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# Prevalence and clinical profile of microcephaly in South America pre-Zika, 2005-14: prevalence and case-control study

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## ABSTRACT

### OBJECTIVE

To describe the prevalence and clinical spectrum of microcephaly in South America for the period 2005-14, before the start of the Zika epidemic in 2015, as a baseline for future surveillance as the Zika epidemic spreads and as other infectious causes may emerge in future.

### DESIGN

Prevalence and case-control study.

### DATA SOURCES

ECLAMC (Latin American Collaborative Study of Congenital Malformations) database derived from 107 hospitals in 10 South American countries, 2005 to 2014. Data on microcephaly cases, four non-malformed controls per case, and all hospital births (all births for hospital based prevalence, resident within municipality for population based prevalence). For 2010-14, head circumference data were available and compared with Intergrowth charts.

### RESULTS

552 microcephaly cases were registered, giving a hospital based prevalence of 4.4 (95% confidence interval 4.1 to 4.9) per 10 000 births and a population based prevalence of 3.0 (2.7 to 3.4) per 10 000. Prevalence varied significantly between countries and between regions and hospitals within countries. Thirty two per cent (n=175) of cases were prenatally diagnosed; 29% (n=159) were perinatal deaths. Twenty three per cent (n=128) were associated with a diagnosed genetic syndrome, 34% (n=189) polymalformed without a syndrome diagnosis, 12% (n=65) with associated neural malformations, and 26% (n=145) microcephaly only. In addition, 3.8% (n=21) had a STORCH (syphilis, toxoplasmosis, other including HIV, rubella, cytomegalovirus, and

herpes simplex) infection diagnosis and 2.0% (n=11) had consanguineous parents. Head circumference measurements available for 184/235 cases in 2010-14 showed 45% (n=82) more than 3 SD below the mean, 24% (n=44) between 3 SD and 2 SD below the mean, and 32% (n=58) larger than -2 SD.

### CONCLUSION

Extrapolated to the nearly 7 million annual births in South America, an estimated 2000-2500 microcephaly cases were diagnosed among births each year before the Zika epidemic began in 2015. Clinicians are using more than simple metrics to make microcephaly diagnoses. Endemic infections are important enduring causes of microcephaly.

## Introduction

After an increase in the number of babies born with congenital microcephaly was observed in northeast Brazil in 2015 and French Polynesia in 2013-14 following epidemics of the Zika flavivirus in these populations,<sup>1,2</sup> the World Health Organization declared a public health emergency of international concern (now no longer in force) as further global spread of the virus was predicted.<sup>3</sup> The aim of a diagnosis of microcephaly at birth is to use small head circumference as an indicator of a high risk of having experienced poor in utero development of the brain.<sup>4</sup> Microcephaly is known to be caused by viral, bacterial, and parasitic infections such as rubella, cytomegalovirus, herpes simplex virus, immunodeficiency virus, toxoplasmosis, and syphilis among others<sup>5-9</sup> (as well as many non-infectious causes), so that an infectious cause of the observed increase in affected babies was plausible. Since then, more evidence has been accumulating to support a causal link,<sup>10,11</sup> including the neurotropic nature of the Zika virus and preliminary reports of epidemiological controlled studies.<sup>12-14</sup> However, the size of the risk for infected pregnant women remains unclear, and the extent of the estimated increase in microcephaly in northeast Brazil has not yet been observed in other populations affected by Zika.

Infection with the Zika virus in pregnancy has also been reported with other brain anomalies, ocular manifestations, arthrogryposis, and intrauterine growth retardation in the newborn.<sup>13,15-18</sup> These observations are leading to a broader definition of "congenital Zika syndrome"<sup>19-28</sup> in a similar way to the well established congenital rubella syndrome.<sup>29</sup> Zika may cause a wider spectrum of abnormalities of the brain,<sup>13,30</sup> which may be diagnosed only later in infancy or childhood. Microcephaly itself may become manifest later owing to restricted postnatal head growth.<sup>15</sup>

## WHAT IS ALREADY KNOWN ON THIS TOPIC

A surge in microcephaly was detected in Brazil at the end of 2015 followed by other South American countries as the Zika virus spread throughout the continent

However, inconsistent definition and registration of this congenital anomaly clouded estimation of the magnitude of increase due to Zika virus

## WHAT THIS STUDY ADDS

This study provides baseline prevalence for microcephaly in 10 South American countries in the pre-Zika period, with which prevalence during the Zika epidemic can be compared

The role of other infectious and non-infectious causes should not be ignored even in times of Zika epidemic

These data can assist both researchers and governments in coping with the Zika induced microcephaly problem

The definition of microcephaly is not uniform in the medical literature.<sup>4</sup> Birth defect surveillance systems in Europe and the United States have noted registration to be poorly standardised,<sup>31 32</sup> probably owing to the lack of an agreed definition.<sup>4</sup> The “metric” microcephaly definition (how small a head size is considered microcephalic) has changed several times in response to the Zika epidemic.<sup>33</sup>

ECLAMC (Latin American Collaborative Study of Congenital Malformations) is a case-control, hospital based congenital anomaly surveillance system in Latin America (Central and South America), which has been operating since 1967.<sup>34</sup> Here, we describe the prevalence and clinical spectrum of microcephaly in South America for the period 2005-14, before the start of the Zika epidemic in 2015, as a baseline for future surveillance as the Zika epidemic spreads and as other infectious causes may emerge in future.

## Methods

### Description of population and ECLAMC data

Between 2005 and 2014, 107 hospitals (fig 1 and supplementary table) belonging to the ECLAMC network in 10 South American countries (Argentina,



Fig 1 | Map of 107 hospitals in ECLAMC 2005-14. Hospital codes are given in supplementary table

Bolivia, Brazil, Chile, Colombia, Ecuador, Paraguay, Peru, Uruguay, and Venezuela) registered congenital anomalies in a total of 1 247 185 births. Data from Central America were not available for this period in ECLAMC. The study period predates the Zika epidemic, which started in 2015. During the 2005-14 period, participation in the ECLAMC network declined, so that 46 of the 107 hospitals were participating by 2010. Fifty seven of the hospitals were university hospitals (supplementary table), and there were also some non-university tertiary referral hospitals. Tertiary hospitals typically serve the local area while also receiving referrals of high risk pregnancies and pregnancies in which malformations have been diagnosed prenatally.

All babies, live and stillborn, with congenital anomalies are registered. Termination of pregnancy for fetal anomaly was illegal during the study period (until 2014) in all countries, except under certain very restricted conditions that would not be specifically relevant to microcephaly.

In each hospital, the paediatrician diagnoses the congenital anomalies and enters the data by choosing the diagnosis from an electronic list, which is coded automatically to ICD-8 (international classification of disease, version 8) codes with a two digit ECLAMC extension for further specification and laterality. This information is transmitted to the ECLAMC central database. All diagnostic information is collected from the maternity unit up to but not beyond discharge.

For each malformed baby born in the hospital, a liveborn non-malformed control is chosen as the next birth of the same sex in the same hospital. If the next birth is not available, the subsequent birth is selected (17.5% of controls). The participation rate in ECLAMC hospitals by mothers of both cases and controls is very high. The mothers of both cases and controls are interviewed by the ECLAMC paediatrician (usually) or by a designated nurse or medical student, depending on hospital, before they are discharged from the maternity unit, according to a standard interview schedule described elsewhere.<sup>34</sup> A data entry programme is accessed via a secure internet portal to the ECLAMC server. Before 2010 data entry was by palm held device and transmitted electronically.

The information collected from medical records for the malformed baby includes the diagnosis, whether the malformation was diagnosed prenatally, gestational age, and birth weight. The information collected from maternal interviews and analysed in this paper includes consanguinity (up to second cousin), exposure to STORCH (syphilis, toxoplasmosis, other including HIV, rubella, cytomegalovirus, and herpes simplex), influenza, diabetes, and demographic information such as residence of the mother, maternal age, and maternal education. Death in the first three days (by the time of data collection) was recorded. “Perinatal deaths” as used in this paper is the addition of stillbirths and deaths in the first three days of life. Each hospital

collects a limited set of data on all births delivered in the hospital, including the total number of live births and stillbirths for calculation of prevalence.

### Definition and classification of microcephaly

The ECLAMC operational manual defines microcephaly (ICD-8 code 743.1) as head circumference more than 3 SD below the mean ( $-3$  SD) for gestational age and sex; any other criteria used instead by hospitals should be specified as part of the diagnostic text. The ECLAMC manual does not specify the growth chart to be used by hospitals, and nor has this been specified in the diagnostic text given by paediatricians. ECLAMC protocols are reinforced by annual meetings of ECLAMC paediatricians, but the specific question of standardisation of the definition and diagnosis of microcephaly did not come up during the study period. This paper includes any diagnosis of microcephaly made by paediatricians, whether or not its diagnosis follows the ECLAMC head circumference recommendation.

Head circumference data are requested to be added to the ECLAMC data for any baby with microcephaly, collected as additional text, but these data were accessible for analysis only from 2010 when data entry was by internet portal. For the purposes of this research, we converted all available head circumference data from 2010 on, when gestational age and sex had also been specified, to z scores (specific to gestational age and sex) according to the Intergrowth charts for very preterm and preterm/term babies by using the International Standards for Size at Birth app (v1.0.5934.26288).<sup>35</sup> As gestational age is recorded in completed weeks, we used the Intergrowth modelled estimates for each gestational week+3 days. The slight discontinuity in the Intergrowth models for very preterm and preterm/term babies giving a higher growth rate at 33 weeks was not smoothed. Birth weight was also classified to z scores by the same Intergrowth charts for 2005–14.

A clinical geneticist (IMO), working centrally on the basis of the diagnostic codes and text descriptions given by the hospital paediatrician (syndromes were diagnosed by the paediatricians, not centrally), further classified microcephaly cases as follows. (A) Syndromes—A1 genetic syndromes (A1.1 chromosomal and A1.2 monogenic syndromes); A2—embryopathies (teratogenic/environmental syndromes). (B) Polymalformed microcephaly—microcephaly with other major non-neural malformations excluding A. (C) Microcephaly with associated malformations of neural origin (brain, spinal cord, ocular) excluding cases classified to A or B. (D) Isolated microcephaly—only microcephaly coded, or microcephaly with minor additional anomalies.

Categories C and D, with the addition of cases from category A in which no major non-neural associated anomalies existed (shown in the last column of table 1), were further analysed together as “neural only cases.” Analysis of “neural only cases,” similar to microcephaly now associated with Zika,<sup>22</sup> would show the frequency

of causative factors associated with such a phenotype in a non-Zika period. However, genetic syndromes or embryopathies (category A) may be underrepresented as they include only those diagnosed in the first few days of life (some microcephaly cases may have received syndrome diagnoses later than this). All cases and controls with STORCH infections had laboratory confirmed evidence of infection. Cases had in addition a STORCH embryopathy diagnosis (classified to A2 above) given by the paediatrician at the hospital. Some brain imaging and genetic syndrome diagnoses for microcephaly cases may be missing if they became available after discharge from the maternity unit. We specifically analysed the presence of arthrogryposis, as it is emerging as part of the definition of congenital Zika syndrome.<sup>17 18</sup>

### Statistical analysis

We calculated prevalence of microcephaly as the number of cases of microcephaly (live and stillborn) divided by the total number of births, and expressed it per 10 000 births.<sup>36</sup> We calculated two estimates for regions and countries—a hospital based estimate and a population based estimate. The hospital based estimate included all cases born in the participating hospitals, with a denominator of all births in the participating hospitals. As some of the participating hospitals are tertiary centres that attract high risk pregnancies or prenatally diagnosed cases, the population based estimate included only babies with microcephaly to residents of the same municipality as the hospital, with a denominator estimated as the number of births multiplied by the proportion of controls (controls of all malformed babies in the same year) who were resident in the same municipality as the hospital.

We used a multilevel Poisson random effects model to analyse secular trend in prevalence. Prevalence by region and country are shown with Poisson exact 95% confidence intervals. We analysed case-control data by conditional logistic regression with matching variables of year, hospital, and sex. For this study, to improve statistical power, we used the microcephaly matched control as well as the three subsequent non-malformed controls chosen for babies with other malformations in the same hospital and year and of the same sex. Odds ratios were adjusted for maternal age and maternal education. We analysed odds ratios for all microcephaly cases and for all neural only cases. We used Stata 12 statistical software for all analyses. The sample size of the primary case-control study allows identification of a twofold risk factor for an exposure frequency of 5% with a test power of 70% and a significance level of 5%, or an odds ratio greater than 4 with a risk factor prevalence of 1%.

### Extrapolation to South America

We extrapolated our prevalence results to all 6.8 million births in South America (2016 estimate from [www.worldpopdata.org](http://www.worldpopdata.org)), using both a simple estimate (the population based prevalence applied to all births in South America) and a country stratified

**Table 1 | Clinical classification of microcephaly cases. Values are numbers (percentages) unless stated otherwise**

Clinical classification	All microcephaly cases	Prevalence (per 10 000)	Male/female sex ratio	Prenatally diagnosed cases	Perinatal deaths*	Neural only cases
A1.1. Chromosomal syndromes:	68 (12.3)	0.55	0.51	30 (44)	52 (76)	0
Patau	41					
Edwards	14					
Other chromosomal anomaly	8					
Down	3					
Turner	1					
5p- deletion	1					
A1.2. Monogenic syndromes:	60 (10.9)	0.48	0.78	29 (48)	22 (37)	2 (3)
Holoprosencephaly†	36					
Meckel	5					
Skeletal dysplasia	3					
Penna-Shokeir I	3					
Seckel	3					
Crouzon	2					
Ictiosis	2					
Cornelia de Lange	1					
Treacher Collins	1					
Freeman-Sheldon	1					
Acrocallosal	1					
Dominant lissencephaly	1					
Ivemark	1					
A2. Embryopathies:	25 (4.5)	0.20	0.67	4 (16)	3 (12)	18 (72)
Embryopathy cytomegalovirus‡	9					
Embryopathy toxoplasmosis‡	4					
Brain disruption with dead twin	2					
Embryopathy alcohol	2					
Embryopathy syphilis‡	2					
Embryopathy rubella‡	2					
Embryopathy HIV+syphilis‡	1					
Embryopathy HIV+toxoplasmosis‡	1					
Embryopathy syphilis+toxoplasmosis‡	1					
Embryopathy rubella+syphilis‡	1					
B. Polymalformed without syndrome diagnosis	189 (34.2)	1.52	0.89	63 (33)	67 (35)	0
C. Microcephaly with associated malformations of neural origin:	65 (11.8)	0.52	0.91	31 (48)	5 (8)	65 (100)
Encephalocele	13					
Brain necrosis/calcification	9					
Hydrocephaly with calcifications, neuronal migration, or other brain anomalies	9					
Neuronal migration abnormality	7					
Corpus callosum anomaly	5					
Spina bifida	4					
Cerebellar anomaly	4					
Hydrocephaly	4					
Hydranencephaly	4					
Other brain defects	4					
An/cryptophthalmos	1					
Corneal opacity	1					
D. Isolated microcephaly	145 (26.3)	1.16	0.56	18 (12)	10 (7)	145 (100)
Total microcephaly cases	552 (100.0)	4.43	0.72	175 (31.7)	159 (28.8)	230 (41.7)

\*Stillbirths and deaths in the first three days of life.

†All holoprosencephaly without chromosomal syndrome or embryopathy were attributed to monogenic category whether or not monogenic, digenic, or copy number variation syndrome diagnosis was made.

‡STORCH (syphilis, toxoplasmosis, other including HIV, rubella, cytomegalovirus, and herpes simplex) embryopathies: six STORCH cases had other major congenital anomalies, four of which were heart defects.

estimate (the country specific, population based prevalence applied to each participating country's annual births).

#### Patient involvement

No patients were involved in setting the research question or the outcome measures, nor were they involved in developing plans for recruitment, design, or implementation of the study. No patients were asked to advise on interpretation or writing up of results.

There are no plans to disseminate the results of the research to study participants or the relevant patient community.

#### Results

##### Prevalence

Five hundred and fifty two cases of microcephaly were recorded, a hospital based prevalence of 4.4 (95% confidence interval 4.1 to 4.9) per 10000 births, including live births and stillbirths. Hospital



based prevalence varied by country and by region (table 2). Variation between hospitals was significant ( $\chi^2=950.5$ ,  $df=106$ ;  $P<0.001$ ) and is shown in fig 2 for 56 hospitals with at least 5000 births. Seventy one per cent (22408/31560) of controls and 49% (268/552) of cases were resident within the municipality of the hospital. The population based prevalence estimate was 3.0 (2.7 to 3.4) per 10000 births. Population based prevalence also varied by country and region (table 2). The highest prevalence was recorded in Brazil, specifically northeast and southeast Brazil, but the region of metropolitan Chile also recorded a high prevalence (table 2). One of the hospitals in Rio de Janeiro (southeast Brazil, hospital code A05; supplementary table) is a highly selective tertiary centre in the Rio municipality with an overall malformation rate of 22%. The population based prevalence for Rio (see footnote to table 2) remains

high, as Rio is a single municipality and there is further selection within municipality that cannot be corrected for; this explains some of the high prevalence in southeast Brazil. We found no significant secular trend in prevalence (prevalence rate ratio: linear term 0.99, 95% confidence interval 0.95 to 1.02; quadratic term 0.99, 0.98 to 1.00).

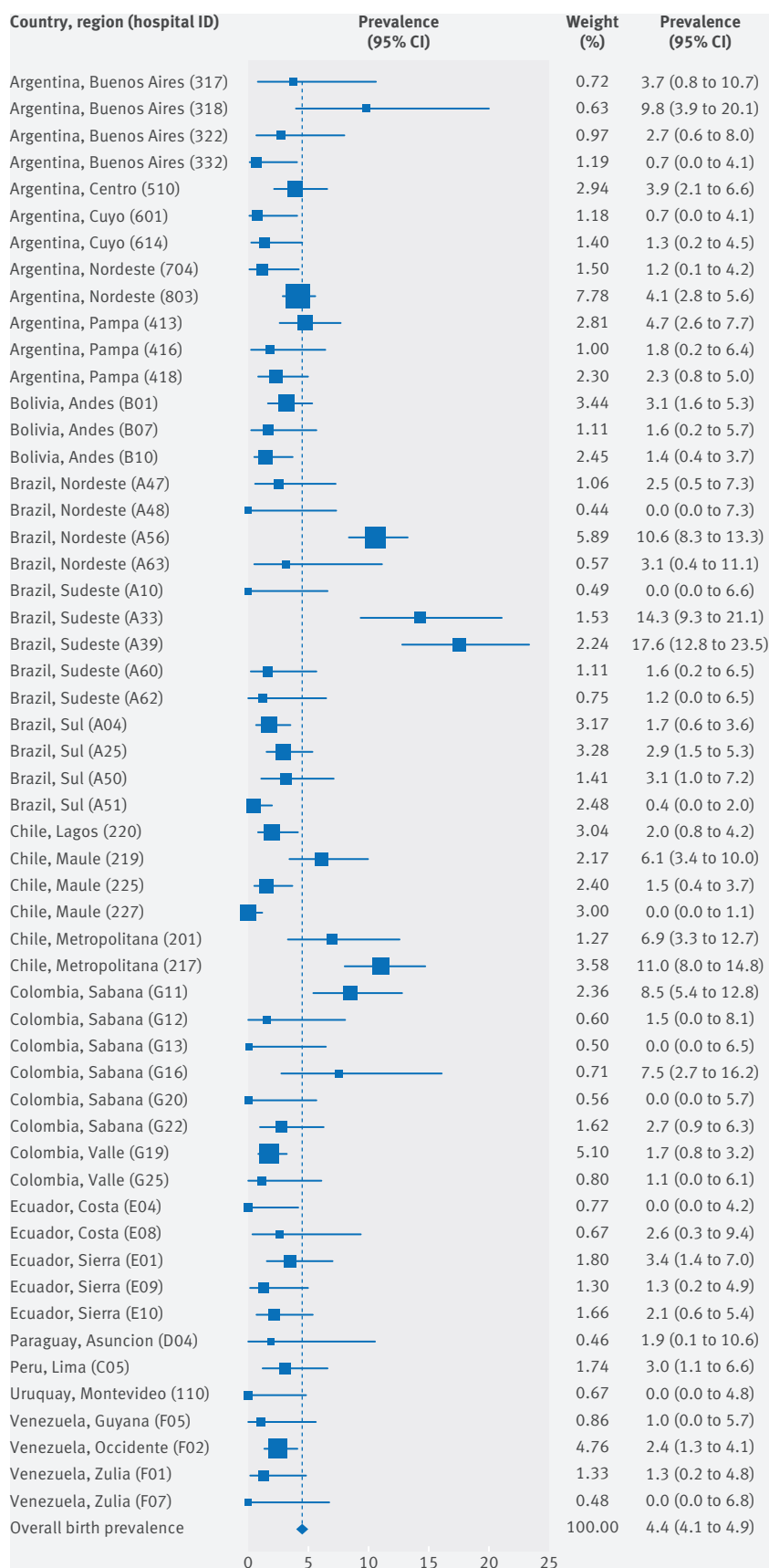
The live birth hospital based prevalence of microcephaly (as a proportion of all live births) was 4.5 (3.9 to 4.9) per 10000; 1.2% of all births were stillbirths, and as a proportion of all stillbirths the microcephaly prevalence was higher at 17.5 (16.9 to 18.1) per 10000. Overall, “perinatal mortality” was 29% among all microcephaly cases and 7% among neural only cases (table 3). Microcephaly or an associated congenital anomaly was prenatally diagnosed in 32% of all cases or 22% of neural only cases (table 3).

**Table 2 | Prevalence of microcephaly by country and region, 2005-14**

Country and region (No of hospitals)	Hospital based			Population based		
	Cases	Total births	Prevalence, per 10 000 (95% CI)	Cases	Births	Prevalence, per 10 000 (95% CI)
Argentina (35):	95	316 771	3.0 (2.4 to 3.7)	39	237 702	1.6 (1.2 to 2.2)
Buenos Aires (9)	17	53 032	3.2 (1.8 to 5.1)	4	43 420	0.92 (0.25 to 2.4)
Centro (2)	13	33 768	3.8 (2.1 to 6.6)	6	27 793	2.2 (0.79 to 4.7)
Cuyo (5)	3	31 566	1.0 (0.20 to 2.8)	2	26 609	0.75 (0.09 to 2.7)
Nordeste (1)	2	17 096	1.2 (0.14 to 4.2)	2	17 096	1.2 (0.14 to 4.2)
Noroeste (4)	36	92 936	3.9 (2.7 to 5.4)	6	45 718	1.3 (0.48 to 2.9)
Pampa (9)	23	80 233	2.9 (1.8 to 4.3)	18	68 926	2.6 (1.6 to 4.1)
Patagonia (5)	1	8140	1.2 (0.03 to 6.8)	1	8140	1.2 (0.03 to 6.8)
Bolivia (5):	18	80 322	2.2 (1.3 to 3.5)	11	58 080	1.9 (0.95 to 3.4)
Andes (4)	14	52 372	2.7 (1.5 to 4.5)	9	40 520	2.2 (1.0 to 4.2)
Subandina (1)	4	27 950	1.4 (0.4 to 3.7)	2	17 560	1.1 (0.14 to 4.1)
Brazil (22):	251	303 922	8.3 (7.3 to 9.4)	114	209 834	5.4 (4.5 to 6.5)
Centro (1)	1	3997	2.5 (0.06 to 13.9)	1	3854	2.6 (0.07 to 14.5)
Nordeste (4)	76	90 891	8.4 (6.6 to 10.5)	34	49 501	6.9 (4.8 to 9.6)
Sudeste (11)	144	88 052	16.4* (13.8 to 19.3)	70	68 178	10.3 (8 to 12.9)
Sul (6)	30	120 982	2.5 (1.7 to 3.5)	9	88 301	1.0 (0.47 to 1.9)
Chile (11):	89	192 401	4.6 (3.7 to 5.7)	43	102 631	4.2 (3.0 to 5.6)
Lagos (1)	7	34 665	2.0 (0.81 to 4.2)	5	23 229	2.2 (0.70 to 5.0)
Maule (4)	19	91 244	2.1 (1.3 to 3.3)	12	49 915	2.4 (1.2 to 4.2)
Metropolitana (6)	63	66 492	9.5 (7.3 to 12.1)	26	29 487	8.8 (5.8 to 12.9)
Colombia (15):	53	152 485	3.5 (2.6 to 4.6)	41	122 437	3.3 (2.4 to 4.5)
Caribe (1)	2	1411	14.2 (1.7 to 51.2)	1	1411	7.1 (0.18 to 39.5)
Sabana (12)	40	83 693	4.8 (3.4 to 6.5)	34	69 115	4.9 (3.4 to 6.9)
Valle (2)	11	67 381	1.6 (0.8 to 2.9)	6	51 911	1.2 (0.4 to 2.5)
Ecuador (11):	23	81 835	2.8 (1.8 to 4.2)	10	63 085	1.6 (0.76 to 2.9)
Costa (4)	2	19 484	1.0 (0.1 to 3.7)	0	16 666	0 (0 to 2.2)
Sierra (7)	21	62 351	3.4 (2.1 to 5.2)	10	46 419	2.2 (1.0 to 3.9)
Paraguay (1):	1	5237	1.9 (0.05 to 10.6)	0	1885	0 (0 to 19.6)
Asuncion (1)	1	5237	1.9 (0.05 to 10.6)	0	1885	0 (0 to 19.6)
Peru (1):	6	19 881	3.0 (1.1 to 6.6)	6	19 881	3.0 (1.1 to 6.6)
Lima (1)	6	19 881	3.0 (1.1 to 6.6)	6	19 881	3.0 (1.1 to 6.6)
Uruguay (2):	0	9553	0 (0 to 3.9)	0	9553	0 (0 to 3.9)
Maldonado (1)	0	1922	0	0	1922	0
Montevideo (1)	0	7631	0	0	7631	0
Venezuela (4):	16	84 778	1.9 (1.1 to 3.1)	4	55 986	0.71 (0.19 to 1.8)
Guyana (1)	1	9868	1.0 (0.03 to 5.7)	1	9868	1.0 (0.03 to 5.7)
Occidente (1)	13	54 317	2.4 (1.3 to 4.1)	2	25 525	0.78 (0.09 to 2.8)
Zulia (2)	2	20 593	1.0 (0.12 to 3.5)	1	20 593	0.49 (0.01 to 2.7)
All countries (107)	552	1 247 185	4.4* (4.1 to 4.9)	268	881 074†	3.0 (2.7 to 3.4)

\*Removal of high prevalence tertiary A05 hospital from Brazil gives estimated hospital based prevalence for Brazil SE of 9.6/10 000 (95% CI 7.5 to 11.9;  $n=75$ ; births=78 470) and all countries prevalence of 3.9/10 000 (3.6 to 4.3;  $n=483$ ; births=1 237 603). A05 alone has hospital based prevalence of 72/10 000 (56.1 to 91;  $n=69$ ; births=9582) and population based prevalence of 52.4/10 000 (36.3 to 76.1). Brazil SE population based prevalence without A05 is 2.7/10 000 (2.3 to 3.0).

†Estimated from 71% of controls residing within municipality of hospital ( $0.71 \times 1\,247\,187=881\,074$ ).



**Fig 2 | Forest plot of hospital based prevalence per hospital (country, region), 2005-14, showing only 55 hospitals with ≥5000 births. Hospital codes are given in supplementary table. Hospital A05 not shown as off scale: hospital based prevalence=72/10 000 (95% CI 56.1 to 91.0; n=69, births=9582)**

Extrapolated to the 6.8 million births in South America, we estimate that the number of microcephaly cases being born per year before Zika was between 2058 (simple estimate) and 2465 (stratified estimate, substituting South American average prevalence for the two countries with zero cases in table 2).

### Clinical and phenotypic spectrum

Table 1 shows the distribution of clinical classification (2005-14). Twenty three per cent or 1.0 per 10 000 births were genetic syndromes, 5% were embryopathies, 34% were polymalformed, 12% were microcephaly with associated neural conditions, and 26% were coded only as microcephaly; 230 cases (42% of all cases) were neural only cases (table 1). The sex ratio was significantly skewed towards females (tables 1 and 3).

Fifty four per cent of microcephaly cases were low birth weight (<2500 g) (table 3); 27% of microcephaly cases had a birth weight lighter than 2 SD below the mean (−2 SD), including 7% with a birth weight lighter than −3 SD (table 3). For the 2010-14 period cases, 45% of head circumferences were less than −3 SD and 24% were between −3 and −2 SD. On inspection of 41/58 cases with head sizes larger than −2 SD, eight (19%) had hydrocephaly, seven (17%) had spina bifida, two (5%) had lissencephaly, and seven (17%) had other cerebral defects; other conditions (2 (1%)) were mentioned in which brain size might not be reflected in head size, such as head deformity and loose/redundant skin over the head, and others were unexplained. Of the 126 cases with a head size smaller than −2 SD, 47 (37%) also had a birth weight smaller than −2 SD. The equivalent figure for −3 SD was 82 cases, of which 10 (12%) also had a birthweight smaller than −3 SD.

Parental consanguinity was present in 2% (n=11) of all cases and 0.7% of controls (table 4). Four of the consanguineous cases had been diagnosed as having genetic syndromes (category A), but consanguinity was also present in seven isolated (category D) cases with no diagnosed genetic syndrome, of which five were first cousins. The association between consanguinity and neural only microcephaly was strong (odds ratio 3.1, 95% confidence interval 1.4 to 7.0) (table 4).

STORCH embryopathies were diagnosed in 3.8% of total cases (table 1), with an odds ratio of 4.4 (2.3 to 8.4) (table 4). Of the 21 STORCH microcephaly cases, 15 were neural only cases (7% of all neural only cases) and six had other congenital anomalies (table 1).

The polymalformed cases included 13 with microcephaly and arthrogryposis, all of which also had other malformations of non-neural origin. There were no cases of microcephaly and arthrogryposis that did not also have non-neural origin anomalies.

### Sociodemographic risk factors

For all cases and neural only cases, we found no association with older maternal age or low maternal education (table 4).

**Table 3 | Description of all microcephaly cases and neural only cases\*. Values are numbers (percentages)**

Characteristic	All microcephaly cases (n=552)	Neural only cases (n=230)
Live born	525 (95.1)	229 (99)
Stillborn	25 (4.5)	1 (0.4)
All perinatal deaths	159 (28.8)	17 (7)
Male	227 (41.1)	92 (40)
Female	315 (57.1)	138 (60)
Intersex	8 (1.4)	0
Twin	14 (2.5)	8 (3)
Prenatally diagnosed	175 (31.7)	51 (22)
Preterm (<37 weeks)	195 (35.3)	51 (22)
Term not specified	32 (5.8)	17 (7)
Low birth weight (<2500 g)	296 (53.6)	94 (41)
Birth weight z score not calculated	46 (8.3)	29 (13)
Birth weight <-2 SD	134/506 (26.5)	40/201 (20)
Birth weight <-3 SD	33/506 (6.5)	7/201 (3)
<b>2010-14 only</b>		
All data:	235 (42.6)	95 (41)
HC z score not calculated†	51/235 (22)	8/95 (8)
HC<-3 SD	82/184 (45)	38/87 (44)
-2 SD≤HC<-3 SD	44/184 (24)	27/87 (31)
-1 SD≤HC<-2 SD	37/184 (20)	18/87 (21)
HC>-1 SD	21/184 (11)	4/87 (5)
Preterm:	55/184 (30)	15/87 (17)
HC<-3 SD	21/55 (38)	8/15 (53)
-2 SD≤HC<-3 SD	8/55 (15)	2/15 (13)
-1 SD≤HC<-2 SD	14/55 (25)	4/15 (27)
HC>-1 SD	12/55 (22)	1/15 (7)
Term:	129/184 (70)	72/87 (83)
HC<-3 SD	61/129 (47)	30/72 (42)
-2 SD≤HC<-3 SD	36/129 (28)	25/72 (35)
-1 SD≤HC<-2 SD	23/129 (18)	14/72 (20)
HC>-1 SD	9/129 (7)	3/72 (4)

\*Two cases were unspecified as to whether live born or stillborn and for sex.

†z score not calculated as sex not specified (2), intersex (4), gestational age not specified (1), gestational age outside Intergrowth range (24-42 weeks+6 days) (8), or head circumference not specified (43).

## Discussion

Before the Zika epidemic, the average hospital based prevalence of microcephaly recorded by ECLAMC in South America for the 2005-14 period was 4.4 per 10 000 births, and the estimated population based prevalence was 3.0 per 10 000. Prevalence varied considerably by country, region, and hospital. As selective referral of prenatally diagnosed cases of congenital anomaly or high risk pregnancies to tertiary hospitals may occur, the prevalence in single hospitals may be inflated compared with the population as a whole. We estimate the average population based prevalence to be 31% less than the hospital based prevalence and recommend that Zika related surveillance takes hospital selection into account and provides population based prevalence. However, we found that hospital selection explained little of the geographical variation in prevalence of diagnosed cases—substantial geographical variation remained in population based prevalence. On the basis of this prevalence of 3.0 per 10 000, we estimate that 2000 to 2500 babies were being diagnosed as having microcephaly each year before the Zika epidemic, among the nearly 7 million annual births in South America.

## Comparison with other studies

Previous estimates of the prevalence of microcephaly in South America and elsewhere have been within a similar range. ECLAMC data showed an average hospital based prevalence of 3.7/10 000 for 1982-2008, ranging from 2.1 in Venezuela to 5.1 in Brazil and 5.5 in Chile.<sup>37</sup> Other studies in South America that overlap to a small extent with ECLAMC data reported a population based prevalence of 3.1/10 000 in the region of Risaralda, Colombia, in 2010-13 and a hospital based prevalence of 2.4/10 000 in Argentina in 2010-14.<sup>38 39</sup> In Europe in 2003-12, a population based prevalence of 1.5/10 000, excluding 23% with genetic syndromes, was recorded,<sup>31</sup> and a hospital based prevalence of 2.3/10 000 was reported in India from 1963 to 1995.<sup>40</sup> The pooled prevalence of microcephaly among the United States population based surveillance systems was higher at 8.7 per 10 000 from 2009 to 2013.<sup>32</sup> Both European and US surveillance systems documented considerable geographical variation, which was due at least in part to diagnostic variability.

## Implications for diagnosis of microcephaly

The ECLAMC surveillance data, like data from other surveillance systems,<sup>31 32</sup> records diagnoses made by clinicians for clinical (not research) purposes, which are only to a limited extent subject to standardisation by belonging to a surveillance network. In our data, about half of cases had a head size larger than -3 SD (according to the Intergrowth curve), including some larger than -2 SD in whom other signs indicated an abnormally small brain. This is despite the ECLAMC recommendation to use a threshold of -3 SD. Some of this discrepancy may relate to differences between the growth charts used (Intergrowth charts were not widely used in this period) and whether they explicitly show the -3 SD threshold measurements. Most readily available charts, including Intergrowth, showed only the third centile during this period,<sup>35 41-43</sup> the third centile being equivalent to the lowest 3% of measurements of the reference population compared with 2.3% for -2 SD and 0.13% for -3 SD in a normal distribution (<https://www.zscorecalculator.com/>). It also follows that use of other charts might have meant that clinicians missed cases that would be microcephalic by the Intergrowth chart. A US analysis of surveillance data found that 42% of microcephaly cases had head circumference larger than the 10th centile (the lowest 10%, equivalent to -1.3 SD) by the Intergrowth curve, an even greater proportion of large head circumferences than we found in South America.<sup>32</sup>

Our data, the literature,<sup>4 31 32 45 46</sup> and discussions with paediatricians suggest that clinical judgment goes beyond a simple metric microcephaly. Paediatricians do (and should) use other signs as well as head circumference at birth to determine which baby has a small brain and whether this is likely to be associated with brain pathology and poor neurodevelopmental outcome. Head circumference is only a screening

**Table 4 | Potential risk factors for microcephaly: number (percentage) exposed for cases and controls, and odds ratio adjusted for maternal age and education (95% CI)**

Exposure	All microcephaly cases			Neural only microcephaly		
	Cases (n=546)*	Controls (n=2176)	Adjusted odds ratio (95% CI)	Cases (n=230)	Controls (n=917)	Adjusted odds ratio (95% CI)
Maternal age:						
<20 years	104 (19.0)	444 (20.4)	0.81 (0.62 to 1.0)	49 (21)	209 (22.8)	0.77 (0.53 to 1.1)
20-34 years	360 (65.9)	1263 (58.0)	1.0	156 (68)	547 (59.7)	1.0
≥35 years	82 (15.0)	241 (11.1)	1.2 (0.88 to 1.6)	25 (11)	97 (10.6)	0.84 (0.51 to 1.4)
Not specified	0	228 (10.5)	–	0	64 (7.0)	–
Maternal education:						
<6 years	111 (20.3)	345 (15.9)	1.2 (0.92 to 1.6)	37 (16)	147 (16.0)	1.0 (0.64 to 1.5)
6-12 years	358 (65.6)	1364 (62.7)	1.0	161 (70)	597 (65.1)	1.0
≥13 years	55 (10.1)	186 (8.5)	1.1 (0.77 to 1.5)	25 (11)	91 (9.9)	1.0 (0.59 to 1.6)
Not specified	22 (4.0)	281 (12.9)	–	7 (3)	82 (8.9)	–
Consanguinity:						
Yes	11 (2.0)	15 (0.7)	3.1 (1.4 to 7.0)	7 (3)	10 (1.1)	3.1 (1.1 to 8.9)
No	535 (98.0)	1865 (85.7)	1.0	216 (94)	818 (89.2)	1.0
Not specified	0	296 (13.6)	–	7 (3)	89 (9.7)	–
STORCH:						
Yes	21 (3.8)	21 (1.0)	4.4 (2.3 to 8.3)	15 (7)	15 (1.6)	6.3 (2.6 to 15.1)
No	525 (96.2)	2155 (99.0)	1.0	215 (93)	902 (98.4)	1.0
Not specified	0	0	–	0	0	–
Diabetes:						
Yes	13 (2.4)	52 (2.4)	1.0 (0.53 to 1.9)	5 (2)	5 (0.5)	1.0 (0.38 to 2.9)
No	533 (97.6)	2124 (97.6)	1.0	225 (98)	912 (99.5)	1.0
Not specified	0	0	–	0	0	–
Influenza:						
Yes	26 (4.8)	57 (2.6)	1.9 (1.1 to 3.3)	12 (5)	12 (1.3)	2.7 (1.8 to 6.4)
No	520 (95.2)	2119 (97.4)	1.0	218 (95)	905 (98.7)	1.0
Not specified	0	0	–	0	0	–

STORCH=syphilis, toxoplasmosis, other including HIV, rubella, cytomegalovirus, and herpes simplex.

\*Six cases were removed as matched controls could not be identified owing to cases being intersex (4) or having unspecified sex (2).

tool, but it may have assumed a status as a sole metric diagnostic arbiter in textbook definitions. Other signs that improve the predictive value may include slowing growth rate of the head in utero and postnatally, presence of associated major or minor malformations (59% of our larger head size cases had neural anomalies that suggested small brain size) or abnormal facies suggesting teratogenic factors, exposure to known risk factors (of which Zika may become one), proportionality of head size to body size,<sup>47</sup> parental head size, brain imaging results,<sup>14 15 48-50 52</sup> and any abnormalities of head shape that invalidate judgments made about occipitofrontal head circumference such as head moulding or water loss.<sup>53-55</sup>

Some under-ascertainment of microcephaly in ECLAMC and other surveillance systems is also possible. When the underlying or an associated pathology is known and diagnosed—for example, a specific syndrome or specific brain malformation—microcephaly may not always be specified as a supplementary diagnosis.

The WHO has recommended in response to the Zika epidemic that microcephaly should be defined as head circumference below  $-2$  SD,<sup>51</sup> rather than  $-3$  SD as used by congenital anomaly surveillance systems,<sup>31 32</sup> particularly as not all Zika affected babies have heads smaller than  $-3$  SD.<sup>14 21 50</sup> As a definition of  $-2$  SD on its own has the disadvantage of low specificity,<sup>50</sup> and unnecessary anxiety, giving many babies with normal brains a microcephaly diagnosis,<sup>22 46</sup> and leading to

a huge potential increase in recorded microcephaly prevalence compared with previous levels, both the Brazilian microcephaly monitoring and the WHO surveillance manual have recommended that a diagnosis of microcephaly should be confirmed at this level for congenital Zika surveillance only if other evidence of structural abnormalities exists.<sup>33 51</sup> The implementation of this definition may lead to a certain amount of standardisation in future but leave several areas of potential variability as discussed above. What is clear is that purely “metric microcephaly” is not a viable definition of a clinical entity, even when restricted to  $-3$  SD, and also that an alert such as Zika has the potential to lead to large artefactual increases in the prevalence of diagnosed microcephaly. Microcephaly surveillance should move beyond measurement based definition to a pathology based definition.

#### Potential explanations and implications

In the period 2005-14 before Zika, the highest microcephaly prevalences in our data were already recorded in Brazil, especially southeast and northeast Brazil, where later (from 2015 on) the Zika epidemic and microcephaly surge began, but a high prevalence was also observed in metropolitan Chile. We have no explanation for this or more generally for the geographic variation observed—a vulnerability to or higher background frequency of teratogenic exposures, or diagnostic factors as discussed above, or both could be hypothesised.



The range of abnormalities associated with congenital Zika syndrome are still being defined.<sup>19-28</sup> Although initially the recognition of this syndrome drove a widening of the microcephaly definition, it is now recognised that microcephaly is only one of a range of neural abnormalities. We found 2.4% (13/552) of cases of microcephaly with arthrogryposis, one of the findings relating to Zika,<sup>17 18</sup> but these were all polymalformed cases that also had other non-neural malformations. At least 34 genetic syndromes present with congenital microcephaly plus arthrogryposis.<sup>56</sup> Intrauterine growth retardation has been a common Zika-related finding,<sup>13 15</sup> including in animal models.<sup>57 58</sup> We found proportionate microcephaly to account for 37% of babies with head size less than -2 SD, but for those with head size more than -3 SD it was uncommon for birth weight to be proportionately reduced. Development of the congenital Zika syndrome definition will need to consider carefully whether to distinguish proportionate growth retardation with normally formed brain from microcephaly with evidence of abnormal brain structure, including consideration of prognosis in terms of neurodevelopmental outcome, which may come from the cohort studies now in progress.

Twenty three per cent of our cases had a diagnosed genetic syndrome, a prevalence of 1 per 10 000, similar to the 23% of genetic cases reported by EUROCAT and 23% reported in the US,<sup>31 32</sup> and it is important that differential diagnosis is maintained during epidemic periods. Autosomal recessive types of microcephaly, well recognised in the microcephaly literature,<sup>59</sup> for which consanguineous parents would be at higher risk, are more difficult to identify at birth as suggested by our very high odds ratio for parental consanguinity among neural only cases, of which few had been diagnosed as having syndromes. As new genetic diagnostic methods become more widely available, the proportion of genetic syndromes is likely to rise in future.

Although Zika is of great relative importance during an epidemic, owing to high numbers of exposed non-immune women, the strong and known associations we found with other infections are of continuing importance. Current estimates of the rate of affected babies after Zika infection in early pregnancy are in the order of 1-13%,<sup>2 60 61</sup> but infections such as rubella are more highly teratogenic (with up to 100% birth defects when exposure occurred before the 11th week or 35% of those infected between the 13th and 16th weeks) but not as specific to neural anomalies.<sup>62</sup> We found in these pre-Zika data that 4% of microcephaly cases (or 7% of neural only cases) were associated with STORCH infections, in some cases more than one infection. A Colombian surveillance study reported that 21% of microcephaly cases tested had evidence of STORCH infections.<sup>63</sup> The STORCH infections in 1.5% of our non-malformed controls, some of whom may have had later neurodevelopmental problems, suggest a significant population level of maternal infection during pregnancy, even though this is likely

to be an underestimate owing to incomplete screening. We also found a significant association with reported influenza, which has not previously been found in the literature,<sup>64</sup> that may also indicate fever relating to other infectious diseases being interpreted as influenza. Zika may be a useful “wake-up call” to the effect of infections during pregnancy. Each STORCH infection is estimated to be associated with between 100 000 and more than a million cases of congenital infection globally each year.<sup>5 65 67</sup>

### Strengths and limitations of study

A strength of the ECLAMC surveillance system is that it uses a common protocol in all countries. However, hospitals receive no funding for their participation and are not able to follow up cases beyond discharge from the maternity unit. This means that some brain imaging and genetic syndrome diagnoses for microcephaly cases may be missing, and some of the polymalformed or isolated cases classified here may have later been given syndrome diagnoses. Diagnostic data are collected within the healthcare possibilities of each hospital, in a “real world” rather than research context. This is common to surveillance systems worldwide,<sup>31 32</sup> which use standardised reporting systems for healthcare data that may arise from non-standardised clinical practice. We found that data were missing for some variables, which indicates the need for improved reporting in ECLAMC. Mothers may not accurately or completely recall information about their exposures when interviewed. Our description of head circumference of microcephaly cases was limited to 2010-14 and may not be representative of the earlier period (2005-09).

Whereas hospital based surveillance has the advantage of ensuring high data quality coming from selected participating hospitals, a strength of the ECLAMC system, selective referral to tertiary or specialist hospitals is problematic for the estimation of population based prevalence. As there are often multiple hospitals within municipalities, and it was not possible to correct for within municipality selection, we could not give fully population based prevalence estimates, but we believe our population based estimates to be an improvement over the hospital based estimates for public health surveillance. Our extrapolation to the entire population of South America is approximate only. All South American countries, except Guyana, Suriname, and French Guiana, were represented in our data.

### Conclusions

We have established a baseline for the prevalence of diagnosed microcephaly before the Zika epidemic with which the prevalence during and after the Zika epidemic can be compared, with strong caveats due to the problems of definition. We have also shown that, even in non-Zika times, the potential effects of infectious and non-infectious causes of abnormal brain development should not be ignored. What may set the Zika epidemic apart from other endemic infections are

## the high exposure rate in the population as it sweeps through non-immune populations and its specificity to neural anomalies.

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**Data sharing:** The data used in this study belong to the individual hospitals. Data are available on request from the corresponding author, who will ask the individual hospitals' permission to use the data.

**Transparency:** The lead author affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned (and, if relevant, registered) have been explained.

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## Supplementary table