AHEAD) did not change the pattern of results, thus supporting the robustness of our findings.

Although we identified a greater reduction in body mass index for each minor allele of FTO genotype in participants aged more than 50 years, this effect was not consistent across body weight and waist circumference outcomes, suggesting that this is likely to be a chance finding. Nevertheless, recent meta-regression estimates from twin studies suggest that heritability of body mass index was 0.07 (P=0.001) higher in children than in adults. Further research into possible subgroup effects of FTO variants and other obesity related genes is warranted to confirm or refute our findings.

Although there is good evidence that the duration of the weight loss trial influences weight lost and regained, with greatest loss occurring at six months followed by gradual regain, we did not identify any effect of study duration on the relation between FTO genotype and weight loss. Behaviour change strategies for weight management seem to have limited effect in preventing relapse. In a pan-European study of 742 adults, multiple single nucleotide polymorphisms were investigated for effects on weight regain over six months ad libitum diet after more than 8% weight loss, and the study investigated whether nutrient sensitive genes modified weight regain in response to glycaemic index and high protein based diets. No significant effects were found for the FTO single nucleotide polymorphisms, and although significant effects were found for several other single nucleotide polymorphisms, associations were no longer significant after adjustment for multiple testing. Furthermore, although a more recent study found evidence for the specific role of FTO genotype in predicting maintenance of body weight, results were based on a smaller (n=128), women only sample, and so the findings should be interpreted with caution. Given the lack of effective strategies for preventing weight regain and the evidence for an interaction between FTO genotype and physical activity on risk of obesity, further research into the role of obesity related genes and their interactions with diet and physical activity on long term weight management are needed. Although the role of FTO, and other obesity associated genes, in weight maintenance was outside the scope of the present meta-analysis, it is an important consideration for maximising the long term health benefits of weight loss and should be considered in future research.

Weight loss in response to a lifestyle intervention requires negative energy balance to be sustained over a considerable period. This process is complex and involves behaviour changes in either or both, food consumption (energy intake) and physical activity (energy expenditure). The mechanisms through which the FTO genotype could influence such processes are not known and might be different from those through

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**Footnotes:**
- Values for mean difference represent linear regression coefficients for differential change between treatment (combined) and control arms in BMI, body weight, or waist circumference by FTO genotype (allele-dose model) and 95% confidence interval.
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Weight loss in response to a lifestyle intervention requires negative energy balance to be sustained over a considerable period. This process is complex and involves behaviour changes in either or both, food consumption (energy intake) and physical activity (energy expenditure). The mechanisms through which the FTO genotype could influence such processes are not known and might be different from those through
which the FTO genotype influences the development of excess adiposity.

**Strength and limitations of this review**

A strength of this meta-analysis is that we jointly analysed individual participant data from eight randomised controlled trials using a standardised analysis strategy. Application of this standardised statistical analysis mitigated differences in study methodologies, utilised data independent of statistical significance or how results were reported initially, and improved the overall reliability of the results. Moreover, our analysis included only randomised controlled trials, considered the ideal study design by avoiding many biases and confounding effects, and this approach allowed us to observe whether any differential effect of FTO genotype on weight loss is specific to weight loss interventions by comparing weight loss between participants randomised to treatment arms and those randomised to the control arm. A further strength of our study was our ability to assess the effect of the FTO genotype on multiple measures of adiposity. Lastly, our study addresses the genetics and gene-lifestyle interactions of weight loss, a problem of public health importance, and fills a major gap in the literature on the putative role of FTO genotype in modulating weight loss in response to weight loss interventions.

An important limitation is that we evaluated the effect of FTO genotype only and, given that obesity risk and weight loss is influenced by multiple genes, the effect of other obesity related genes, such as MC4R and TMEM18, on weight loss in response to intervention remains to be determined. None the less, our findings are in line with those from the DPP and Finnish DPS trials, where participants with greater genetic risk for several known type 2 diabetes benefitted from lifestyle intervention equally well compared with those with lower risk. Furthermore, we found a relatively small number of randomised controlled trials reporting weight loss by FTO genotype, all of which originated from North America and Europe, with predominantly white participants, which limited our power to investigate differential effects between ethnicities. Given evidence that the relation between FTO genotype and obesity varies by ethnicity, further well powered studies on this topic in ethnically different populations are warranted. In addition, we applied the same body mass index cut points for adiposity classification in Asian and other individuals, but, because Asians made up less than 5% of our total sample, this is unlikely to have impacted on our results. Given the evidence that both obesity risk and weight loss are modified by multiple genetic variants, our findings for FTO genotype should not be considered in isolation. A further limitation is that the present analysis did not account for the diversity of intervention designs, where the type of diet or exercise intervention may induce differential weight loss through effects on metabolism, appetite, and thermogenesis although we found no interaction between intervention type (diet versus exercise) and FTO genotype on weight loss. In addition, weight loss interventions may be biased by higher dropout rates than in other randomised controlled trials, which may have influenced our ability to identify a relation between FTO genotype and weight loss. Lastly, the authors of three studies identified in the systematic review did not agree to provide individual participant or aggregate data and thus were excluded from the present analyses. These studies were small (total participants n=441) and thus their exclusion is unlikely to have affected our conclusions.

**Implications of the findings**

We found that the FTO genotype had no detectable effect on weight loss in overweight and obese adults in response to intervention. Importantly, our findings show that the genetic predisposition to obesity associated with the FTO minor allele can be at least partly counteracted through dietary, exercise, or drug based weight loss interventions and that those carrying the minor allele respond equally well to such interventions. Moreover, our results suggest that screening for the FTO genotype in routine clinical work would not predict weight loss success. Future public health strategies for the management of obesity should aim to induce long term improvements in lifestyle behaviours, principally eating patterns and physical activity, since these will be effective in achieving sustained weight loss irrespective of FTO genotype.

**Conclusions**

This systematic review and meta-analysis of individual participant data reveals that carriage of the FTO minor allele, associated with risk of obesity in the general population as well as baseline adiposity in the present study, was not associated with changes in body mass index, body weight, or waist circumference in response to weight loss intervention. Our findings show that weight loss in those carrying the FTO minor allele is similar to the rest of the population after dietary, exercise, or drug based interventions. Future studies should investigate the possibility that a panel of genetic variants, including other FTO genotypes, may modulate weight loss in obese people in response to lifestyle and other interventions.

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Supplementary file: supplementary appendices