Effect of prolonged and exclusive breast feeding on risk of allergy and asthma: cluster randomised trial

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ABSTRACT

Objective To assess whether exclusive and prolonged breast feeding reduces the risk of childhood asthma and allergy by age 6.5 years.

Design Cluster randomised trial.

Setting 31 Belarussian maternity hospitals and their affiliated polyclinics.

Participants A total of 17 046 mother-infant pairs were enrolled, of whom 13 889 (81.5%) were followed up at age 6.5 years.

Intervention Breastfeeding promotion intervention modelled on the WHO/UNICEF baby friendly hospital initiative.

Main outcome measures International study of asthma and allergies in childhood (ISAAC) questionnaire and skin prick tests of five inhalant antigens.

Results The experimental intervention led to a large increase in exclusive breast feeding at 3 months (44.3% v 6.4%; $P<0.001$) and a significantly higher prevalence of any breast feeding at all ages up to and including 12 months. The experimental group had no reduction in risks of allergic symptoms and diagnoses or positive skin prick tests. In fact, after exclusion of six sites (three experimental and three control) with suspiciously high rates of positive skin prick tests, risks were significantly increased in the experimental group for four of the five antigens.

Conclusions These results do not support a protective effect of prolonged and exclusive breast feeding on asthma or allergy.

Trial registration Current Controlled Trials ISRCTN37687716.

INTRODUCTION

Whether breast feeding protects against the development of allergy and asthma has been frequently studied and hotly debated for more than 70 years.¹⁻⁸ Research findings indicating a beneficial effect have been most consistent for atopic eczema during infancy, but the evidence on asthma and other atopic outcomes (including hay fever, food allergies, and positive skin tests) has been far more mixed. Some studies have

reported greater degrees of protection with more exclusive and prolonged breast feeding,⁹⁻¹² and several have noted a larger protective effect in children prone to atopy.^{4,5} Other studies, however, have reported no reduction in risk or even an increase in risk with breast feeding.¹³⁻²⁰

A variety of methodological problems are likely to have affected the evidence base on infant feeding and allergic disease. Virtually all of the evidence is based on observational studies. Case-control studies are prone to recall bias, which may be systematic if the people who ascertain the previous feeding history are aware of the case versus control status of the study participants. Cohort studies, on the other hand, are prone to biased assessment of outcomes when observers are not blinded to the previous infant feeding history, a particularly important problem for outcomes such as atopic eczema, hay fever, and asthma, for which no definitive laboratory measures are available to provide a "gold standard" for diagnosis. Misclassification of infant feeding history, particularly with respect to differences in degree or duration of breast feeding, is always a problem in studying the health effects of breast feeding but is even more problematic in studies of atopic disease. This is because hypersensitivity reactions in particular and atopic disease in general may not show the dose-response effects seen with other outcomes related to infant feeding and may be affected not only by the quantity of exposure but also by its frequency and timing. One recent study has even reported evidence of reverse causality, whereby early signs of atopic disease led to earlier discontinuation of exclusive breast feeding.²¹ Finally, publication bias may well have affected the evidence base, with selective submission and publication of studies with positive findings (that is, showing that breast feeding is protective).

One potential solution to these methodological problems is a randomised controlled trial. Although randomising healthy mothers and infants to breast feeding versus formula feeding is infeasible and probably unethical, randomising them to an intervention that

Table 1 | Baseline characteristics of children followed up at age 6.5 years in experimental and control groups. Values are numbers (percentages) unless stated otherwise

Variable	Experimental group (n=7108)	Control group (n=6781)
Maternal age (years):		
<20	1017 (14.3)	897 (13.2)
20-34	5783 (81.4)	5598 (82.6)
≥35	308 (4.3)	286 (4.2)
Maternal education:		
Incomplete secondary	310 (4.4)	202 (3.0)
Complete secondary	2441 (34.3)	2011 (29.7)
Advanced secondary or partial university	3395 (47.8)	3693 (54.5)
Complete university	962 (13.5)	875 (12.9)
Older children living in household:		
0	4176 (58.8)	3695 (54.5)
1	2368 (33.3)	2450 (36.1)
≥2	564 (7.9)	636 (9.4)
Positive family history of atopy	359 (5.1)	226 (3.3)
Maternal smoking during pregnancy	184 (2.6)	108 (1.6)
Male child	3653 (51.4)	3528 (52.0)
Mean (SD) birth weight (g)	3440 (418)	3441 (423)

promotes breast feeding is both feasible and ethical. One potential strategy would be to intervene prenatally to increase the proportion of mothers who start breast feeding, but such a strategy poses logistical problems of reaching and influencing women during or even before pregnancy. An alternative and more practical strategy is to promote exclusivity and duration of breast feeding among those mothers who have already decided to start breast feeding. This is the strategy we used in designing and implementing the promotion of breastfeeding intervention trial (PROBIT), a cluster randomised trial in the Republic of Belarus.²² In this paper, we describe the methods and results of measures of allergic symptoms and diagnoses and skin prick tests after 6.5 years of follow-up among Belarusian children enrolled in PROBIT.

Table 2 | Audit results

	κ (95% CI)
ISAAC questions (n=190)	
Ever had wheezing	0.77 (0.61 to 0.92)
Wheezing in past 12 months	0.74 (0.59 to 0.90)
Ever had asthma	0.66 (0.52 to 0.81)
Ever had hay fever symptoms	0.88 (0.73 to 1.00)
Hay fever symptoms in past 12 months	0.92 (0.77 to 1.00)
Recurrent itchy rash	0.53 (0.39 to 0.67)
Ever had eczema	0.50 (0.36 to 0.63)
Skin prick test antigens (n=54)	
House dust mite	0.78 (0.49 to 1.00)
Cat	0.59 (0.30 to 0.87)
Birch pollen	0.49 (0.21 to 0.78)
Mixed northern grasses	0.94 (0.65 to 1.00)
<i>Alternaria</i>	0.55 (0.26 to 0.84)

ISAAC=international study of asthma and allergy in childhood.

METHODS

The detailed methods of PROBIT and the results during the first year of follow-up have been previously reported.²² The units (clusters) of randomisation were maternity hospitals and one affiliated polyclinic (out-patient clinic where children are followed for well child and illness care) for each hospital, with double randomisation based on both a random numbers table and a coin flip.²² The experimental intervention was based on the baby friendly hospital initiative, which was developed by the World Health Organization (WHO) and the United Nations Children's Fund (UNICEF) to promote and support breast feeding, particularly among mothers who have chosen to start breast feeding.²³ The control maternity hospitals and polyclinics continued the practices and policies in effect at the time of randomisation. The trial results are based on a total of 17 046 healthy breastfed infants from 31 maternity hospitals/polyclinics; all were born at term in 1996-7, weighed at least 2500 g, and were enrolled during their postpartum stay.²² To our knowledge, PROBIT is the largest randomised trial ever done in the area of human lactation. It conforms to the CONSORT recommendations for the design, analysis, and reporting of cluster randomised trials.²⁴

As previously reported,²² the two randomised groups were similar in baseline sociodemographic and clinical variables, including maternal age, education, number of other children at home, proportion who had breast fed a previous child for at least three months, family history of atopy, caesarean delivery, maternal smoking during pregnancy, birth weight, gestational age, and five minute Apgar score. The experimental intervention led to a substantial difference in the duration of any breast feeding that was maintained throughout the first year of follow-up: 72.7% versus 60.0% were still breast feeding at 3 months, 49.8% versus 36.1% at 6 months, 36.1% versus 24.4% at 9 months, and 19.7% versus 11.4% at 12 months in the experimental and control groups. In addition, the prevalence of exclusive breast feeding was sevenfold higher in the experimental group at 3 months (43.3% *v* 6.4%; $P<0.001$), although low in both groups at 6 months (7.9% *v* 0.6%; $P=0.01$).²²

One paediatrician in each of 24 of the 31 polyclinics did follow-up interviews and examinations at age 6.5 years from December 2002 to April 2005; in the remaining seven high volume clinics, follow-up visits were shared by two paediatricians. Allergy symptoms and diagnoses were ascertained with the international study of asthma and allergy in childhood (ISAAC) questionnaire, which had already been translated and validated by the ISAAC investigators.²⁵ In addition, the paediatricians did skin prick tests to five antigens (Allergy Canada): house dust mite, cat, birch pollen, mixed northern grasses, and *Alternaria*. Saline was included as a negative control and histamine (1 mg/ml) as a positive control. The criteria for a positive result were a mean wheal ≥ 3 mm or flare ≥ 10 mm, calculated as the mean of the longest diameter and orthogonal diameter after subtracting the mean of these

Table 3 | Positive skin prick test wheal remeasurements (n=119). Values are intraclass correlation coefficients (95% confidence intervals)

Antigen	Longest diameter	Orthogonal diameter
House dust mite	0.96 (0.95 to 0.98)	0.96 (0.94 to 0.97)
Cat	0.93 (0.90 to 0.95)	0.86 (0.81 to 0.94)
Birch pollen	0.94 (0.91 to 0.95)	0.97 (0.96 to 0.98)
Mixed northern grasses	0.95 (0.92 to 0.96)	0.98 (0.98 to 0.99)
<i>Alternaria</i>	0.94 (0.91 to 0.96)	0.92 (0.89 to 0.95)

diameters for the saline control. A negative test result required a positive histamine test: mean diameter with histamine minus mean diameter with saline ≥ 3 mm for wheal or ≥ 5 mm for flare. A line in marker pen was drawn around each flare and wheal and transferred to paper tape, which was then stuck on to the data forms sent to the PROBIT data centre in Minsk.

Because blinding of paediatricians to the experimental versus control assignment was not feasible, we randomly selected five children per paediatrician (n=38) for audit, for a total of 190 audited children. To ensure that all children followed up in PROBIT II were eligible to be audited, we did the audits after completion of follow-up visits, a mean of 17.7 (range 5.3–32.6) months after the initial polyclinic visit. One of the Minsk based paediatrician collaborators (LM, ZS, ID, or GS) did the audits; they were blind to the responses to the ISAAC questionnaire and results of the skin prick tests obtained at the initial clinic visit but not to the experimental versus control treatment allocation. We used the κ coefficient (and its 95% confidence interval) to assess inter-observer agreement. As a further data quality control check, we remeasured the wheal and flare results recorded on transfer to paper tape on a randomly selected sample of 1% (n=119) of the children with positive test results. We used the intraclass correlation coefficient (and its 95% confidence interval) to quantitate the agreement between the two measurements.

We based all statistical analyses on intention to treat. We used the GLIMMIX procedure in SAS (version 8.2) to analyse differences in outcome between the experimental and control groups. This is a multilevel statistical modelling procedure that uses penalised quasi-likelihood for generalised linear mixed models and accounts for the clustered randomisation.²⁶ The

procedure estimates odds ratios and 95% confidence intervals for dichotomous outcomes, thus permitting inference at the level of the individual child, rather than at the level of the cluster (maternity hospital and polyclinic). The cluster adjusted differences presented in the text and tables are very similar to those obtained from a multivariate model that adjusts for stratum level variables (geographic region (west *v* east) and urban *v* rural location), as well as the following individual level covariates: age at follow-up, sex, birth weight, family history of atopy, pet ownership, contact with farm animals, and previous exposure to probiotics (multivariate results available on request). We did not impute missing data. Finally, because several previous studies have suggested that breast feeding is more protective in infants with a positive first degree (parent or sibling) family history of atopy,^{4,5} we also analysed mixed models that included terms for family history and a multiplicative interaction term for family history and treatment.

RESULTS

A total of 13 889 children were seen in follow-up for PROBIT II, representing 81.5% of the 17 046 originally randomised. Of the 3157 children randomised but not followed up, 88 had died, 2938 were lost to follow-up, and 131 were unable or unwilling to come for their PROBIT II visit. Follow-up rates were similar in the experimental (80.2%) and control (82.9%) polyclinics but varied considerably by polyclinic—from 56.1% at one of the Minsk polyclinics to 94.6% at Klimovich, a small rural based polyclinic. The mean age at follow-up was 6.6 (SD 0.3) years. As shown in table 1, the children followed up in the experimental and control groups were similar in baseline characteristics, with small differences paralleling those seen (and previously reported²²) at randomisation.

Table 2 shows the audit results. The data shown are the κ coefficients (and their 95% confidence intervals) between the results at the initial clinic visit and the results at the audit visit. These κ values are high with respect to wheezing and hay fever symptoms and moderate for reported diagnosis of asthma or symptoms and diagnosis of atopic eczema. Concordance was high for the skin prick test results, but only 54 (28%) of the 190 audited children (or their parents) agreed to the repeat skin tests. The results of the random skin prick test wheal remeasurements indicated exceptionally high intraclass correlation coefficients for both the longest diameter and orthogonal diameter for all five antigens tested (table 3).

Table 4 shows the trial results for the ISAAC questionnaire, the number and proportion of those with positive responses in the experimental and control groups, the intraclass correlation coefficients reflecting the degree of within polyclinic clustering, and the cluster adjusted odds ratios and their 95% confidence intervals. The results showed a very low degree of clustering—that is, very little tendency for mothers of children attending the same polyclinic to respond more similarly to each other than those of children attending

Table 4 | ISAAC results. Values are numbers (percentages) positive unless stated otherwise

Question	Experimental group (n=7101)	Control group (n=6763)	ICC	Cluster adjusted odds ratio* (95% CI)
Ever had wheezing	778 (11.0)	651 (9.6)	0.03	1.1 (0.6 to 1.8)
Wheezing in past 12 months	238 (3.4)	188 (2.8)	0.01	1.0 (0.7 to 1.6)
Ever had asthma	97 (1.4)	68 (1.0)	0.00	1.2 (0.7 to 1.9)
Ever had hay fever symptoms	384 (5.4)	257 (3.8)	0.02	1.1 (0.6 to 1.9)
Hay fever symptoms in past 12 months	262 (3.7)	192 (2.8)	0.01	1.0 (0.6 to 1.8)
Recurrent itchy rash	350 (4.9)	241 (3.6)	0.02	1.3 (0.7 to 2.2)
Ever had eczema	69 (1.0)	72 (1.1)	0.00	1.0 (0.5 to 1.8)

ICC=intraclass correlation coefficient.

*For experimental group versus control group.

Table 5 | Skin prick test results. Values are numbers (percentages) positive unless stated otherwise

Antigen	Experimental group (n=5551)	Control group (n=5595)	ICC	Cluster adjusted odds ratio* (95% CI)
House dust mite	805 (14.5)	603 (10.8)	0.14	1.1 (0.5 to 2.4)
Cat	648 (11.7)	491 (8.8)	0.20	1.2 (0.5 to 2.8)
Birch pollen	526 (9.5)	393 (7.0)	0.18	1.2 (0.5 to 2.9)
Mixed northern grasses	712 (12.8)	491 (8.8)	0.17	1.0 (0.5 to 2.3)
<i>Alternaria</i>	480 (8.6)	340 (6.1)	0.18	1.5 (0.5 to 4.4)
≥1 positive	1496 (27.0)	1013 (18.1)	0.19	1.2 (0.5 to 2.6)

ICC=intraclass correlation coefficient.

*For experimental group versus control group.

different polyclinics. Most of the symptoms and diagnoses were slightly more prevalent in the experimental group than in the control group, but the cluster adjusted odds ratios in the experimental versus control groups were close to unity for all of the symptoms and diagnoses.

Table 5 shows the skin prick test results. Of the 13 889 PROBIT children seen at follow-up, 11 772 (85%) agreed to the skin prick tests, of whom 11 146 (95%) had valid results, as judged by the criterion of a positive test to histamine (see methods). Positive skin test results were slightly but consistently more frequent in the experimental group for all five test antigens, although none of the differences was statistically significant. We also analysed the combination of one or more positive skin prick tests plus each of the allergic symptoms and diagnoses from the ISAAC questionnaire. The cluster adjusted odds ratios were very close to 1.0 for all of these combinations (data not shown).

As suggested by the high intraclass correlation coefficients and the wide confidence intervals for the odds ratios (table 5), skin prick test results were extremely variable and highly clustered among the 31 polyclinics. In fact, the range among the 31 clinics was from 4% to 85% with one or more positive tests. Six of the polyclinics had positivity rates of ≥10% to each of the five test antigens, which was considerably higher than those at the 25 other polyclinics and inconsistent with the low prevalence of atopic symptoms and diagnoses as revealed by the ISAAC questionnaire. Despite repeated questioning, the paediatricians indicated no problems with the technique used for administering

Table 6 | Results of sensitivity analysis for skin prick test results. Values are numbers (percentages) positive unless stated otherwise

Antigen	Experimental group (n=4100)	Control group (n=4906)	ICC	Cluster adjusted odds ratio* (95% CI)
House dust mite	504 (12.3)	299 (6.1)	0.04	2.0 (1.2 to 3.4)
Cat	347 (8.5)	182 (3.7)	0.05	2.1 (1.1 to 3.9)
Birch pollen	273 (6.7)	125 (2.5)	0.03	2.3 (1.3 to 4.1)
Mixed northern grasses	369 (9.0)	209 (4.3)	0.06	1.5 (0.8 to 2.8)
<i>Alternaria</i>	258 (6.3)	77 (1.6)	0.05	3.5 (1.6 to 7.7)
≥1 positive	929 (22.7)	579 (11.8)	0.07	2.0 (1.1 to 3.4)

ICC=intraclass correlation coefficient.

*For experimental group versus control group.

the skin prick test, nor were any problems detected during monitoring visits carried out soon after data collection had begun. Redefining positive skin prick tests on the sole basis of the size of the wheal did not reduce the high positivity rates in the six discrepant clinics, nor did it uncover any protective effects of the experimental intervention (data available on request).

We did a sensitivity analysis (n=9006) after excluding the six polyclinic sites with high rates of positive skin prick test results (three of these polyclinics were in the experimental group and three were in the control group). Table 6 shows the results of the sensitivity analysis. Intraclinic clustering of skin prick test results was reduced substantially. The proportions of positive test results were considerably lower in both the experimental and control groups, yet the differences between the two groups increased, with significantly elevated odds ratios in the 2-3 range for all but one antigen (mixed northern grasses).

Statistical models with interaction terms showed no evidence that the overall effects of the intervention on allergic symptoms and diagnoses were modified by family history of atopy, except for a history of ever wheezing, for which the experimental intervention increased the risk to a greater extent among those with a positive family history (P=0.005). For the skin prick tests, we found a significant interaction in the same direction (greater risk increase in those with a positive family history) only for mixed northern grasses (P=0.02).

Finally, we also examined the observational associations between the duration and exclusivity of breast feeding and the prevalence of allergic symptoms and diagnoses and skin prick test results. For allergic symptoms and diagnoses, we found borderline significant reductions in history of eczema both with more prolonged any breast feeding and with more prolonged exclusive breast feeding (P=0.08 for both associations, based on χ^2 tests for trend). Associations were stronger and in the opposite direction for skin prick test results, especially after elimination of the six sites with high rates of positive tests; we found highly significant increases in positive skin prick test results with exclusive breast feeding for 3 to <6 months and ≥6 months versus <3 months for house dust mite, cat, birth pollen, mixed northern grasses, and *Alternaria* (P<0.001 for all five antigens, based on χ^2 tests for trend).

DISCUSSION

The results from this large cluster randomised trial indicate that the experimental intervention to promote breast feeding did not reduce the risk of asthma, hay fever, or eczema at age 6.5 years despite large increases in the duration and exclusivity of breast feeding; nor did the intervention succeed in reducing the prevalence of positive skin prick tests. We observed high inter-paediatrician variability (and consequently wide confidence intervals around the adjusted odds ratios) in results of skin prick tests. After exclusion of six polyclinics with suspiciously high rates of positive skin prick test results, the relative odds of positive skin

WHAT IS ALREADY KNOWN ON THIS TOPIC

Evidence is conflicting as to whether prolonged and exclusive breast feeding increases, decreases, or has no effect on the risks of asthma and allergy. All of the available evidence is based on observational studies.

WHAT THIS STUDY ADDS

Prolonged and exclusive breast feeding had no protective effect on allergic symptoms and diagnoses or on positive skin prick tests.

prick tests were twofold to threefold higher in the experimental group than in the control group.

These results conflict with some previous studies suggesting that prolonged and exclusive breast feeding reduces the risk of asthma, other allergic diseases, and atopy skin tests.^{4-6,9-12,27} On the other hand, as discussed in the introduction, the evidence is far from uniform, and several recent studies have even found breast feeding (or more exclusive and prolonged breast feeding) to be associated with increased risks of these outcomes.¹⁵⁻²⁰ Some investigators have found stronger protective effects of breast feeding in offspring with a positive family history of atopy,^{4,5} but we did not. The prevalences of all allergic symptoms and diagnoses were lower among PROBIT children than are customarily seen in Western industrialised countries but similar to those previously reported from Eastern Europe.²⁸ Caution is therefore advised in extrapolating our results to settings where atopic diseases occur more frequently.

The absence of a protective effect against eczema also conflicts with our previous findings based on follow-up during infancy.²² The extremely low reported histories of eczema at 6.5 years (table 4), however, are likely to be gross underestimates. These histories may reflect the absence of an eczema “label” transmitted to the parents but almost certainly constitute a less objective result than that previously reported from PROBIT, in which the diagnosis was based on a standard algorithm using physical findings and symptom histories documented by paediatricians on six occasions during the first 12 months of life.²²

The finding of exceedingly high rates of positive skin tests at six of the study sites, equally divided between experimental and control groups, is not easily explained. Something systematic clearly occurred at these sites, leading to our decision to eliminate them in the sensitivity analysis. These six study sites had no obvious links. They were dispersed geographically; two were in Vitebsk region (north), and one each were in Grodno region (northwest), Gomel region (south), Mogilev region (east), and Minsk region (centre). Two were in urban areas (Retchitsa and Mogilev), whereas the remaining four were rural (Glubokoe, Berestovitsa, Dokshitsy, and Minsk Regional). More importantly, however, the six sites did not have higher prevalences of allergic symptoms and diagnoses (on the basis of the ISAAC questionnaire) than the 25 sites with much lower skin prick test positivity rates. Skin testing is influenced by the technique of the tester,

the device used for administration, the potency of the extracts, the spacing between the antigens, the position of the arm used for placement, and even the time of day.^{29,30} All testers were trained by the same investigator (BM), all centres were provided with the same testing equipment, and the extracts were prepared from the same lots for all participants. Other possible explanations include failure to wipe the testing needle after histamine or antigens, contamination of the extract preparation with histamine, trauma with the skin prick test needle, and placement of the test antigens too closely together. The small number ($n=54$) of children with repeat skin tests at the audit, and the auditing paediatricians’ use of the polyclinics’ antigen solutions, prevented use of the audit data to confirm or refute these possibilities.

Given that we found significantly increased risks of positive skin prick tests in the experimental group only after excluding the six suspect polyclinics, we cannot be confident that the experimental intervention actually caused the increased risks. We feel on safer ground in inferring no reduction in risk. Given these results based on a large randomised trial and the inconsistent benefits reported in previous studies, public health measures to increase the initiation, duration, and exclusivity of breast feeding seem unlikely to have a major impact on reducing the incidence of atopic diseases. The fact that most atopic outcomes have increased in incidence over the past several decades,³¹ simultaneous with the renaissance in breast feeding, strongly suggests that breast feeding does not have a potent protective effect at the population level. Thus, our results underline the importance of seeking other explanations for the recent epidemic of allergy and asthma and of investigating other potential causative factors to develop and test new preventive interventions.

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